Genetic and environmental variance and covariance parameters for some reproductive traits of Holstein and Jersey cattle in Antioquia (Colombia)

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ABSTRACT - The objective of this study was to estimate the genetic, phenotypic and environmental parameters for calving interval (CI), days open (DO), number of services per conception (NSC) and conception rate (CR) in Holstein and Jersey cattle in Antioquia (Colombia). Variance and covariance component estimates were obtained by an animal model that was solved using the derivative-free restricted maximum likelihood method. The means and standard deviations for CI, DO, NSC and CR were: 430.32±77.93 days, 127.15±76.96 days, 1.58±1.03 services per conception and 79.88±28.66% in Holstein cattle, and 409.33±86.48 days, 125.62±86.09 days, 1.48±0.98 services per conception and 84.08±27.23% in Jersey cattle, respectively. The heritability estimates (standard errors) were: 0.088(0.037), 0.082(0.037), 0.040(0.025) and 0.030(0.026) in Holstein cattle and 0.072(0.098), 0.090(0.104), 0.093(0.097) and 0.147(0.117) in Jersey cattle, respectively. The results show that the genetic, phenotypic and permanent environmental correlations in the two evaluated breeds were favorable for CI × DO, CI × NSC and DO × NSC, but not for CI × CR, DO × CR and NSC × CR. Genetic and permanent environmental correlations were high in most cases in Holstein cattle, whereas in Jersey cattle they were moderate. In contrast, phenotypic correlations were very low in both breeds, except for CI × DO and NSC × CR, which were high. Overall, the genetic component found was very low (≤8%) in both evaluated breeds and this implies that their selection would take long time and that a good practical management of the herd will be essential in order to improve the reproductive performance.

Key Words: fertility, genetic correlation, heritability, repeatability

Introduction

Fertility traits are considered very important because of their impact on the economy of dairy cattle breeding. Economic losses due to fertility problems are mainly caused by low dairy yield, prolonged calving intervals, increased insemination costs, few calves per cow per year, increased culling, high replacement costs and shorter reproductive lifespans (Abe et al., 2009; González-Recio and Alenda, 2005). Nonetheless, for many years, genetic improvement programs worldwide did not include reproductive performance, since the selection was mainly focused on milk yield. An exception was Scandinavia, whose selection indices included not only milk yield, but also health and reproductive traits (Miglior et al., 2005). Now, the increase in milk yield without considering the reproductive performance is a problem, because it produced an important decline in the reproductive efficiency over time (Pryce et al., 2004; Melendez and Pinedo, 2007).

It is difficult to determine which traits must be included in the genetic evaluation of fertility, since they have very low heritability values, i.e., close to 0.1 (Thaller, 1997; Jamrozik et al., 2005). However, over the last decade, reproductive traits have increasingly been included in the selection indices for reproductive traits in genetic evaluations in different countries, thus highlighting the importance of including fertility in improvement programs of dairy cattle (Miglior et al., 2005). Nevertheless, genetic improvement of reproductive traits is almost non-existent in Colombia. Even though some improvement programs for increasing dairy yield have been implemented, they have not survived due to the lack of financial resources. As a consequence, there is currently a lack of information on all the genetic components involved, and specifically on reproductive variables such as calving interval, age at first service, age at first calving, and days open, among others (García et al., 2002; Grajales et al., 2006). The existing literature is therefore rather limited. Further research on how to increase dairy yield, reproduction and health of dairy cattle is required.

The objective of this study was to estimate the genetic, phenotypic and environmental parameters for calving interval (CI), days open (DO), number of services per conception (NSC) and conception rate (CR) in Holstein and Jersey cattle in Antioquia (Colombia).
Material and Methods

This research was approved by the Ethics Committee of Universidad Nacional de Colombia sede Medellín, considering that it meets the standards for this type of research (Approval letter number: CEMED-015 May 2012).

The data for this study was obtained from 88 Holstein and Jersey dairy herds in 18 municipalities of Antioquia (Colombia). Specific health and feeding management conditions were different for each dairy farm, as well as topography and geographical location.

All herd information regarding breeding or health — calving dates, dry dates, pregnancy detection dates, insemination dates, cow entry/exit, etc — was obtained from historical records, notebooks and cattle management software. This information was entered, analyzed and saved using the Control 1 Software, version 1.0 of the Animal Genetic Improvement Laboratory at Universidad Nacional de Colombia sede Medellín. Data selection was based upon reliability: information whose validity could not be guaranteed was deleted from the final data set before analysis, as were extreme values that were considered physiologically abnormal or erroneously coded.

