



Limiting behavior of delayed sums under a non-identically distribution setup

CHEN PINGYAN

Department of Mathematics, Jinan University, Guangzhou, 510630, P.R. China

*Manuscript received on July 28, 2007; accepted for publication on June 9, 2008;
 presented by DJAIRO G. FIGUEIREDO*

ABSTRACT

We present an accurate description the limiting behavior of delayed sums under a non-identically distribution setup, and deduce Chover-type laws of the iterated logarithm for them. These complement and extend the results of Vasudeva and Divanji (Theory of Probability and its Applications, 37 (1992), 534–542).

Key words: stable distribution, laws of iterated logarithm, delayed sum.

1 INTRODUCTION AND MAIN RESULTS

The distribution function F of a real valued random variable X is called stable law with exponent α ($0 < \alpha < 2$), if for some $\sigma > 0$, $-1 \leq \beta \leq 1$, its characteristic function is of the form

$$E \exp(itX) = \exp \left\{ -\sigma |t|^\alpha \left(1 + i\beta \frac{t}{|t|} \omega(t, \alpha) \right) \right\}, \quad t \in \mathbb{R} \quad (1.1)$$

where

$$\omega(t, \alpha) = \begin{cases} \tan \frac{\pi\alpha}{2}, & \text{if } \alpha \neq 1, \\ \frac{2}{\pi} \ln |t|, & \text{if } \alpha = 1. \end{cases}$$

If $\beta = 0$, X is a symmetric random variable. It is well-known, if F is a stable law with exponent α ($0 < \alpha < 2$), we have the following tail behavior:

$$\lim_{t \rightarrow \infty} t^\alpha (1 - F(t) + F(-t)) = c(\alpha, \sigma), \quad (1.2)$$

where $c(\alpha, \sigma) > 0$ only depends on α and σ (cf. e.g. Feller 1971). This property will play an important role in this paper.

Let $\{X_n, n \geq 1\}$ be a sequence of independent random variables with its partial sums $S_n = \sum_{i=1}^n X_i$. Let $\{a_n, n \geq 1\}$ be a positive integer subsequence. Set $T_n = S_{n+a_n} - S_n$ and $\gamma_n = \log(n/a_n) + \log \log n$.

The sum T_n is called a forward delayed sum (see Lai 1974). Suppose X_n 's involve of two distributions F_1 and F_2 which are stable laws with exponents α_1 and α_2 ($0 < \alpha_1 \leq \alpha_2 < 2$). For each $n \geq 1$, let $\tau_1(n)$ denote the number of random variables in the set $\{X_1, X_2, \dots, X_n\}$ with distribution function F_1 , then $\tau_2(n) = n - \tau_1(n)$ is the number of random variables with distribution function F_2 in the set $\{X_1, X_2, \dots, X_n\}$. Then $(\tau_1(n), \tau_2(n))$ is called the sample scheme of the sequence $\{X_n, n \geq 1\}$. Assume that $\tau_1(n) = [n^{\alpha_1/\alpha_2}]$ and $B_n = n^{1/\alpha_2}$, where $[x]$ is the integer part of x . By Sreehari (1970), S_n/B_n converges weakly to a composition of the two stable laws.

Let $U_{\tau_1(n)}$ be the sum of those $\{X_1, X_2, \dots, X_n\}$ with distribution function F_1 and $V_{\tau_2(n)}$ be the sum of those $\{X_1, X_2, \dots, X_n\}$ with distribution function F_2 . Then $S_n = U_{\tau_1(n)} + V_{\tau_2(n)}$. One can note that in T_n there are $[(n + a_n)^{\alpha_1/\alpha_2}] - [n^{\alpha_1/\alpha_2}]$ random variables with distribution function F_1 and $n + a_n - [(n + a_n)^{\alpha_1/\alpha_2}] - (n - [n^{\alpha_1/\alpha_2}])$ random variables with distribution function F_2 .

The motivation of this paper is to extend and complement the results of Vasudeva and Divanji (1992). They obtained the following theorem in the special case that F_1 and F_2 are positive stable laws with exponents $0 < \alpha_1 \leq \alpha_2 < 1$.

