

Do subsidies drive technical efficiency? The case of portuguese firms in the agribusiness sector

Os subsídios geram eficiência técnica? O caso das empresas portuguesas no setor do agronegócio

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Abstract: Successive reforms in the Common Agricultural Policy (CAP), enlargements of the European Union (EU) and pressures in the public budget increase the need for empirical studies to assess the role of subsidies in shaping firms' performance and, hence, to support policy-makers decisions and decision-making in the productive sector. This paper assesses whether subsidies affect the technical efficiency of 1,943 Portuguese agribusiness firms over 2007-2015. By applying stochastic frontier models and fixed-effects models, to calculate the average efficiency of these firms, the results show different effects of subsidies among agribusiness sectors. The stochastic frontier model revealed that the subsidies have a positive and significant impact on the technical efficiency of three sectors, animal production, manufacture of food, and manufacture of paper. Further, except for firms in the forestry and logging sector, Portuguese agribusiness firms had higher average efficiency levels when subsidised, which provides evidence that subsidies could be important and contribute positively to the technical efficiency of Portuguese firms.

Keywords: agribusiness, stochastic frontier, efficiency, subsidies.

Resumo: As sucessivas reformas na Política Agrícola Comum (PAC), os alargamentos da União Europeia (UE) e as pressões no orçamento público aumentam a necessidade de estudos empíricos para avaliar o papel dos subsídios na definição do desempenho das empresas e, dessa forma, apoiar os tomadores de decisões, tanto para definições de políticas públicas quanto para o setor produtivo. Este artigo avalia se os subsídios afetam a eficiência técnica de 1.943 empresas portuguesas do agronegócio no período de 2007 a 2015. Utilizaram-se modelos de fronteira estocástica e modelo de efeitos fixos para calcular a eficiência média dessas empresas. Os resultados mostram diferentes efeitos dos subsídios entre os setores do agronegócio. O modelo de fronteira estocástica revelou que os subsídios têm um impacto positivo e significativo na eficiência técnica de três setores: produção animal, fabricação de alimentos e fabricação de papel. Além disso, com exceção das empresas dos setores florestal e madeireiro, as empresas do agronegócio português apresentaram níveis médios de eficiência mais altos quando subsidiadas, o que fornece evidências de que os subsídios podem ser importantes e contribuir positivamente para a eficiência técnica das empresas portuguesas.

Palavras-chave: agronegócio, fronteira estocástica, eficiência, subsídios.

1 Introduction

Apart from being a major food and fibre supplier, agriculture is the main contributor to the rural vitality, because it generates employment and it prevents rural depopulation whilst keeping traditions alive (Organisation for Economic Co-operation and Development, 2013; McManus et al., 2012). While the abandonment of agricultural land represents significant



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threats to the environment in terms of biodiversity and loss of natural capital (MacDonald et al., 2000; Cramer et al., 2008), it may also fuel negative socio-economic impacts in rural areas. Additionally, various activities and businesses are created and developed in sectors related to agriculture, resulting in the agribusiness sector.

To avoid the abandonment of agricultural land, rural depopulation, and to foster sustainable agricultural activities, Giannakis & Bruggeman (2015) highlight the importance of maintaining compensatory subsidies to farmers with permanent natural handicaps, to increase rural income and prevent the rural exodus. The Common Agricultural Policy (CAP) within the framework of the European Union, is the main political instrument designed to support and develop the Member States agriculture. Ever since its launch in 1962, CAP has gone through a series of reforms, which modify producers' decisions regarding the choices of inputs and production, affecting their economic performance (Zhu & Lansink, 2010; Giannakis & Bruggeman, 2015; Daugbjerg et al., 2005). Moreover, further enlargements of the European Union (EU) and the pressure on the public budget increase the need to investigate the impact of subsidies on the efficiency of production units. Hence, empirical studies, personal experiments, and reliable studies are essential for policymakers (Minviel & Latruffe, 2017) as well as for decision-making in the agricultural sector.