The reproductive traits evaluated in this study were: calving interval (CI), measured as the number of days between two consecutive calving; days open (DO), measured as the days between calving and conception; number of services per conception (NSC), defined as the number of inseminations until pregnancy; and conception rate (CR), defined as the percentage of successful inseminations [CR = (1/NSC)*100]. The number of Holstein animals was 7,937, of which 275 were bulls. For Jersey cattle, there were a total of 833 animals, including 56 bulls.

To (co)variance components were estimated by Derivative-free Restricted Maximum Likelihood Method using the MTDFREML software (Boldman et al., 1995), which estimates fixed and random effect solutions by solving the mixed models equations (MME), described by Henderson (1984).

Variance components for CI, DO, NSC and CR were estimated through univariate analysis using an animal model considering the effects of herd, number of calving and contemporary group as fixed, and the permanent environmental, animal additive genetic and residual effects as random. The contemporary group included municipality, year of calving and calving season. The model used (Mrode, 1996) can be described as:

\[ y = Xb + Wpe + Za + e \]

in which \( y \) = vector of observations; \( b \) = vector of fixed effects (herd, number of calving, and contemporary group); \( pe \) = vector of random permanent environmental effects; \( a \) = vector of random animal effects; \( e \) = vector of random residual effects; \( X, W, \) and \( Z \) = incidence matrices that establish relationships between the records and the effects.

It is assumed that permanent environmental, animal, and residual effects are independently distributed with mean zero and variance:

\[
\begin{bmatrix}
V \\
pe \\
e
\end{bmatrix} = 
\begin{bmatrix}
A\sigma_a^2 & 0 & 0 \\
0 & I\sigma_{pe}^2 & 0 \\
0 & 0 & I\sigma_e^2
\end{bmatrix}
\]

Considering that \( A = \) relationship matrix, \( I\sigma_e^2 = R \), then \( \text{Var}(y) = \text{Var}(Z'\sigma_a^2 + W\sigma_{pe}^2 + W'R + R) \). Thus, the mixed model equations for the best linear unbiased estimator (BLUE) of estimable functions of \( b \) and for the best linear unbiased prediction (BLUP) of \( pe \) and \( a \) are:

\[
\begin{bmatrix}
b \\
pe \\
a
\end{bmatrix} = 
\begin{bmatrix}
X'X & X'Z & X'W \\
Z'X & Z'Z + A_{a1} & Z'W \\
W'X & W'Z & W'W + I_{a2}
\end{bmatrix}^{-1} 
\begin{bmatrix}
X'y \\
Z'y \\
W'y
\end{bmatrix}
\]

in which \( A = \) relationship matrix; and \( a_1 = \frac{\sigma_a^2}{\sigma_e^2} \) and \( a_2 = \frac{\sigma_{pe}^2}{\sigma_e^2} \).

Heritability was estimated as the ratio of the additive genetic variance to total phenotypic variance; and repeatability, as the ratio of the sum of the additive genetic variance plus permanent environmental variance to phenotypic variance, as described by Falconer and Mackay (2001):

\[
h^2 = \frac{\sigma_a^2}{\sigma_a^2 + \sigma_{pe}^2 + \sigma_e^2}; \\
R = \frac{\sigma_a^2 + \sigma_{pe}^2}{\sigma_a^2 + \sigma_{pe}^2 + \sigma_e^2}.
\]

To estimate genetic, phenotypic, residual and permanent environmental correlations, a bivariate model was used including herd, number of calving and contemporary group (which included municipality, year of calving and calving season) as the fixed effects, and the permanent environmental and additive genetic direct effects as random. The matrix model used was:

\[
\begin{bmatrix}
y_1 \\
y_2
\end{bmatrix} = 
\begin{bmatrix}
X_1 & 0 & h_1 & Z_1 & 0 & a_1 \\
0 & X_2 & 0 & Z_2 & a_2 \\
W_1 & 0 & pe_1 \\
0 & W_2 & pe_2
\end{bmatrix} + 
\begin{bmatrix}
e_1 \\
e_2
\end{bmatrix}
\]

in which \( y \) = vector of N observations; \( h_1 = \) vector of fixed effects (herd, number of calving, contemporary group); \( pe_1 = \) vector of random permanent environmental effects; \( a_1 = \) vector of random animal effects; \( e_i = \) vector of random residual effects; \( X, W, \) and \( Z \) = incidence matrices establishing relationships between the records and the fixed and random effects, respectively.