THEOREM A. Let $\{a_n, n \geq 1\}$ be a nondecreasing sequence with $0 < a_n \leq n$ and a_n/n non-increasing. Let F_1 and F_2 are positive stable law and $0 < \alpha_1 \leq \alpha_2 < 1$.

(i) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = +\infty$, then

$$\limsup_{n \rightarrow \infty} \left(\frac{T_n}{B_{a_n}} \right)^{1/\gamma_n} = e^{1/\alpha_2} \text{ a.s.}$$

(ii) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = 0$, then

$$\limsup_{n \rightarrow \infty} \left(\frac{T_n}{B_{a_n}} \right)^{1/\gamma_n} = e^{1/\alpha_1} \text{ a.s.}$$

(iii) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = s \in (0, +\infty)$, then

$$\limsup_{n \rightarrow \infty} \left(\frac{T_n}{B_{a_n}} \right)^{1/\gamma_n} = \exp \left\{ \frac{\alpha_1 s + \alpha_2}{(s + 1)\alpha_1 \alpha_2} \right\} \text{ a.s.}$$

They only discuss the case that F_1 and F_2 are positive stable law with exponents $0 < \alpha_1 \leq \alpha_2 < 1$. But by their method, it is impossible to discuss the rest case. In this paper, by a new method, we will complement and extend Theorem A in three directions, namely:

- (i) We will obtain more exact results.
- (ii) We will discuss not only that the distributions is the positive stable laws, but also that the distributions is not necessary positive stable laws and the exponents of the stable laws in $(0, 2)$, not only in $(0, 1)$.
- (iii) We will replace the restrictions $0 < a_n \leq n$ and a_n/n non-increasing of the sequence $\{a_n, n \geq 1\}$ by a more general assumption $\limsup_{n \rightarrow \infty} a_n/n < +\infty$.

Recall that the kind of type law of the iterated logarithm (LIL) was first obtained by Chover (1966) for symmetric stable law, and is called Chover-type LIL. By far, some papers concern with the Chover-type LIL, for example, Chen (2002) for the weighted sums of symmetric stable law, Chen and Yu (2003) for the weighted sums of stable law without symmetric assumption, Peng and Qi (2003) for the weighted sums of law in the domain of attraction of stable law, and Chen (2004) for geometric weighted sums and Cesàro weighted sums of stable law, etc.

First we give an accurate description of the limiting behavior of S_n .

THEOREM 1.1. Let $f > 0$ be a nondecreasing function. Then with probability one

$$\limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(f(n))^{1/\alpha_1}} = \begin{cases} 0, \\ +\infty, \end{cases} \Leftrightarrow \int_1^{+\infty} \frac{dx}{xf(x)} \begin{cases} < +\infty, \\ = +\infty. \end{cases} \quad (1.3)$$

By Theorem 1.1, we have the following Corollary at once.

COROLLARY 1.1. For every $\delta > 0$, we have

$$\limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(\log n)^{(1+\delta)/\alpha_1}} = 0 \quad a.s.$$

and

$$\limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(\log n)^{1/\alpha_1}} = +\infty \quad a.s.$$

In particular

$$\limsup_{n \rightarrow \infty} \left| \frac{S_n}{B_n} \right|^{1/\log \log n} = e^{1/\alpha_1} \quad a.s. \quad (1.4)$$

REMARK 1.1. If $\alpha_1 = \alpha_2$, Corollary 1.1 extends the result of Chover (1966).

THEOREM 1.2. Let $\{a_n, n \geq 1\}$ be a subsequence of positive integers with $\limsup_{n \rightarrow \infty} a_n/n < +\infty$. Let $f > 0$ be a nondecreasing function. Then with probability one

$$\limsup_{n \rightarrow \infty} \frac{|T_n|}{B_n(f(n))^{1/\alpha_1}} = \begin{cases} 0, \\ +\infty, \end{cases} \Leftrightarrow \int_1^{+\infty} \frac{dx}{xf(x)} \begin{cases} < +\infty, \\ = +\infty. \end{cases} \quad (1.5)$$