It is worth stressing that theoretical results on the link between subsidies and efficiency are rather ambiguous. If, on the one hand, subsidies can reduce the efforts of producers (Martin & Page Junior, 1983) or change their attitudes when facing risk (Serra et al., 2008) thereby resulting in a reduction of technical efficiency, on the other, one should expect that subsidies increase technical efficiency while offering producers the means to maintain cutting-edge technologies and to improve organisational efficiency (Roe et al., 2003; Zhu & Lansink, 2010; Zhu et al., 2012; Góral, 2015). In this light, empirical evidence unfolding the path (significantly negative, significantly positive, or not significant) of the observed effects is sensitive to how subsidies are specified in the studies (Minviel & Latruffe, 2017; Latruffe & Desjeux, 2016; McCloud & Kumbhakar, 2008). Due to controversial results found in the literature on subsidies and efficiency, several authors such as Serra et al. (2008), Kumbhakar & Lien (2010), Zhu & Lansink (2010), argue that investigating the effects of subsidies on efficiency is an essentially empirical matter.

Overall, available studies portray subsidies negatively associated with efficiency and productivity (Giannakas et al., 2001; Rezitis et al., 2003; Iraizoz et al., 2005; Karagiannis & Sarris, 2005; Zhengfei & Lansink, 2006). In contrast, there are also empirical studies whose results shed light on the positive impact of subsidies on technical efficiency and productivity (McCloud & Kumbhakar, 2008; Kumbhakar & Lien, 2010; Kumbhakar et al., 2014). Most of the studies are focused on the United States and Europe. However, efficiency analyses considering Portuguese firms operating in the agribusiness sector are quite rare (Giannakis & Bruggeman, 2015; Daugbjerg et al., 2005; Tranter et al., 2007), even though the recent dynamism of Portuguese agribusiness sector. In this context, to the best of our knowledge, this is the first study examining the effects of subsidies on technical efficiency of firms belonging to the Portuguese agribusiness sector, mainly firms in the sector of animal production, processing and preservation of meat, milk and its by-products, the animal feed production, logging, paper manufacture, and related products. To this end, we use the stochastic frontier model with panel data as well as the fixed effects model to calculate the average efficiency of the production units. The sample embodies data obtained in the Amadeus platform by selecting 1,943 firms whilst covering for nine years (2007-2015).

The remaining part of the paper is organised as follows. Section 2 discusses the relationship between subsidies and efficiency. Section 3 presents the data, empirical variables, and the two econometric approaches applied to estimate technical efficiency and its main explanatory factors. Section 4 provides the discussion on empirical results, while Section 5 presents the main conclusions and offers some research avenues for further work.

2 Subsidies and technical efficiency: Is there a link?

It is interesting to underline that research into the impacts of subsidies on the technical efficiency of productive units has stood out significantly in the last 10 years. Most of the studies have mainly addressed developed countries, especially Europe, Canada, and the United States, while there is somewhat restricted literature on the emerging countries about

this matter. The exception is only India, China, and Brazil to which there can be found some evaluating literature (Minviel & Latruffe, 2017). The empirical results differ not only in the context of the studies (e.g., country, period, level of education of producers and production types) but also on the data used (for example, number of production units, cross-section data or panel data) and also in what relates to the used methodology (for example, parametric or non-parametric approach). Moreover, the empirical results are not conclusive, as there can be found significant effects, both positive and negative, of subsidies on the firm's technical efficiency, as well as effects without any statistical significance. Hence, the orientation of the subsidy-efficiency relationship found in the empirical literature may be related to a variety of factors, empirical contexts, data, and methodology used, so it is recommended some precaution when interpreting these results (Zhu & Lansink, 2010; McCloud & Kumbhakar, 2008; Minviel & Latruffe, 2017).

Overall, subsidies do not explicitly envision to improve efficiency, but rather increase productivity, promote environmental sustainability, animal health and welfare, and food safety, whilst improving the income of the producers or encouraging the production of certain resources, including environmental resources (European Union, 2013). The concepts of productivity and efficiency are two different phenomena (Coelli et al., 2005). While the productivity of a farm or firm refers to a given level of output per unit of input, the productive efficiency is to better decision-making processes. It can be classified as technical efficiency (TE) geared for input or output, whereas the allocative efficiency (AE) represents the ability of a firm to use inputs in optimal proportions, given their respective prices and production technology. Technical efficiency when input-oriented shows how a firm can reduce its costs or inputs without changing its production level. Otherwise, the measurement of TE can also be output-oriented when it analyses the amount of revenue that could increase without adding any input. The methodology for measuring efficiency was firstly used by Farrell (1957) and a more detailed treatment has been offered by Färe et al. (1985), Lovell (1996), Cooper et al. (2011).