It is assumed that random permanent environmental, animal and error effects are independently distributed with mean of zero and variance:

\[
\begin{bmatrix}
A\sigma_a^2 & 0 & 0 \\
0 & I\sigma_{pe}^2 & 0 \\
0 & 0 & I\sigma_e^2
\end{bmatrix}
\]
Genetic and environmental variance and covariance parameters for some reproductive traits of Holstein and Jersey cattle in Antioquia...

\[
V = \begin{bmatrix} a \\ \phantom{a} pe \\ e \end{bmatrix} = \begin{bmatrix} G_0 \otimes A & 0 & 0 \\ 0 & Q_0 \otimes I & 0 \\ 0 & 0 & R_0 \otimes I \end{bmatrix}
\]

\[
G_0 = \begin{bmatrix} \sigma^2_{ai} & \sigma_{aij} \\ \sigma_{aij} & \sigma^2_{aj} \end{bmatrix}, \quad Q_0 = \begin{bmatrix} \sigma^2_{peij} & \sigma_{peijj} \\ \sigma_{peijj} & \sigma^2_{pejj} \end{bmatrix}, \quad R_0 = \begin{bmatrix} \sigma^2_{ei} & \sigma_{eiij} \\ \sigma_{eiij} & \sigma^2_{eij} \end{bmatrix}
\]

in which \( \otimes \) = direct or Kronecker product; \( I \) = identity matrix equal to number of observations; \( A \) = relationship matrix among all animals in the pedigree; \( G_0 \) = variance and covariance matrix of random animal effects; \( \sigma^2_{ai} \) = animal additive genetic variance for trait \( i \); \( \sigma^2_{aj} \) = animal additive genetic variance for trait \( j \); \( \sigma_{aij} \) = animal additive genetic covariance between traits \( i \) and \( j \); \( Q_0 \) = variance and covariance matrix of random permanent environmental effects; \( \sigma^2_{peij} \) = permanent environmental variance for trait \( i \); \( \sigma^2_{pejj} \) = permanent environmental variance for trait \( j \); \( \sigma_{peijj} \) = permanent environmental covariance between traits \( i \) and \( j \); \( R_0 \) = variance and covariance matrix of residual effects; \( \sigma^2_{ei} \) = residual variance for trait \( i \); \( \sigma^2_{eij} \) = residual variance for trait \( j \); and \( \sigma_{eiij} \) = residual covariance between traits \( i \) and \( j \).

The mixed model equations for the best linear unbiased estimator (BLUE) of estimable functions of \( a \) and for the best linear unbiased prediction (BLUP) of \( a \) and \( ap \) are:

\[
\begin{bmatrix} \hat{b} \\ \hat{\alpha} \\ \hat{\rho}_e \end{bmatrix} = \begin{bmatrix} X^T R^{-1} X & X^T R^{-1} Z & X^T R^{-1} W \\ Z^T R^{-1} X & Z^T R^{-1} Z + k_i & Z^T R^{-1} W \\ W^T R^{-1} X & W^T R^{-1} Z & W^T R^{-1} W + I \cdot Q^{-1} \end{bmatrix}^{-1} \begin{bmatrix} X^T R^{-1} y \\ Z^T R^{-1} y \\ W^T R^{-1} y \end{bmatrix}
\]

in which \( k_i = G_0 \ast A^{-1} \).

The estimates of genetic (\( r_g \)), permanent environmental (\( r_p \)), residual (\( r_r \)), and phenotypic (\( r_y \)) correlations were obtained from the estimation of covariance components using the following equations:

\[
r_g = \frac{\sigma_{aij}}{\sqrt{\sigma^2_{ai} \sigma^2_{aj}}}, \quad r_p = \frac{\sigma_{peijj}}{\sqrt{\sigma^2_{peij} \sigma^2_{pejj}}}, \quad r_r = \frac{\sigma_{eij}}{\sqrt{\sigma^2_{eij} \sigma^2_{eij}}}, \quad r_y = \frac{\sigma_{peijj}}{\sqrt{\sigma^2_{peij} \sigma^2_{pejj}}}
\]

in which \( \sigma_{peijj} \) = phenotypic covariance between traits \( i \) and \( j \); \( \sigma^2_{peij} \) = genetic variance for trait \( i \); \( \sigma^2_{pejj} \) = genetic variance for trait \( j \); \( \sigma_{eij} \) = environmental variance for trait \( j \); \( \sigma^2_{eij} \) = environmental covariance between traits \( i \) and \( j \); and \( \sigma^2_{eij} \) = phenotypic variance for trait \( j \).