COROLLARY 1.2. Let $\{a_n, n \geq 1\}$ as Theorem 1.2. Then for every $\delta > 0$, we have

$$\limsup_{n \rightarrow \infty} \frac{|T_n|}{B_n(\log n)^{(1+\delta)/\alpha_1}} = 0 \quad a.s.$$

and

$$\limsup_{n \rightarrow \infty} \frac{|T_n|}{B_n(\log n)^{1/\alpha_1}} = +\infty \quad a.s.$$

In particular

$$\limsup_{n \rightarrow \infty} \left| \frac{T_n}{B_n} \right|^{1/\log \log n} = e^{1/\alpha_1} \quad a.s. \quad (1.6)$$

COROLLARY 1.3. Let $\{a_n, n \geq 1\}$ as Theorem 1.2.

(i) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = +\infty$, then

$$\limsup_{n \rightarrow \infty} \left| \frac{T_n}{B_{a_n}} \right|^{1/\gamma_n} = e^{1/\alpha_2} \text{ a.s.} \tag{1.7}$$

(ii) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = 0$, then

$$\limsup_{n \rightarrow \infty} \left| \frac{T_n}{B_{a_n}} \right|^{1/\gamma_n} = e^{1/\alpha_1} \text{ a.s.} \tag{1.8}$$

(iii) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = s \in (0, +\infty)$, then

$$\limsup_{n \rightarrow \infty} \left| \frac{T_n}{B_{a_n}} \right|^{1/\gamma_n} = \exp \left\{ \frac{\alpha_1 s + \alpha_2}{(s + 1)\alpha_1 \alpha_2} \right\} \text{ a.s.} \tag{1.9}$$

COROLLARY 1.4. Let $\{a_n, n \geq 1\}$ as Theorem 1.2. If $\alpha_1 = \alpha_2 = \alpha$, then

$$\limsup_{n \rightarrow \infty} \left| \frac{T_n}{B_{a_n}} \right|^{1/\gamma_n} = e^{1/\alpha} \text{ a.s.} \tag{1.10}$$

REMARK 1.2. Corollary 1.4 extends the result of Zinchenko (1994).

2 PROOFS OF THE MAIN RESULTS

We need the following lemmas.

LEMMA 2.1 (see Lemma 2.1 of Chen 2004). Let $f > 0$ be a non-decreasing function with

$$\int_1^\infty \frac{dx}{xf(x)} < +\infty,$$

then there exists a non-decreasing function $g > 0$ such that

$$g(x) \leq f(x), \quad \limsup_{x \rightarrow +\infty} g(2x)/g(x) < +\infty \quad \text{and} \quad \int_1^\infty \frac{dx}{xg(x)} < +\infty.$$

LEMMA 2.2 (see Lemma 2.2 of Chen 2002). Let $f > 0$ be a non-decreasing function satisfying

$$\int_1^\infty \frac{dx}{xf(x)} = +\infty.$$

Then there exists a non-decreasing function $h > 0$ such that

$$h(x) \rightarrow +\infty \text{ as } x \rightarrow +\infty \quad \text{and} \quad \int_1^\infty \frac{dx}{xf(x)h(x)} = +\infty.$$

LEMMA 2.3 (see Lemma 3 of Chow and Lai 1973). Let $\{W_n, n \geq 1\}$ and $\{Z_n, n \geq 1\}$ be two sequences of random variables such that $\{W_i, 1 \leq i \leq n\}$ and Z_n are independent for each $n \geq 1$. Suppose $W_n + Z_n \rightarrow 0$ a.s. and $Z_n \rightarrow 0$ in probability, then $W_n \rightarrow 0$ a.s. and $Z_n \rightarrow 0$ a.s.

In the rest of this paper, we denote C as a generic positive number which may be different at different places, and $a(n) \sim b(n)$ means $\lim_{n \rightarrow \infty} a(n)/b(n) = 1$. For the sake of simplicity, we denote random variable Y_1 with distribution function F_1 and random variable Y_2 with distribution function F_2 .

