Differences in walking attainment ages between low-risk preterm and healthy full-term infants

Diferenças na idade de aquisição da marcha entre lactentes pré-termo de baixo risco e a termo saudáveis

Ana P. Restiffe¹, José Luiz D. Gherpelli²

ABSTRACT

Objective: To compare gross motor development of preterm infants (PT) without cerebral palsy with healthy full-term (FT) infants, according to Alberta Infant Motor Scale (AIMS); to compare the age of walking between PT and FT; and whether the age of walking in PT is affected by neonatal variables. Methods: Prospective study compared monthly 101 PT and 52 FT, from the first visit, until all AIMS items had been observed. Results: Mean scores were similarity in their progression, except from the eighth to tenth months. FT infants were faster in walking attainment than PT. Birth weight and length and duration of neonatal nursery stay were related to walking delay. Conclusion: Gross motor development between PT and FT were similar, except from the eighth to tenth months of age. PT walked later than FT infants and predictive variables were birth weight and length, and duration of neonatal intensive unit stay.

Key words: Alberta Infant Motor Scale, child development, infant, premature, motor activity.

RESUMO

Objetivo: Comparar o desenvolvimento motor de lactentes pré-termo sem paralisia cerebral (PT) com lactentes normais nascidos a termo (T), de acordo com Escala Motora Infantil de Alberta (AIMS); comparar idade da marcha entre PT e T e se a idade da marcha em PT é passível de ser afetada. Métodos: Estudo prospectivo com 101 PT e 52 T, seguidos mensalmente até que todos os itens da AIMS tivessem sido observados. Resultados: Os escores médios apresentaram semelhanças entre os grupos, com exceção do oitavo ao décimo meses. Os lactentes T iniciaram marcha antes dos PT. Peso, estatura ao nascimento e tempo de internação na unidade de terapia intensiva neonatal (UTIN) foram preditivos. Conclusão: O desenvolvimento motor entre PT e T foi semelhante, exceto no oitavo e décimo meses de idade. PT andaram mais tardiamente e variáveis preditivas foram peso, estatura ao nascimento e o tempo de permanência na UTIN. Palavras-Chaves: Escala Motora Infantil de Alberta, desenvolvimento infantil, lactente, prematuro, atividade motora.

During the past decades there was an increase in the survival rate of infants born very preterm (gestational age <32 weeks) and with very low birth weight (VLBW). However, this was associated with an increased rate of adverse developmental outcome, which in turn has raised important issues regarding the quality of life of these infants¹.

The incidence cerebral palsy (CP) ranges from 5 to 15% and is inversely related to birth weight and gestational age. CP is often detected during the first year of life². Minor neurological dysfunctions (MND) such as motor coordination disorder or clumsiness, behavioral problems, and learning disorders range from 40 to 70% and usually become more apparent as the child grows older³.

Whether delay in achieving developmental milestones is a marker of MND⁴ or a variation of normal motor development in preterm infants is a matter still under debate⁵. Developmental abnormalities and delays diagnosed early in life in preterm infants may be transient and may fade away as a result of the interaction between genetic and environmental factors⁶, and also the interaction of various systems, including the neuromuscular, sensory, biomechanical, and central nervous systems⁷.

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Walking is considered an important motor developmental milestone and has a predictive value for future motor abilities, but virtually nothing is known about such relationship. Studies that compared walking onset ages between preterm infants without major neurological disabilities and normal full-term infants were inconclusive. Some studies showed that preterm infants walk at older ages, while others reported walking at similar ages as term infants. Normal full-term infants usually attain independent walking at about 12 months of age, and continuously refine their movement in the following years to establish a mature walking pattern. Accumulating data on preterm infants have shown that the measure of age at onset of walking reflects various degrees of motor delay. These findings suggest that age at onset and quality of walking movement may be useful measures for diagnosing future minor motor abnormalities in preterm infants.