Theoretical studies on the effects of subsidies on efficiency support either the negative or the positive impacts. The first may result from income effects, and it means that a stable income can distort investment decisions (Zhu & Lansink, 2010; Young & Westcott, 2000; Hubbard, 1998) and reduce incentives to produce efficiently. Insurance effects can exacerbate that result when producers are risk-averse. Any measure that reduces risk or increases the expected income will affect production (Burfisher & Hopkins, 2003; Lopez, 2001). Hennessy (1998) highlighted that the agricultural income support policies affect directly risk-averse producers' decisions in contexts of uncertainty. On the other hand, subsidies may also help to overcome those limitations that prevent an efficient restructuring or modernisation of firms. Therefore, firms may increase technical efficiency, thereby improving their production capacity, if they have available the financial resources to realise the required investments to maintain or update firm's production technologies (Roe et al., 2003; Zhu & Lansink, 2010; Zhu et al., 2012; Góral, 2015).

Giannakis & Bruggeman (2015) investigated the factors that explain the different performance of agriculture among the 27 countries of the European Union (EU) for five years (2007-2011). It is one of the scarce studies that addressed the issue of efficiency in Portugal, which is why some results are used in this study. The investigation focused on the productivity indicators of labour and land, and also gross value added. The results pinpointed that the characteristics of the farming population, the location in more favourable environments, and financial investments do have significant effects on the economic performance of the agricultural sector. The research has shown that the odds of achieving high economic performance are almost 9 times higher for those countries with a highly trained young agricultural population, namely the Netherlands (72%) and Germany (69%) than countries where farmers hold a poor rural training and there is a high proportion of farmers with more than 55 years, for example in Portugal (72%), Bulgaria (66%) and in Italy (64%). These countries are 92% less likely to be outperforming. It is also of paramount importance to note that environmental conditions did also show significant impacts, such as the countries with a high share of agricultural land used in less-favoured areas are 94% less likely to achieve high performance. This is the case of countries where the participation of

agricultural area used in Less Favoured Areas (LFA) is greater than 50% as in Portugal (92%) in Spain (82%) in Greece (78%) and Malta (100%). Another important issue is the economic disparities among EU members. The results suggested that countries displaying high economic performance are those who receive greater financial support to stimulate rural exploitation, such as Denmark (23,665 euros/farm), Luxembourg (17,700 euros/farm), and France (17,185 euros/farm). These countries are 17 times more likely to achieve high performance as compared to countries displaying a low-performance level. The wide variation in government support to rural income deteriorates further the position of the agricultural regions with chronic structural constraints and with small economies of scale. As per Smit et al. (2001), the high economic resources increase adaptive capacity, while the lack of these financial resources limits the ability to adapt to adverse conditions.

3 Data, empirical variables, and econometric approach

The latest techniques used to measure the efficiency and its relationship with subsidies refer to the parametric frontier and non-parametric frontier. The most popular non-parametric approach is Data Envelopment Analysis (DEA) which does not impose restrictions while specifying functional forms of production and restrain from making assumptions about one side of the error term because the non-parametric frontiers use approaches of linear programming. This is consistent with the line of thought as offered by Charnes et al. (1978), Fried et al. (2008) and Cooper et al. (2011). Indeed, the parametric frontier approach imposes constraints on the specification of the functional form of production and materialises assumptions about the unilateral distribution of the error term. The most common functional forms are the translog production function and Cobb-Douglas. Another distinction between the two models is that the non-parametric frontier is deterministic, while the parametric frontier is stochastic. Deterministic frontiers assume that all deviations from the frontier are the result of the inefficiency of the decision-makers, while in the stochastic frontier, it is assumed that part of the bypass is due to random shocks (reflecting on measurement errors and statistical noise) and part is due to the firm-specific inefficiency (Førsund et al., 1980; Battese, 1992; Coelli et al., 1998, 2005; Ajibefun, 2008; Greene, 2008).