PROOF OF THEOREM 1.1. Assume that $\int_1^\infty \frac{dx}{xf(x)} < \infty$. First of all, we show that

$$\frac{S_n}{B_n(f(n))^{1/\alpha_1}} \rightarrow 0 \text{ in probability.} \tag{2.1}$$

Note that by (1.1), $(\tau_1(n))^{-1/\alpha_1}(U_{\tau_1(n)} - b_{\tau_1(n)})$ has the same distribution as Y_1 and $(\tau_2(n))^{-1/\alpha_2}(V_{\tau_2(n)} - d_{\tau_2(n)})$ has the same distribution as Y_2 , where $b_n = 0$ if $\alpha_1 \neq 1$ and $b_n = bn \log n$ for some $b \in (-\infty, +\infty)$ if $\alpha_1 = 1$, and $d_n = 0$ if $\alpha_2 \neq 1$ and $d_n = dn \log n$ for some $d \in (-\infty, +\infty)$ if $\alpha_2 = 1$. $\int_1^\infty \frac{dx}{xf(x)} < \infty$ implies that $f(n) \rightarrow \infty$ and $\log n/f(n) \rightarrow 0$ as $n \rightarrow \infty$. Hence for all $\varepsilon > 0$

$$\begin{aligned} P(|U_{\tau_1(n)} - b_{\tau_1(n)}| \geq \varepsilon B_n(f(n))^{1/\alpha_1}) &= P(|Y_1| \geq \varepsilon B_n(f(n))^{1/\alpha_1}/(\tau_1(n))^{1/\alpha_1}) \\ &\sim Cn^{-\alpha_1/\alpha_2}(f(n))^{-1} \tau_1(n) \\ &\sim C(f(n))^{-1} \rightarrow 0, \quad n \rightarrow \infty \end{aligned}$$

and

$$\begin{aligned} P(|V_{\tau_2(n)} - d_{\tau_2(n)}| \geq \varepsilon B_n(f(n))^{1/\alpha_1}) &= P(|Y_2| \geq \varepsilon B_n(f(n))^{1/\alpha_1}/(\tau_2(n))^{1/\alpha_2}) \\ &\sim Cn^{-1}(f(n))^{-\alpha_2/\alpha_1} \tau_2(n) \\ &\sim C(f(n))^{-\alpha_2/\alpha_1} \rightarrow 0, \quad n \rightarrow \infty. \end{aligned}$$

Hence (2.1) holds. So by standard symmetric argument (see Lemma 3.2.1 of Stout 1974), we need only to prove the result for $\{X_n, n \geq 1\}$ symmetric.

By Lemma 2.1 of Chen (2002),

$$\frac{U_{\tau_1(n)}}{(\tau_1(n)f(\tau_1(n)))^{1/\alpha_1}} \rightarrow 0 \text{ a.s.} \quad \text{and} \quad \frac{V_{\tau_2(n)}}{(\tau_2(n)f(\tau_2(n)))^{1/\alpha_2}} \rightarrow 0 \text{ a.s.}$$

Note that

$$\limsup_{n \rightarrow \infty} \frac{(\tau_1(n)f(\tau_1(n)))^{1/\alpha_1}}{B_n(f(n))^{1/\alpha_1}} < \infty \quad \text{and} \quad \limsup_{n \rightarrow \infty} \frac{(\tau_2(n)f(\tau_2(n)))^{1/\alpha_2}}{B_n(f(n))^{1/\alpha_1}} < \infty.$$

Hence

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(f(n))^{1/\alpha_1}} &\leq \limsup_{n \rightarrow \infty} \frac{|U_{\tau_1(n)}|}{B_n(f(n))^{1/\alpha_1}} + \limsup_{n \rightarrow \infty} \frac{|V_{\tau_2(n)}|}{B_n(f(n))^{1/\alpha_1}} \\ &\leq \limsup_{n \rightarrow \infty} \frac{(\tau_1(n)f(\tau_1(n)))^{1/\alpha_1}}{B_n(f(n))^{1/\alpha_1}} \times \frac{|U_{\tau_1(n)}|}{(\tau_1(n)f(\tau_1(n)))^{1/\alpha_1}} \\ &\quad + \limsup_{n \rightarrow \infty} \frac{(\tau_2(n)f(\tau_2(n)))^{1/\alpha_2}}{B_n(f(n))^{1/\alpha_1}} \times \frac{|V_{\tau_2(n)}|}{(\tau_2(n)f(\tau_2(n)))^{1/\alpha_2}} \\ &= 0 \text{ a.s.} \end{aligned}$$

So we complete the proof of the convergent part.