We hypothesized that infants born preterm without major neurological disabilities have a different pattern of early gross motor development compared to healthy full-term infants. More specifically, the onset of independent walking among preterm infants is later than full-term infants. We also hypothesized that these variations are related to specific biological and/or non-biological variables, considered in this investigation as predictor variables.

Herein the objectives of this study were: to compare prospectively gross motor performance of preterm infants with low risk for major neurological disabilities with healthy full-term infants based on the Alberta Infant Motor Scale (AIMS) to compare the age at onset of independent walking between low risk preterm infants and healthy infants born at term and to investigate whether the age of onset walking in infants born preterm was affected by perinatal and/or sociodemographic variables.

**METHODS**

From June 2004 to August 2005, the researcher (A.P.R.) recruited and assessed consecutively and in a non-randomized way, 101 preterm infants with low risk for neurological problems, and 52 healthy full-term infants, at the first hospital follow-up visit. Preterm and full-term infants were followed prospectively and examined monthly, until 18 months of corrected and chronological age, respectively. At this age, all infants underwent a full neuromotor evaluation performed by a neurologist (J.L.D.G.) to exclude a major neurological abnormality (cerebral palsy, severe motor delay, reflex abnormality, muscle tone abnormality, focal neurological signs, persisting asymmetry, or seizures).

All preterm infants were inborn at the Clinics Hospital of the Universidade de São Paulo (USP) Medical School. Inclusion criteria were: gestational age (GA) <37 weeks and the mother’s willingness to participate in the study. Exclusion criteria were: presence of congenital or chromosomal anomalies and major neonatal diseases that could impair neurodevelopment, such as grade III and IV intraventricular hemorrhages, periventricular leukomalacia, moderate or severe hypoxic ischemic encephalopathy, meningitis, encephalitis and cerebral malformation, severe retinopathy of prematurity (stage III-IV), visual and hearing deficits, seizures, and absence on more than three consecutive appointments.

Full-term infants were born at the Hospital of the USP. The inclusion criteria were: singleton, inborn, gestational age between 37 and 42 weeks, Apgar score ≥5 points at 10 minutes, birth weight ≥2.5 kg and absence of maternal and perinatal complications and of significant medical problems in the neonatal period and/or during the follow-up period, plus mother's willingness participate in the study. The exclusion criteria were: presence of abnormal neurological impairments during follow-up or at 18 month chronological-age; presence of delay in the spheres of adaptive, language, and social development; and failure to attend more than three consecutive appointments.

We defined independent walking onset as the moment when the infants were able to move for five successive steps without support.

If an abnormal neurological impairment was detected during follow-up and/or at 18 months of corrected and chronological age the child was excluded. None of children underwent rehabilitation during follow-up.

Gestational age (GA) was estimated based on the date of the mother last menstrual period and confirmed by early ultrasound scan performed before 20 weeks of gestation. Some cases had their GA estimated by the New Ballard assessment scores (preterm infants) or by the Capurro assessment (at term). Corrected age was calculated by subtracting to the degree of prematurity from the chronological age and was used during the study period for all preterm infants.

This study was approved by the Ethics Committees of University Hospital of the USP and the Clinics Hospital of the University of São Paulo Medical School, Brazil before recruitment of participants. All parents signed an informed consent form.

Monthly evaluation assessments were scheduled, either at home or at the hospital, depending on parent’s convenience. One of the authors (A.P.R.), who had been trained for reliability for research use and had achieved 82% of agreement, in a pilot-study which took place before the data collection of the present study, was the only investigator to interact with the child, apart from the parents. Infant motor performance was recorded through videotape that lasted between 20 and 40 minutes and displayed the motor performance repertoire in four positions (prone, supine, sitting, and standing). During the assessment, the infants were laid undressed either on the examining table or on a firm mat. Neither facilitation nor handling was allowed. In order to motivate the infants to move and explore the environment, verbal cues and toys were used. All videotapes recorded were
replayed thoroughly, without being edited, so that the only investigator could score the infant motor function, according to the AIMS assessment tool criteria\textsuperscript{15}.