The parametric frontier uses an econometric model to estimate the parameters of the stochastic frontier production function and also the model that affects inefficiency. The major advantage of the stochastic frontier approach is the introduction of a stochastic variable representing exogenous shocks, such as statistical noise, measurement errors, *etcetera*, which is not part of the component inefficiency and they are not beyond the control of the production unit. The parametric approach proves to be more applicable to the productive sector than a non-parametric approach, because the production units are subject to uncertainties (Battese & Coelli, 1995).

3.1 Data

Data were collected from Amadeus and SABI (Analysis of Iberian Balance Sheets) databases and withdrawn from the balance sheets of 1,943 Portuguese agribusiness firms. These databases cover the population of Portuguese firms as they encompass all firms that are obligated to report balance sheets and financial results. Therefore, our sample comprises all firms in the database with available information on the selected empirical variables, which could be seen as representing adequately the population of Portuguese firms in the agribusiness sector. Specifically, the sample gathers 208 firms from the manufacture of the paper sector, 18% (37) of which have been subsidised, 563 firms from the forestry and logging sector, of which 25% (140) have been allocated subsidies, 442 firms from the manufacture of food sector (processing and preserving of meat, dairy products, and prepared animal foods), of which 30% (134) have received subsidies, and 730 firms from the livestock industry sector (raising of swine, poultry, dairy cattle, sheep and other animals) of which 56% (408) have been subsidised.

For each sampled firm, we obtained a series on nominal values of operating revenue, intermediate inputs, and factor inputs (labour and capital) over the period 2007-2015.

Additionally, data on public subsidies and exports were collected from SABI. As both databases use the same firm identifier, matching the data proved to be an easy and reliable task.

The monetary values are calculated in euros and they were properly deflated using the GDP deflators for deflating the operating income, Gross Fixed Capital Formation to deflate capital and materials, the wages were deflated using the Private Consumer index, whereas we used the Public Consumption index to deflate the subsidies and the export used its respective index. Table 1 summarises the definition of each variable used in the empirical analysis and the respective descriptive statistics.

Table 1 – Variables description and descriptive statistics.

| Variables | Description | Mean | Max. | Min. | Std. Dev. |
|-----------|--|-------|--------|--------|-----------|
| Output | Logarithm of annual operating revenue as a proxy for the firm's production | 5.977 | 13.809 | -3.853 | 1.964 |
| Capital | Logarithm of fixed assets stock. | 5.185 | 13.693 | -5.423 | 2.156 |
| Material | Logarithm of intermediate inputs (materials, energy and service inputs) | 5.436 | 12.974 | -8.186 | 2.417 |
| Labour | Logarithm of wages and other employees' earnings | 4.229 | 11.275 | -5.161 | 1.595 |
| Subsidies | Logarithm of monetary amount transfer to firms from public agencies | 0.955 | 8.051 | -6.968 | 1.671 |
| Exports | Logarithm of total sales in international markets | 0.915 | 13.565 | -5.326 | 2.137 |

Notes: Monetary variables are in the 2011 prices.

The descriptive statistics (Table 1) show the heterogeneity of firms under scrutiny, when observing the values of the standard deviation, mainly for exports and subsidies variables. This is an expected result since they involve different phases of the production chain. As released in Appendix 1, the manufacture of food and manufacture of paper sectors, besides their being far more labour intensive, capital, and material display an export average of (1.862) and (2.684) respectively. These values are far beyond the average as achieved by the other two sectors, animal production (0.302) and forestry and logging (0.321). Another relevant issue is the average of subsidies for firms in different sectors. At the beginning of the production chain, animal production (1.535) and forestry and logging (0.628) are far superior to the average of subsidies for firms operating the manufacturing sectors, namely manufacture of food (0.683) and manufacture of paper (0.377).

Table 2 – Correlation coefficients among selected variables

| Variable | Output | Capital | Material | Labour | Subsidy | Exports |
|----------|--------|---------|----------|--------|---------|---------|
| Output | 1 | | | | | |
| Capital | 0.687 | 1 | | | | |
| Material | 0.888 | 0.558 | 1 | | | |
| Labour | 0.848 | 0.659 | 0.690 | 1 | | |
| Subsidy | 0.088 | 0.289 | 0.036 | 0.059 | 1 | |
| Exports | 0.549 | 0.433 | 0.484 | 0.556 | 0.042 | 1 |

Looking at the correlation among empirical variables, Table 2 summarises some high correlations between the inputs. Nonetheless, it does not represent a problem because the main interest of this research is not to estimate the contribution of each input to the production. In this light, the correlations do not indicate the existence of any significant econometric problem.