### Results

Regarding the statistical summary for the four evaluated traits CI (days), DO (days), NSC and CR (%) in the Holstein and Jersey breeds (Table 1), the means for Holstein cattle were: 410.32±77.93 days, 127.5±76.96 days, 1.58±1.03 and 79.88±28.66%, respectively. Very similar results were obtained for Jersey cattle: 409.33±86.48 days, 125.62±86.09 days, 1.48±0.98 and 84.08±27.23%. The traits with the highest variability for both breeds according to the variation coefficient were DO and NSC (Holstein: 60.53, 65.07% and Jersey: 68.53, 66.00%) and those with the lowest variation, CI and CR (Holstein: 18.99, 35.88% and Jersey: 21.13, 32.38%).

The heritability and repeatability estimates for Holstein and Jersey cattle (Table 2) were lower than 0.15, and the heritability estimates for CI, DO, NSC and CR were: 0.088±0.037, 0.082±0.037, 0.040±0.025 and 0.030±0.026 for the Holstein and 0.072±0.098, 0.090±0.104, 0.093±0.097 and 0.147±0.117 for the Jersey breeds, respectively. The heritability values for CI and DO were similar for the two breeds, but higher for NSC and CR in Jersey cattle.

The repeatability estimates for CI, DO, NSC and CR were: 0.093±0.037, 0.101±0.037, 0.069±0.026 and 0.076±0.027 for Holstein cattle, and 0.074±0.116, 0.091±0.123, 0.094±0.105 and 0.147±0.121 for Jersey cattle. The repeatability values were similar for the two breeds for CI, DO and NSC, but they were higher for CR of Jersey cattle.

Permanent environmental variance was lower than genetic variance in both breeds in all cases (Table 3). Permanent environmental variance was higher for Holstein than Jersey cattle, since the number of studied lactations was lower in the latter. It should be noted that all the variances change for one trait depending on the trait with which it is correlated, because the number of observations of the two correlated traits must be balanced to calculate the standard errors of their genetic parameters.

According to the results obtained in this study, genetic, permanent environmental and phenotypic correlations can be classified into two groups for both breeds: group 1, or positive correlations; CI × DO, CI × NSC and DO × NSC, and group 2, or negative correlations: CI × CR, DO × CR and NSC × CR (Table 4). Genetic correlations were high in both groups (group 1: 0.980, 0.955 and 0.988; group 2: −1.00, −0.993 and −1.00) for Holstein cattle, respectively, whereas for Jersey cattle they were high in some cases and moderate in others (group 1: 0.50, 0.541 and 0.49; group 2: −0.250, −0.250 and −0.890), respectively. Permanent environmental correlations showed a similar pattern (group 1: 0.941, 0.512 and 0.576; group 2: −0.988, −0.810 and −0.978 for Holstein cattle, but for Jersey cattle they were relatively lower in some cases: group 1: 0.951, 0.261 and 0.260; group 2: −0.710, −0.490 and −0.810). Phenotypic correlations were very low in both breeds, with the exceptions of CI × DO and NSC × CR (group 1: 1.00, 0.047 and 0.050; group 2: −0.047, −0.053, −0.880 for Holstein cattle, whereas for Jersey cattle they were: group 1: 1.00,
Finally, residual correlations were very low in both breeds, with the exceptions of CI × DO and NSC × CR (group 1: 1.00, –0.036 and –0.039; group 2: 0.014, 0.040 and –0.877 for Holstein cattle; and for Jersey cattle, group 1: 1.00, –0.005 and 0.001; group 2: 0.001, –0.005 and –0.930).