Perinatal data were collected from medical records. Socioeconomic variables included race and mother’s education and occupational level and socioeconomic classification. The latter was expressed as four indicators: family’s monthly income according to multiples of one minimum wage, household resources, housing quality, and parental literacy\textsuperscript{20}.

**Statistical analysis**

A total sample size of 40 infants was needed to demonstrate comparison group differences in the age of walking attainment equal or greater than two months. It was assumed that all infants (preterm and full term) had a constant chance of a specific event (i.e. independent walking attainment) to happen throughout the observation period. The power of the study was 90% and the significance level was 0.05.

For each month, individual full-term and preterm infants’ AIMS raw scores according to chronological and corrected ages, were clustered to calculate monthly mean scores\textsuperscript{21}. To demonstrate statistical difference between the groups, the variability in the AIMS mean scores was measured by calculating monthly standard error. If monthly standard error scores between comparison groups overlapped, it indicated absence of statistical difference in the AIMS scores between the groups in that specific month. The AIMS mean score was calculated only until children reached 16 months of corrected age because after that almost all infants had completed all the 58 AIMS items.

Walking onset age between preterm and full-term infants was estimated by Kaplan-Meier function and Turnbull methods\textsuperscript{21}.

The relationships between independent variables (perinatal and/or socio-demographic and walking attainment age) were analyzed using Cox’s proportional hazard regression models.

Statistical analyses were performed using SPSS version 10.0, Microsoft Excel version 2002, MINITAB version 14, R version 2.2.0, and SAS System, version 8.0.

**RESULTS**

**Sample characteristics**

Forty-nine full-term infants and 77 preterm infants completed the study. All infants had a normal neurological examination at 18 months. Three full-term infants were excluded from the study: 2 were lost to follow-up and 1 had a mild motor delay at 18 months. Twenty-four preterm infants were excluded: ten infants (10/24; 42%) were diagnosed as having cerebral palsy; six preterm infants (6/24; 25%) were still not walking at the end of the study period; two infants (2/24; 8.5%) died and two (2/24; 8.5%) were lost to follow-up. There were four pairs of identical twins and one infant of each pair (4/24; 16%) was randomly excluded, due to sibling dependence.

Table 1 shows perinatal characteristics of full-term and preterm infants and preterm perinatal morbidities which variables were also used in the Cox’s proportional hazard regression models. Analysis of socioeconomic characteristics revealed similarly distribution across the studied groups as shown in Table 2.