3.2 Econometric approach

To analyse the efficiency of productive units of the Portuguese agribusiness sector, this study has resorted to the Stochastic Frontier method of evaluation of efficiency, which is based on regression analysis for Maximum Likelihood. This method defines a specific functional form for a firm's technology and a probability distribution for the coefficient of efficiency and regression error (Bauer et al., 1998). As for the former one, it used the translog functional form, and as for the probability distributions, we opted for the half-normal one.

It should be noted that the results in terms of inefficiency obtained under certain conditions may actually offer more consistent values. Bauer (1990) and Rossi & Ruzzier (2000) propose a set of criteria to evaluate whether the estimated inefficiency obtained according to different models are consistent with each other, that is, comparing levels of inefficiency among the surveyed units. Therefore, we used the stochastic frontier model and the fixed effects model for analysing the effects of subsidies on the efficiency of firms.

Among the frontier methods for the assessment of efficiency, the stochastic statistical method is based on an econometric estimation technique, in which the deviations observed in the existing frontier are the result of both trends of random factors, which the firm does not control for, and the statistical noise. In this analysis, we used the method of the stochastic frontier through the measurement of total production, given the firm's operating revenues.

Whilst appraising the technical efficiency of the Portuguese firms operating in the agribusiness sector, the study used a production function, which assumes a stochastic production frontier model embodying technical effects that affect the component inefficiency. Also, we used a logarithmic transformation for all variables of the equation.

The specification of the production function is taken from the Cobb-Douglas family and we used a normal model-half-normal for the error structure. The normal components of random errors show zero mean and variances depending on the effects of the firms. The error components of inefficiency are independent of the random error components, with variations depending upon a set of technical effects, according to Aigner et al. (1977), Kumbhakar & Lovell (2003), Coelli et al. (2005), Greene (2008) and Fried et al. (2008). The model materialises into the following equation:

$$\ln y_{it} = \vec{x}_{it}' \beta + v_{it} - u_{it} \quad i = I, \dots, I,$$

Where y_i represents the log of the outputs of the firms' sample, x_i is a vector containing $K \times 1$ logarithm of the inputs, β is a vector of unknown parameters (elasticities of inputs), the $v_i \sim N(0, \sigma_{v_i}^2)$ are the components of random errors which are independent of the inefficiency errors components $u_i \sim N^+(0, \sigma_{u_i}^2)$. The total component error is $\varepsilon_i = v_i - u_i$. Based on the stochastic frontier model, the technical efficiency of the firms' sample is given by $E_i = E\{\exp(-u_i) | \varepsilon_i\}$.

Alternatively, the efficiency analysis can be realised by using the regression models applied to Panel data. This analysis can substantially improve the quality of the estimates, in particular, because over time, it allows the analysis of the efficiency of firms; moreover, it permits more accurate estimates as it allows to incorporate fixed and individual random effects. Finally, the main gain in using panel data is to offer consistent estimators of the individual parameters of efficiency (Kumbhakar & Lovell, 2003).

The fixed-effects model, the calculation of average efficiency is computed according to Kumbhakar et al. (2015, p 244). This model assumes that the most efficient unit in the sample is 100% efficient and thus, it is considered as the reference for the calculation of the average efficiency of the other production units. In this study, subsidies are used as a contextual factor, that is, as an explanatory variable of production and technical efficiency as found in several publications such as empirical studies by Zhu & Lansink (2010), Bojneč & Latruffe (2013), Kumbhakar et al. (2014) and Sipiläinen et al. (2014).

4 The impact of subsidies on the firm's technical efficiency

We estimated alternative empirical models to assess whether public subsidies can boost agribusiness firms' technical efficiency. Table 3 shows the estimation results of firms' production function based on the fixed-effects model, while Table 4 displays the estimates based on a stochastic frontier model. In both cases, the empirical results were obtained using the entire sample (i.e. all firms) and sub-samples for each agribusiness sector under analysis. This allows us to do a robustness check on the main findings.