Discussion

The results obtained for the mean of each trait are consistent with the results of previous studies, although some authors have reported different values. The mean values for CI obtained in this study for Holstein and Jersey

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**Table 1** - Descriptive statistical analysis for calving interval (CI) days open (DO), number of services per conception (NSC) and conception rate (CR) of Holstein and Jersey cattle in Antioquia, Colombia

<table>
<thead>
<tr>
<th></th>
<th>Holstein</th>
<th></th>
<th>Jersey</th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>CI</td>
<td>DO</td>
<td>NSC</td>
<td>CR</td>
</tr>
<tr>
<td>Mean</td>
<td>410.3</td>
<td>127.15</td>
<td>1.58</td>
<td>79.88</td>
</tr>
<tr>
<td></td>
<td>Standard deviation</td>
<td>77.93</td>
<td>76.96</td>
<td>1.03</td>
</tr>
<tr>
<td></td>
<td>Coefficient of variation</td>
<td>18.99</td>
<td>60.53</td>
<td>1.0</td>
</tr>
<tr>
<td></td>
<td>Minimum</td>
<td>250</td>
<td>1</td>
<td>1.0</td>
</tr>
<tr>
<td></td>
<td>Maximum</td>
<td>700</td>
<td>415</td>
<td>7</td>
</tr>
<tr>
<td></td>
<td>Number of lactations</td>
<td>10,303</td>
<td>10,156</td>
<td>723</td>
</tr>
</tbody>
</table>

**Table 2** - Variance components and genetic parameters for calving interval (CI) days open (DO), number of services per conception (NSC) and conception rate (CR) in Holstein and Jersey breeds in Antioquia, Colombia

<table>
<thead>
<tr>
<th>Breed</th>
<th>Trait</th>
<th>V_a (SE)</th>
<th>V_pe (SE)</th>
<th>V_e (SE)</th>
<th>V_2 (SE)</th>
<th>R (SE)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Holstein</td>
<td>CI</td>
<td>516.0(0.037)</td>
<td>5.265.0(0.037)</td>
<td>5.503.0(0.037)</td>
<td>5.850.0(0.037)</td>
<td>0.088(0.037)</td>
</tr>
<tr>
<td></td>
<td>DO</td>
<td>469.0(0.037)</td>
<td>5265.0(0.037)</td>
<td>5.126.0(0.037)</td>
<td>5.700.0(0.037)</td>
<td>0.094(0.037)</td>
</tr>
<tr>
<td></td>
<td>NSC</td>
<td>0.03 (0.025)</td>
<td>0.028 (0.025)</td>
<td>0.088 (0.025)</td>
<td>0.040 (0.025)</td>
<td>0.04 (0.025)</td>
</tr>
<tr>
<td></td>
<td>CR</td>
<td>22.0 (0.025)</td>
<td>33.0 (0.025)</td>
<td>665.0 (0.025)</td>
<td>719.0 (0.025)</td>
<td>0.97 (0.025)</td>
</tr>
<tr>
<td>Jersey</td>
<td>CI</td>
<td>547.0 (0.116)</td>
<td>5034.0 (0.116)</td>
<td>5574.0 (0.116)</td>
<td>5531.0 (0.116)</td>
<td>0.093 (0.116)</td>
</tr>
<tr>
<td></td>
<td>DO</td>
<td>678.0 (0.116)</td>
<td>4895.0 (0.116)</td>
<td>5126.0 (0.116)</td>
<td>5541.0 (0.116)</td>
<td>0.147 (0.116)</td>
</tr>
<tr>
<td></td>
<td>NSC</td>
<td>0.087 (0.105)</td>
<td>0.001 (0.105)</td>
<td>0.084 (0.105)</td>
<td>0.093 (0.105)</td>
<td>0.094 (0.105)</td>
</tr>
</tbody>
</table>

V_a - additive genetic variance; V_pe - permanent environmental variance; V_e - residual variance; V_2 - phenotypic variance; h^2 - heritability; R - repeatability; SE - standard error.
Table 4 - Genetic, permanent environmental, phenotypic, and residual correlations for CI × DO, CI × NSC, CI × CR, DO × NSC, DO × CR and NSC × CR in Holstein and Jersey cattle in Antioquia, Colombia