Now we assume that $\int_1^\infty \frac{dx}{xf(x)} = +\infty$. If

$$\sum_{n=1}^\infty P(|X_n| \geq MB_n(f(n))^{1/\alpha_1}) = +\infty, \quad \forall M > 0 \tag{2.2}$$

holds, then by the Borel-Cantelli lemma, we have

$$\limsup_{n \rightarrow \infty} \frac{|X_n|}{B_n(f(n))^{1/\alpha_1}} = +\infty \text{ a.s.}$$

and note that

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{|X_n|}{B_n(f(n))^{1/\alpha_1}} &\leq \limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(f(n))^{1/\alpha_1}} + \limsup_{n \rightarrow \infty} \frac{B_{n-1}(f(n-1))^{1/\alpha_1}}{B_n(f(n))^{1/\alpha_1}} \times \frac{|S_{n-1}|}{B_{n-1}(f(n-1))^{1/\alpha_1}} \\ &\leq 2 \limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(f(n))^{1/\alpha_1}}, \end{aligned}$$

hence we have

$$\limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(f(n))^{1/\alpha_1}} = +\infty \text{ a.s.}$$

Now we prove (2.2). Note that

$$\begin{aligned} \sum_{n=1}^\infty P(|X_n| \geq MB_n(f(n))^{1/\alpha_1}) &= \sum_{k=0}^\infty \sum_{n=2^k}^{2^{k+1}-1} P(|X_n| \geq MB_n(f(n))^{1/\alpha_1}) \\ &\geq \sum_{k=0}^\infty \sum_{n=2^k}^{2^{k+1}-1} P(|X_n| \geq MB_{2^{k+1}}(f(2^{k+1}))^{1/\alpha_1}) \\ &\geq \sum_{k=0}^\infty (\tau_1(2^{k+1} - 1) - \tau_1(2^k - 1)) P(|Y_1| \geq MB_{2^{k+1}}(f(2^{k+1}))^{1/\alpha_1}) \\ &\geq C \sum_{k=0}^\infty (\tau_1(2^{k+1} - 1) - \tau_1(2^k - 1)) (2^{k+1})^{-\alpha_1/\alpha_2} (f(2^{k+1}))^{-1} \\ &\geq C \sum_{k=0}^\infty (f(2^{k+1}))^{-1} \end{aligned}$$

and $\int_1^\infty \frac{dx}{xf(x)} = +\infty$ implies $\sum_{k=0}^\infty (f(2^{k+1}))^{-1} = +\infty$, so (2.2) holds.

PROOF OF THEOREM 1.2. Assume that $\int_1^\infty \frac{dx}{xf(x)} < \infty$, by Lemma 2.1, without loss of generality, we can assume that $\limsup_{x \rightarrow \infty} f(2x)/f(x) < \infty$. By Theorem 1.1, we have

$$\limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(f(n))^{1/\alpha_1}} = 0 \text{ a.s.} \quad \text{and} \quad \limsup_{n \rightarrow \infty} \frac{|S_{n+a_n}|}{B_{n+a_n}(f(n+a_n))^{1/\alpha_1}} = 0 \text{ a.s.}$$

Note that $\limsup_{n \rightarrow \infty} \frac{B_{n+a_n}(f(n+a_n))^{1/\alpha_1}}{B_n(f(n))^{1/\alpha_1}} < \infty$, hence

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{|T_n|}{B_n(f(n))^{1/\alpha_1}} &\leq \limsup_{n \rightarrow \infty} \frac{|S_{n+a_n}|}{B_n(f(n))^{1/\alpha_1}} + \limsup_{n \rightarrow \infty} \frac{|S_n|}{B_n(f(n))^{1/\alpha_1}} \\ &= \limsup_{n \rightarrow \infty} \frac{B_{n+a_n}(f(n+a_n))^{1/\alpha_1}}{B_n(f(n))^{1/\alpha_1}} \times \frac{|S_{n+a_n}|}{B_{n+a_n}(f(n+a_n))^{1/\alpha_1}} \\ &= 0 \text{ a.s.} \end{aligned}$$