In Table 3, the estimates were based on Ordinary Least Squares applied to panel data with fixed effects. The average efficiency was computed according to Kumbhakar et al. (2015, p. 244-245). Where the average efficiency was calculated using $\hat{u}_i = \max_i \{\hat{x}_i\} - \hat{x}_i \geq 0, i = 1, \dots, N$ and $\widehat{TE}_i = \exp\{-\hat{u}_i\}, i = 1, \dots, N$. The values of \hat{x}_i were recovered from the mean of the residual for each cross-sectional unit.

Overall, the estimates indicate that production is strongly driven by the intensity of the use of inputs as in the theory. Furthermore, public subsidies seem to be an important driver of technical efficiency, even though there are significant differences among agribusiness sectors.

Looking at the fixed effects model, firms with subsidies displayed an average efficiency higher than those without subsidies. The results show that the largest effect was on the livestock industry (production cattle, pigs, poultry, goats, sheep) in which 56% of firms are funded, and those that received subsidies outperformed in terms of average efficiency: over three times higher than those without subsidies. The manufacture of paper and the manufacture of food sectors unveil the largest average efficiency levels among the sectors examined, but subsidised firms are 1.40 and 1.129 times higher in efficiency terms than firms without subsidies, respectively. This seems to indicate that subsidies could have a decreasing impact on technical efficiency. If so, the role of subsidies in shaping technical efficiency seems to lose power as sectors improve their technical efficiency levels.

Table 3 – OLS estimates of Cobb-Douglas production function with fixed effects

| Variables | All firms | Manufacture of paper | Forestry and logging | Manufacture of food | Animal production | | | | | |
|--------------------|---------------------|----------------------|----------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|
| Capital | 0.080*** (0.015) | 0.073*** (0.009) | -0.008 (0.038) | 0.029*** (0.009) | 0.062 (0.046) | 0.091*** (0.017) | -0.023 (0.037) | 0.034** (0.014) | 0.092*** (0.018) | 0.079*** (0.023) |
| Material | 0.298*** (0.022) | 0.313*** (0.020) | 0.754*** (0.082) | 0.565*** (0.060) | 0.163*** (0.037) | 0.235*** (0.021) | 0.531*** (0.065) | 0.447*** (0.078) | 0.354*** (0.029) | 0.333*** (0.037) |
| Labour | 0.230*** (0.034) | 0.364*** (0.029) | 0.232** (0.096) | 0.345*** (0.050) | 0.303*** (0.105) | 0.422*** (0.045) | 0.429*** (0.075) | 0.350*** (0.078) | 0.169*** (0.034) | 0.258*** (0.047) |
| Subsidy | 0.049*** (0.007) | | 0.010** (0.005) | | 0.073*** (0.025) | | 0.016** (0.007) | | 0.061*** (0.012) | |
| Constant | 3.370*** (0.176) | 2.781*** (0.138) | 1.223** (0.518) | 1.502*** (0.247) | 3.468*** (0.520) | 2.874*** (0.188) | 1.739*** (0.336) | 2.295*** (0.346) | 2.927*** (0.170) | 2.898*** (0.277) |
| Average Efficiency | 0.072 | 0.058 | 0.457 | 0.326 | 0.045 | 0.086 | 0.280 | 0.248 | 0.139 | 0.045 |
| Obs. | 5,528 | 8,696 | 340 | 1,467 | 805 | 2,651 | 1,194 | 2,564 | 3,189 | 2,014 |
| R ² | 0.453 | 0.575 | 0.873 | 0.843 | 0.216 | 0.529 | 0.662 | 0.683 | 0.574 | 0.559 |

Notes: All regressions include time effects. Standard errors in parentheses. *, **, and *** mean that coefficients are statistically significant at 10%, 5%, and 1% level, respectively

A different finding emerges from the forestry and logging sectors, where the subsidised firms seem to unfold a lower average efficiency to those without subsidies. In this sector, most firms were not allocated subsidies, only 25% received subsidies and they have

unwrapped the lowest average efficiencies among the sectors examined. Several factors, such as environmental context, climatic conditions, manager training, could explain such poor performance in terms of technical efficiency. Giannakis & Bruggeman (2015) and Tranter et al. (2007) have already discussed the performance specificities of that agribusiness sector in Portugal and other European countries.