<table>
<thead>
<tr>
<th>Breed</th>
<th>Correlations</th>
<th></th>
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</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>CI × DO</td>
<td>CI × NSC</td>
<td>CI × CR</td>
<td>DO × NSC</td>
</tr>
<tr>
<td>Holstein</td>
<td></td>
<td>r_g (SE)</td>
<td>r_pe (SE)</td>
<td>r_g (SE)</td>
<td>r_pe (SE)</td>
</tr>
<tr>
<td></td>
<td></td>
<td>0.980(0.243)</td>
<td>0.955(0.450)</td>
<td>-1.000(0.529)</td>
<td>0.988(0.449)</td>
</tr>
<tr>
<td></td>
<td></td>
<td>0.941(1.096)</td>
<td>0.512(1.325)</td>
<td>-0.988(0.010)</td>
<td>0.576(0.656)</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>-0.993(0.532)</td>
<td></td>
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<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>-1.000(0.123)</td>
<td></td>
</tr>
<tr>
<td>Jersey</td>
<td></td>
<td>1.000(0.002)</td>
<td>0.541(1.357)</td>
<td>-0.250(1.300)</td>
<td>0.490(1.109)</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>-0.710(0.569)</td>
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<td></td>
<td></td>
<td></td>
<td>-0.490(0.905)</td>
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<td></td>
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<td></td>
<td></td>
<td>-0.890(0.220)</td>
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(0.101, 0.091), are lower than, yet close to those reported by M’hamedi et al. (2010), Estrada-León et al. (2008) and Demeke et al. (2004), which were in a very close range (0.135-0.190). For NSC, the repeatability estimates were 0.069 and 0.094, similar to those reported by Estrada-León et al. (2008) and Demeke et al. (2004), but higher than those reported by M’hamedi et al. (2010), Biffani et al. (2005) and Kadarmideen et al. (2000), which were in a range between 0.022 and 0.037. Finally, for CR, the repeatability values found for each breed (0.076, 0.174) were higher than those reported by Ríos-Ultrera et al. (2010b), which were 0.03 in Brown Swiss cattle.

The correlations between CI and DO in this study for Holstein and Jersey cattle were 0.98 and 1.00, respectively. These values are similar to those reported by González-Recio and Alenda (2005), Gredler et al. (2007), and Ghiasi et al. (2011), who reported a nearly perfect genetic correlation (0.99, 0.98, and 0.99, respectively). These results suggest that these two reproductive traits are genetically equivalent, i.e., they are influenced by the same genes. This is known as pleiotropic effect (Falconer and Mackay, 2001).

The joint analysis of CI and NSC indicates that the genetic association between these two traits is positive and high, according to the results obtained for Holstein cattle (0.955), consistent with those reported by Ríos-Ultrera et al. (2010c) and González-Recio and Alenda (2005) (0.97 and 0.89, respectively). For Jersey cattle, a genetic correlation between CI and NSC of 0.541 was estimated, corroborating the values obtained by Biffani et al. (2005), Ghiasi et al. (2011) and Wall et al. (2003) for Holstein cattle (0.61, 0.69 and 0.61, respectively). However, lower values were reported by Kadarmideen et al. (2000) for Holstein cattle (0.22) and by Gredler et al. (2007) for Simmental cattle (0.41).

Likewise, the genetic correlation between DO and NSC was high and favorable for Holstein cattle (0.988), and medium and favorable (0.490) for Jersey cattle. Equivalent results were reported by González-Recio and Alenda (2005) and Ghiasi et al. (2011): 0.94 and 0.72, respectively. Gredler et al. (2007) reported a medium genetic correlation (0.42), equivalent to the correlation found in this study for Jersey cattle. These results suggest the notorious presence of pleiotropic effects between these traits.

The genetic correlation between CI and CR was −1.00 for Holstein cattle. Considering that the pregnancy rate is closely related to the conception rate, the result obtained in this study is similar to that obtained by González-Recio and Alenda (2005) and Ghiasi et al. (2011), who both reported a correlation of −0.99 between CI and pregnancy rate. For Jersey cattle, the genetic correlation between CI and CR that we obtained was medium (−0.447), similar to the findings obtained by Haile-Mariam et al. (2003), who reported a correlation of −0.57 between CI and pregnancy rate.

Similarly, the genetic correlation obtained in this study between DO and CR was high and negative for Holstein cattle (−0.993); this is consistent with the results reported by González-Recio and Alenda (2005), Ghiasi et al. (2011) and Abe et al. (2009) (−0.99, −0.99 and −0.76, respectively). For Jersey cattle, in contrast, we obtained a moderately low genetic correlation (−0.244).