<table>
<thead>
<tr>
<th>Variables</th>
<th>Full-term</th>
<th>Preterm</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total (%)</td>
<td>49 (100)</td>
<td>77 (100)</td>
</tr>
<tr>
<td>Mean gestational age (range) (weeks)</td>
<td>39.6 (37.1–42.0)</td>
<td>31.9 (25.7–36.0)</td>
</tr>
<tr>
<td>Moderate premature (%)/extreme premature (%)</td>
<td>–</td>
<td>30 (39/36 (47)</td>
</tr>
<tr>
<td>Mean birth weight (range) (grams)</td>
<td>3,178 (2,500–4,020)</td>
<td>1,505 g (590–2,500)</td>
</tr>
<tr>
<td>Low birth weight (%)/Very-low birth weight (%)</td>
<td>–</td>
<td>33 (43/31 (40)</td>
</tr>
<tr>
<td>Mean birth stature (range) (cm)</td>
<td>49 (44–52)</td>
<td>39 (30–47)</td>
</tr>
<tr>
<td>Birth head circumference (range) (cm)</td>
<td>34 (32–37)</td>
<td>29 (23–33)</td>
</tr>
<tr>
<td>1-minute Apgar score median (range)</td>
<td>9 (5–7)</td>
<td>7 (1–10)</td>
</tr>
<tr>
<td>5-minute Apgar score median (range)</td>
<td>9 (6–10)</td>
<td>8 (3–10)</td>
</tr>
<tr>
<td>Gestational age vs. birth weight</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Adequate for gestational age (%)</td>
<td>47 (96)</td>
<td>49 (64)</td>
</tr>
<tr>
<td>Small for gestational age (%)</td>
<td>2 (4)</td>
<td>28 (36)</td>
</tr>
<tr>
<td>Mean duration of mechanical ventilation (range)</td>
<td>0 day (0)</td>
<td>4.4 days (0–78)</td>
</tr>
<tr>
<td>Mean duration of neonatal nursery stay (range)</td>
<td>3 days (2–13)</td>
<td>40 days (2–125)</td>
</tr>
<tr>
<td>Number of hospital admission</td>
<td></td>
<td></td>
</tr>
<tr>
<td>None (%)</td>
<td>40 (82)</td>
<td>49 (64)</td>
</tr>
<tr>
<td>Once (%)</td>
<td>8 (16)</td>
<td>27 (35)</td>
</tr>
<tr>
<td>Twice (%)</td>
<td>1 (2)</td>
<td>1 (1)</td>
</tr>
<tr>
<td>Neonatal respiratory distress (%)</td>
<td>63 (82)</td>
<td></td>
</tr>
<tr>
<td>Hyaline membrane disease (%)</td>
<td>20 (26)</td>
<td></td>
</tr>
<tr>
<td>Bronchopulmonary dysplasia (%)</td>
<td>11 (14)</td>
<td></td>
</tr>
<tr>
<td>Intraventricular hemorrhage (grades I and II) (%)</td>
<td>23 (30)</td>
<td></td>
</tr>
<tr>
<td>Retinopathy of prematurity (I and II) (%)</td>
<td>20 (26)</td>
<td></td>
</tr>
</tbody>
</table>
Longitudinal Alberta Infant Motor Scale scores

Fig 1 shows the mean and standard error score of the AIMS curves of preterm versus full-term infants throughout the follow-up period. Both curves were similar in their progression pattern. Full-term infants mean scores were higher than those for preterm infants during follow-up, except in the first month. The greatest difference in scores between groups was observed from the 8th to the 11th months. From the 12th up to the 16th month, the difference gradually decreased. Fig 1 shows that there was less variability in AIMS scores in the first 6 months for both infant groups, while the greatest variability was observed from 7 to 12 months. From this point onwards, the variability tended to decrease as the infants reached the end of the evaluation. Fig 1 shows that the AIMS standard errors overlapped throughout the follow-up period, with the exception of the eighth, ninth, and tenth months of age.

Predictive value of multiple variables on walking attainment

Fig 2 shows Kaplan-Meier function plots of age of walking attainment distribution for preterm and full-term infants. Fig 2 shows that full-term infants were faster in walking attainment than preterm infants. Mean age of onset of walking among full-term was 368.6 days (299–436), whereas among preterm it was 381.6 days (288–470). The difference remained significant when 95% confidence interval age of walking attainment between preterm and full-term infants was estimated (95%CI -1.268–0.162; p<0.05). The percentage of likelihood ratio of full-term infant to walk earlier than preterm is 51% (risk ratio=1.51; 95%CI 1.176–3.554).

Regarding mean age at the first assessment among full-term and preterm infants was respectively, 45 days (21–111) and 75 days (11–164), while mean age of completing all 58

Table 2. Socio-demographic characteristics of full-term and preterm infants.