Now we assume that $\int_1^\infty \frac{dx}{xf(x)} = +\infty$. Suppose

$$\limsup_{n \rightarrow \infty} \frac{|T_n|}{B_n(f(n))^{1/\alpha_1}} = +\infty \text{ a.s.}$$

does not hold, then by Kolmogorov 0-1 law, there exists a constant $c_0 \in [0, \infty)$ such that

$$\limsup_{n \rightarrow \infty} \frac{|T_n|}{B_n(f(n))^{1/\alpha_1}} = c_0 \text{ a.s.}$$

Hence

$$\lim_{n \rightarrow \infty} \frac{T_n}{B_n(f(n)h(n))^{1/\alpha_1}} = 0 \text{ a.s.}$$

where $h(x)$ is given by Lemma 2.2. It is easy to show that

$$\frac{X_{n+1}}{B_n(f(n)h(n))^{1/\alpha_1}} \rightarrow 0 \text{ in probability,}$$

i.e.

$$\frac{T_n - X_{n+1}}{B_n(f(n)h(n))^{1/\alpha_1}} \rightarrow 0 \text{ in probability.}$$

By Lemma 2.3, we have

$$\frac{X_{n+1}}{B_n(f(n)h(n))^{1/\alpha_1}} \rightarrow 0 \text{ a.s.}$$

By the Borel-Cantelli lemma

$$\sum_{n=1}^\infty P(|X_n| \geq B_n(f(n)h(n))^{1/\alpha_1}) < \infty.$$

But by the same argument in the proof of Theorem 1.1, we have

$$\sum_{n=1}^\infty P(|X_n| \geq B_n(f(n)h(n))^{1/\alpha_1}) = \infty.$$

This leads to a contradiction, so we complete the proof.

PROOF OF COROLLARY 1.3. By Theorem 1.2, we have

$$\limsup_{n \rightarrow \infty} \frac{|T_n|}{B_n(\log n)^{(1+\delta)/\alpha_1}} = 0 \text{ a.s. } \forall \delta > 0 \quad \text{and} \quad \limsup_{n \rightarrow \infty} \frac{|T_n|}{B_n(\log n)^{1/\alpha_1}} = +\infty \text{ a.s.}$$

Hence we have

$$P(|T_n| \geq B_n(\log n)^{(1+\delta)/\alpha_1}, \text{ i.o.}) = 0, \quad \forall \delta > 0 \quad \text{and} \quad P(|T_n| \geq B_n(\log n)^{1/\alpha_1}, \text{ i.o.}) = 1,$$

where $P(A_n, \text{ i.o.}) = P(\limsup_{n \rightarrow \infty} A_n)$ and A_n is a sequence of events. So we have

$$P\left(\log \left| \frac{T_n}{B_{a_n}} \right| \geq (1/\alpha_2) \log(n/a_n) + ((1+\delta)/\alpha_1) \log \log n, \text{ i.o.}\right) = 0, \quad \forall \delta > 0,$$

and

$$P\left(\log \left| \frac{T_n}{B_{a_n}} \right| \geq (1/\alpha_2) \log(n/a_n) + (1/\alpha_1) \log \log n, \text{ i.o.}\right) = 1.$$

(i) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = \infty$, then

$$P\left(\log \left| \frac{T_n}{B_{a_n}} \right| \geq (1+\delta_1)\gamma_n/\alpha_2, \text{ i.o.}\right) = 0, \quad \forall \delta_1 > 0$$

and

$$P\left(\log \left| \frac{T_n}{B_{a_n}} \right| \geq (1-\delta_2)\gamma_n/\alpha_2, \text{ i.o.}\right) = 1, \quad \forall \delta_2 > 0,$$

hence we have

$$\limsup_{n \rightarrow \infty} \left| \frac{T_n}{B_{a_n}} \right|^{1/\gamma_n} = e^{1/\alpha_2} \text{ a.s.}$$