Table 5 - Genetic parameters reported by several authors for calving interval (CI), days open (DO), number of services per conception (NSC) and conception rate (CR)

<table>
<thead>
<tr>
<th>Authors</th>
<th>Breed</th>
<th>CI h^2</th>
<th>CI R</th>
<th>DO h^2</th>
<th>DO R</th>
<th>NSC h^2</th>
<th>NSC R</th>
<th>CR h^2</th>
<th>CR R</th>
</tr>
</thead>
<tbody>
<tr>
<td>Ghiasi et al. (2011)</td>
<td>H</td>
<td>0.074</td>
<td>-</td>
<td>0.076</td>
<td>-</td>
<td>0.046</td>
<td>-</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>Pantelic et al. (2011)</td>
<td>H</td>
<td>-</td>
<td>-</td>
<td>0.105</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>M’hamedi et al. (2010)</td>
<td>H</td>
<td>0.063</td>
<td>0.152</td>
<td>0.041</td>
<td>0.135</td>
<td>0.027</td>
<td>0.034</td>
<td>-</td>
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</tr>
<tr>
<td>Ríos-Ultrera et al. (2010a)</td>
<td>H</td>
<td>0.130</td>
<td>0.130</td>
<td>0.010</td>
<td>0.160</td>
<td>0.030</td>
<td>-</td>
<td>-</td>
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</tr>
<tr>
<td>Ríos-Ultrera et al. (2010b)</td>
<td>BS</td>
<td>0.030</td>
<td>0.120</td>
<td>0.000</td>
<td>0.140</td>
<td>-</td>
<td>-</td>
<td>0.049</td>
<td>-</td>
</tr>
<tr>
<td>Abe et al. (2009)</td>
<td>H</td>
<td>0.074</td>
<td>-</td>
<td>0.074</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>Sun et al. (2010)</td>
<td>H</td>
<td>0.067</td>
<td>-</td>
<td>0.067</td>
<td>-</td>
<td>0.028</td>
<td>-</td>
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<tr>
<td>Tsruta et al. (2009)</td>
<td>H</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>0.052</td>
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<tr>
<td>Restrepo et al. (2008)</td>
<td>H</td>
<td>0.090</td>
<td>0.180</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
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<tr>
<td>Estrada-León et al. (2008)</td>
<td>BS</td>
<td>0.110</td>
<td>0.180</td>
<td>0.057</td>
<td>0.190</td>
<td>0.040</td>
<td>0.070</td>
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<tr>
<td>Gredler et al. (2007)</td>
<td>S</td>
<td>0.035</td>
<td>-</td>
<td>0.040</td>
<td>-</td>
<td>0.022</td>
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<tr>
<td>Biffani et al. (2005)</td>
<td>H</td>
<td>0.065</td>
<td>0.154</td>
<td>-</td>
<td>-</td>
<td>0.027</td>
<td>0.037</td>
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<tr>
<td>Sven et al. (2005)</td>
<td>H</td>
<td>0.020</td>
<td>-</td>
<td>0.030</td>
<td>-</td>
<td>0.010</td>
<td>-</td>
<td>-</td>
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</tr>
<tr>
<td>González-Recio and Alenda (2005)</td>
<td>H</td>
<td>0.040</td>
<td>-</td>
<td>0.040</td>
<td>-</td>
<td>0.020</td>
<td>-</td>
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</tr>
<tr>
<td>Demeke et al. (2004)</td>
<td>H</td>
<td>0.080</td>
<td>0.140</td>
<td>0.040</td>
<td>0.140</td>
<td>0.070</td>
<td>0.080</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>Haile-Mariam et al. (2003)</td>
<td>H</td>
<td>0.09</td>
<td>-</td>
<td>0.040</td>
<td>-</td>
<td>0.030</td>
<td>-</td>
<td>0.05</td>
<td>-</td>
</tr>
<tr>
<td>Wall et al. (2003)</td>
<td>H</td>
<td>0.033</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>0.020</td>
<td>-</td>
<td>-</td>
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</tr>
<tr>
<td>Veerkamp et al. (2001)</td>
<td>H</td>
<td>0.032</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>0.034</td>
<td>-</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>Ojango and Pollott (2001)</td>
<td>H</td>
<td>0.047</td>
<td>0.060</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>-</td>
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<tr>
<td>Kadarmideen et al. (2000)</td>
<td>H</td>
<td>0.022</td>
<td>0.049</td>
<td>-</td>
<td>0.010</td>
<td>0.032</td>
<td>-</td>
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</tr>
</tbody>
</table>

H - Holstein; J - Jersey; BS - Brown Swiss; S - Simmental; h^2 - heritability; R - repeatability.
Likewise, the genetic correlation between NSC and CR obtained in this study was high and negative for both Holstein (−0.877) and Jersey cattle (−0.972), similar to the results reported by González-Recio and Alenda (2005) and Ghiasi et al. (2011) for the genetic correlation between NSC and pregnancy rate (−0.94 and −0.73, respectively). These results suggest that these two traits, CI and CR, are not genetically favored, as could be logically expected.