<table>
<thead>
<tr>
<th>Variables</th>
<th>Full-term (%)</th>
<th>Preterm (%)</th>
<th>p-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total</td>
<td>49 (100)</td>
<td>77 (100)</td>
<td>0.6978</td>
</tr>
<tr>
<td>Male gender</td>
<td>25 (51)</td>
<td>43 (56)</td>
<td>0.6978</td>
</tr>
<tr>
<td>Caucasian</td>
<td>25 (51)</td>
<td>41 (53)</td>
<td>0.9378</td>
</tr>
<tr>
<td>Maternal education</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incomplete elementary grade</td>
<td>2 (4)</td>
<td>25 (32)</td>
<td>0.3223</td>
</tr>
<tr>
<td>Complete elementary grade*</td>
<td>6 (12)</td>
<td>6 (8)</td>
<td></td>
</tr>
<tr>
<td>Incomplete high school</td>
<td>6 (12)</td>
<td>14 (18)</td>
<td>0.8217</td>
</tr>
<tr>
<td>Complete high school</td>
<td>12 (25)</td>
<td>27 (35)</td>
<td></td>
</tr>
<tr>
<td>Incomplete college level</td>
<td>17 (35)</td>
<td>3 (4)</td>
<td></td>
</tr>
<tr>
<td>College level</td>
<td>6 (12)</td>
<td>2 (3)</td>
<td></td>
</tr>
<tr>
<td>Maternal occupation</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Unemployed/housewife**</td>
<td>29 (59)</td>
<td>46 (60)</td>
<td>0.5025</td>
</tr>
<tr>
<td>Unskilled labor</td>
<td>15 (31)</td>
<td>24 (31)</td>
<td></td>
</tr>
<tr>
<td>Technical labor</td>
<td>3 (6)</td>
<td>4 (5)</td>
<td></td>
</tr>
<tr>
<td>Professional</td>
<td>2 (4)</td>
<td>3 (4)</td>
<td></td>
</tr>
<tr>
<td>Socioeconomic status (family income)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>B1-B2 (US$ 532.5–1,471.50 per month)**</td>
<td>4 (8)</td>
<td>4 (6)</td>
<td>0.8484</td>
</tr>
<tr>
<td>C (US$ 248.5–532.00 per month)</td>
<td>27 (55)</td>
<td>35 (45)</td>
<td></td>
</tr>
<tr>
<td>D (US$ 248.00–131.50 per month)</td>
<td>17 (35)</td>
<td>35 (45)</td>
<td></td>
</tr>
<tr>
<td>E (≤ US$ 131.00 per month)</td>
<td>1 (2)</td>
<td>3 (4)</td>
<td></td>
</tr>
</tbody>
</table>

*pReference variable: complete high-school; **Reference variable: Unemployed/housewife; ***Reference variable: C/D/E. p<0.05%.
AIMS items among full-term and preterm infants was respectively, 414.6 days (333–488) and 424.3 days (328–526).

The perinatal independent variables that were statistically significant related to walking attainment delay were: birth weight (p=0.0022), birth length (p=0.022), and duration of neonatal nursery stay (p=0.0465). None of the socioeconomic variables were statistically related to walking attainment delay (p-value in Table 2).

Univariate proportional hazard regression analysis of perinatal risk factors for late walking attainment among preterm infants showed that a 100 g increment in birth weight increases by 11% the likelihood ratio of walking attainment (risk ratio=1.108; 95%CI 1.044–1.170) and an 1 cm increment to birth length increases in 12% the likelihood ratio of walking attainment (risk ratio=1.116; 95%CI 1.033–1.207). A decrease of 10 days of neonatal nursery stay increases by 14% the likelihood ratio of walking attainment (risk ratio=1.137; 95%CI 1.025–1.260).

**DISCUSSION**

We identified differences in early gross motor development between low risk preterm infants compared with normal full-term infants. There was no statistical difference between preterm and full-term infant AIMS mean scores, with the exception of the eighth, ninth and tenth month periods.

Comparison of the AIMS scores between the preterm and full-term groups was previously reported\(^{11,22-26}\), however the methodologies were distinct among studies, jeopardizing data comparison.

Formiga and Linhares\(^{24}\) found that the preterm infants presented greater delay in the AIMS mean scores, especially from 9 to 12 months, while van Haastert et al.\(^{22}\) reported significantly lower mean scores at all age levels. In a previous study\(^{25}\), we did not find statistical differences in the AIMS mean scores between the preterm infant group and the AIMS normative data, during the first year of life.