(ii) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = 0$, then

$$P\left(\log \left| \frac{T_n}{B_{a_n}} \right| \geq (1+\delta_3)\gamma_n/\alpha_1, \text{ i.o.}\right) = 0, \quad \forall \delta_3 > 0$$

and

$$P\left(\log \left| \frac{T_n}{B_{a_n}} \right| \geq (1-\delta_4)\gamma_n/\alpha_1, \text{ i.o.}\right) = 1, \quad \forall \delta_4 > 0,$$

hence we have

$$\limsup_{n \rightarrow \infty} \left| \frac{T_n}{B_{a_n}} \right|^{1/\gamma_n} = e^{1/\alpha_1} \text{ a.s.}$$

(iii) If $\lim_{n \rightarrow \infty} \log(n/a_n)/\log \log n = s \in (0, \infty)$, then

$$P\left(\log \left| \frac{T_n}{B_{a_n}} \right| \geq \left(\frac{\alpha_1 s + \alpha_2}{\alpha_1 \alpha_2 (s+1)} + \delta_5 \right) \gamma_n, \text{ i.o.}\right) = 0, \quad \forall \delta_5 > 0$$

and

$$P\left(\log \left| \frac{T_n}{B_{a_n}} \right| \geq \left(\frac{\alpha_1 s + \alpha_2}{\alpha_1 \alpha_2 (s+1)} - \delta_6 \right) \gamma_n, \text{ i.o.}\right) = 1, \quad \forall \delta_6 > 0,$$

hence we have

$$\limsup_{n \rightarrow \infty} \left| \frac{T_n}{B_{a_n}} \right|^{1/\gamma_n} = \exp\left(\frac{\alpha_1 s + \alpha_2}{\alpha_1 \alpha_2 (s+1)}\right) \text{ a.s.}$$

ACKNOWLEDGMENTS

Research supported by the National Natural Science Foundation of China. The author is thankful to the referee for her/his helpful remarks which improved the presentation of the paper.

RESUMO

Apresentamos uma descrição precisa do comportamento limite de somas retardadas, e deduzimos leis do tipo Chover de logaritmo iterado para as mesmas. Isso completa e estende os resultados de Vasudeva e Divanji (Theory of Probability and its Applications, 37 (1992), 534–542).

Palavras-chave: distribuição estável, leis do logaritmo iterado, somas retardadas.

REFERENCES

- CHEN PY. 2002. Limiting behavior of weighted sums with stable distributions. *Statist Probab Letters* 60: 367–375.
- CHEN PY. 2004. On Chover's LIL for Weighted Sums of Stable random Variables. *Sankhya: Indian J Statist* 66: 49–62.
- CHEN PY AND YU JH. 2003. On Chover's LIL for the weighted sums of stable random variables. *Acta Math Sci Ser B Engl*, ed. 23: 74–82.
- CHOVER J. 1966. A law of the iterated logarithm for stable summands. *Proc Amer Math Soc* 17: 441–443.
- CHOW YS AND LAI TL. 1973. Limiting behavior of weighted sums of independent random variables. *Ann Probab* 1: 810–824.
- FELLER W. 1971. An introduction to probability theory and its applications. Vol II. J. Wiley & Sons, New York.
- LAI TL. 1974. Limit theorems for delayed sums. *Ann Probab* 2: 432–440.
- PENG L AND QI YC. 2003. Chover-type laws of the iterated logarithm for weighted sums. *Statist Probab Letters* 65: 401–410.
- SREEHARI M. 1970. On a class of limit distributions for normalized sums of independent random variables. *Theory Probab Appl* 15: 269–290.
- STOUT WF. 1974. Almost sure convergence. Academic Press, New York.
- VASUDEVA R AND DIVANJI G. 1992. LIL for delayed sums under a non-identically distribution setup. *Theory Probab Apl* 37: 534–542.
- ZINCHENKO NM. 1994. A modified law of iterated logarithm for safe stable random variables. *Theory Probab Math Statist* 49: 69–76.