The phenotypic correlation between CI and DO in this study was perfect and favorable, with values of 1.00 for both breeds. González-Recio and Alenda (2005) and Ghiasi et al. (2011) reported similar estimates (0.91 and 0.95, respectively) for Holstein cattle. In contrast, the phenotypic correlation between CI and NSC was very low both for both Holstein (0.047) and Jersey (0.041) cattle. These results are consistent with those obtained by Ageeb and Hayes (2000) and El Amin et al. (1986) (0.05 and 0.01, respectively). Nevertheless, Ghiasi et al. (2011), González-Recio and Alenda (2005), and Kadarmideen et al. (2000) report higher values (0.70, 0.68, and 0.69, respectively). Similarly, the results obtained in this study regarding the phenotypic correlation between DO and NSC are very low (0.050) for both breeds. These results are very similar to those reported by El Amin et al. (1986) (0.01). However, Ríos-Ultrera et al. (2010c) and Ghiasi et al. (2011) report higher phenotypic correlations (0.56 and 0.73).

The differing results of previous studies may be due to the fact that in our study there was a high number of cows with high days open and calving intervals, and low services per conception because of different factors: failure to record unsuccessful inseminations, low heat detection efficiency, or cows not served due to reproductive problems or, in some cases, to low body condition. Additionally, environmental conditions might have influenced phenotypic correlations: this study, as well as those which obtained low phenotypic correlations between DO and NSC, and between CI and NSC, was carried out in countries with an equatorial climate, unlike most of the studies where higher correlations between these traits were obtained, which were carried out in zones with a subtropical climate.

The phenotypic correlations between NSC and CR obtained in this study were high and negative (−0.88 and −0.92 for Holstein and Jersey cattle, respectively), and they were equivalent to the results obtained by González-Recio and Alenda (2005) and Ghiasi et al. (2011) (−0.75 and −0.73, respectively). In contrast, the phenotypic correlations between CI and CR were −0.047 for Holstein and −0.030 for Jersey cattle, unlike the higher values reported by Ghiasi et al. (2011) and González-Recio and Alenda (2005) (−0.95 and −0.91, respectively). In relation to the phenotypic correlation between DO and CR, the values obtained in this study were −0.053 for Holstein cattle and −0.032 for Jersey cattle, which is consistent with the result reported by Abe et al. (2009) (−0.07), which indicates that the degree of phenotypic association between these two traits is low and negative. The differing results may also be explained by the factors discussed above.

The residual correlations between CI and DO obtained in the present study were favorable and highly associated, presenting 1.00 in the two breeds studied. These results are similar to the findings of Ríos-Ultrera et al. (2010c), who reported 0.90. Similarly, residual correlations between NSC and CR were negative and highly associated with values of −0.877 and −0.930 in Holstein and Jersey cattle, respectively. Ríos-Ultrera et al. (2010c) and Kadarmideen et al. (2000) reported residual correlation values which were moderately high and negative (−0.71 and −0.7829) and corroborate the findings of this study.

As for permanent environmental correlations between CI and NSC, the values obtained were medium and positively associated (0.512 and 0.261 for Holstein and Jersey cattle, respectively). Biffani et al. (2005) and Kadarmideen et al. (2000) reported similar results (0.512 and 0.2502, respectively) in Holstein cattle. The permanent environmental correlation estimates obtained in this study between CI and CR were also high yet negative in both Holstein and Jersey cattle (−0.988 and −0.701, respectively). Kadarmideen et al. (2000) reported a result of −0.380, which is medium-low and negative. Regarding the correlation between NSC and CR, the values were −0.978 and −0.810 in Holstein and Jersey cattle, respectively, which are similar to the −0.990 reported by Kadarmideen et al. (2000).

Conclusions

This study confirms that reproductive traits present low heritability, <10% in most cases, in Holstein and Jersey cattle, suggesting that fertility is affected mostly by the environment. Therefore, good management of the fertility traits must be considered in order to improve reproductive efficiency. The high and negative genetic correlations for CI × CR, DO × CR, and NSC × CR, and high and positive genetic correlations for CI × NSC, CI × DO, and DO × NSC in this study suggest that these reproductive traits are genetically equivalent, i.e., they are influenced by the same genes, which reveals the presence of pleiotropic effects. This obviously favors the selection of these traits, as we can predict what will happen to several of the reproductive traits after performing selection on one of them. In this manner, we can integrate information on different traits to propose more efficient selection strategies.
References