Regarding the study design, we followed our infants longitudinally from the first hospital visit until all infants had completed the 58 AIMS items and had undergone a comprehensive neurological examination. As we intended to draw the typical motor development pattern of the preterm infants over time, we chose to test both infant groups monthly, rather than doing it twice or at every three months up to 18 months\(^{22}\). Interestingly, we found differences between the groups only at eight, nine, and ten months, when preterm infant mean scores were statistically lower than full-term infants. A longitudinal study of infant performance on the AIMS revealed that measurements (raw scores, percentile ranks) varied from age to age with no systematic pattern within individual infants\(^{27}\). Therefore, this finding may suggest that preterm infants have a different pattern of score progression than full-term infants, which could be a biological variability between preterm and full-term infants. Alternatively, it is possible that the period between the 8th and 12th months is of special importance for early gross motor development due to the greatest score increase and the greatest difference in mean scores between our comparison groups. Darrah et al.\(^{27}\) observed that the rate of motor development in normal developing infants is characterized by within-subject variability and non-linearity of skill emergence. There are peaks, valleys, and plateaus in the neuro-developmental trajectories of standardized measures to provide information about child skill development, which indicates periods of acceleration, deceleration, or quiescence of activity.

Preterm infant motor trajectories resemble those of full-term infants, as characterized by gradual increase in the AIMS scores with advancing age, alternating with periods of more or less quiescence. Both of our groups exhibited the greatest score increase between the 8th and 12th months of age. In the first three months and between the 12th and 16th months, the score slope of preterm infants showed a slow increase with a tendency of quiescence, characterized by attainment of few new motor skills. The small score increase or variation is limited towards the two ends of the measurement (in the first trimester and from 12 to 16 months), whereas the period between 8 and 12 months shows greater score variation. During this period, the AIMS items aggregate at the middle range of difficulty level, which suggests that the AIMS is sufficiently precise to discriminate among infants whose ability levels are in the middle range, but not at the lowest nor at the highest ability ends. Small score variation in the first three months may be due to the fact that there are only a few items in prone and supine positions that are commonly observed in infants at this age. However, after walking attainment (between 12 and 16 months), there are only standing items available to complete the 58 items\(^{28}\).

Alternatively, this variability in the AIMS scores in these specific months may be related to limitations of the scaling of the AIMS items, as there is a discontinuity in item difficulty\(^{28}\). There are gaps at several difficulty levels, which indicate that a large jump in ability level is required to pass one or more items around the gap.

We observed delay of approximately one month in walking attainment among preterm infants (the mean age of onset walking in preterm: 381.6 days; and in full-term infants: 368.6 days), which agreed with some studies\(^{8,13,14}\). The age of onset of walking of our full-term infants was close to that reported by Allen and Alexander\(^{8}\). Some studies\(^{8,12}\) suggested that the age of walking attainment is related to the rate of maturation and development of nervous and musculoskeletal systems and therefore it may have some predictive value as a later neuromotor outcome measure. It is suggested for future research that the late walkers should be followed up to 4-5 years.
to verify the predictive value in other areas of the neuro-
motor development. Johnson, Goddard and Ashurst\textsuperscript{12} ob-
served that more than half of the late walkers had an asso-
ciated neurological abnormality diagnosed before the age
of 3, whereas 15\% had no neurological dysfunction. It is
a matter of speculation whether delay in walking among
extremely low birth weight children can indicate subtle
damage to motor control functions seen in “soft” or MND.
Bartlett et al.\textsuperscript{29} reported that a significant proportion
of children born preterm without major neurological con-
dition were delayed in the acquisition of antigravity postur-
al control and motor development, such as walking mile-
stone. We found that birth weight and length and duration
of nursery stay in the neonatal period had strong associ-
ation with delay in preterm walking attainment.

Our results agree with studies\textsuperscript{10,11,30} that found a negative
relation between late attainment of walking and extreme
prematurity, small-for-gestational-age, and very-low birth
weight infants.

We conclude that preterm infants walk later than full-
term infants, and that very low birth weight and length, and
duration of neonatal nursery stay were predictive of late
walking in preterm infants.

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