In accordance to Roe et al. (2003) under the assumption of imperfect markets, in the short and medium run, subsidies are likely to have a positive effect but small on the production, but technical progress should be achieved in the long run, because subsidised producers have fewer restrictions to credit and they are much more prone to invest more in cutting-edge technologies and development. McCloud & Kumbhakar (2008) have also found a positive effect when they assess the impact of subsidies on agricultural productivity in Denmark, Sweden, and Finland. They have pinpointed the fact that the effect on technical efficiency sheds light upon the link between subsidies and production.

Evidence showed that the estimated average efficiency in rural activities such as forestry and logging and animal production is between 4% and 14%. Although these values are low, they may be considered reasonable when compared to the values estimated by Zhu & Lansink (2010) for agriculture efficiency in Germany (64%), Netherlands (76%) and Sweden (71%) which are considered to be countries that hold the highest levels of efficiency in Europe. It is also important to mention the insightful Giannakis & Bruggeman (2015) study, which unveils that countries such as Germany, Sweden, and the Netherlands are 9 times more likely to outperform in rural activities countries such as Portugal, Spain, Greece, and Malta. These differences may be explained if we consider the rural population profile, the plot of land used in areas considered less favourable and also due to unequal financial support.

For that, Table 4 reports estimates from a stochastic frontier model in which firm's technical efficiency is specified as a function of public financial support – *subsidy* –, the variable of interest, and a firm's specific characteristics – *export* – that can be seen as a synthesis of several firm's specific characteristics, apart from subsidies, that would affect its efficiency level. Various studies offer robust evidence on the distinctive characteristics of exporters relative to non-exporters (see, e.g., Alvarez & López, 2005; Garcia-Marin & Voigtlander, 2019).

Overall, the estimates from the stochastic frontier model (Table 4) suggest that subsidies can drive the technical efficiency of agribusiness firms, even though there is significant heterogeneity among sectors. In three sectors - manufacture of paper, manufacture of food and animal production – subsidies appear to be a significant factor in reducing technical inefficiency, indicating that such agribusiness firms seem to use public financial support to foster investments in innovation and/or new technologies as a way to gain efficiency and competitiveness. In turn, firms in the forestry and logging sector seem to fail to use subsidies to restructuring production and improving efficiency through investing in efficiency-enhancing technologies. In fact, for those firms, subsidies seem to distort the incentives to improve efficiency. Those firms may use subsidies to improve other performance measures, such as exporting intensity, without a significant effect on efficiency. Also, on average, the estimated efficiency of those firms (see Table 3) is smaller than the estimated efficiency for other firms in the Portuguese agribusiness sector. Jointly, these findings suggest that the role of subsidies to improve efficiency may be restricted to specific sectors.

Table 4 – Estimates of Cobb-Douglas production function with half-normal distribution for inefficiency term: stochastic frontier half-normal model

| Variables | All firms | Manufacture of paper | Forestry and logging | Manufacture of food | Animal production |
|--|----------------------|----------------------|----------------------|----------------------|----------------------|
| Capital | 0.105*** (0.003) | 0.068*** (0.004) | 0.126*** (0.007) | 0.043*** (0.005) | 0.097*** (0.005) |
| Material | 0.399*** (0.003) | 0.551*** (0.006) | 0.297*** (0.005) | 0.595*** (0.005) | 0.451*** (0.004) |
| Labour | 0.419*** (0.005) | 0.349*** (0.008) | 0.487*** (0.011) | 0.354*** (0.008) | 0.311*** (0.008) |
| Constant | 2.078*** (0.029) | 1.446*** (0.035) | 2.497*** (0.063) | 1.356*** (0.031) | 2.180*** (0.042) |
| Explanatory variables for technical inefficiency variance function | | | | | |
| Subsidy | -0.078** (0.030) | -0.449** (0.148) | 0.281*** (0.033) | -0.781*** (0.210) | -0.189*** (0.039) |
| Export | -0.260*** (0.031) | -0.164*** (0.056) | -0.226*** (0.050) | -0.338*** (0.081) | -0.416*** (0.114) |
| Constant | -2.696*** (0.163) | -4.212*** (0.324) | -1.645*** (0.148) | -3.756*** (0.171) | -2.619*** (0.227) |
| No. obs. | 14,138 | 1,762 | 3,451 | 3,740 | 5,185 |
| Wald Chi2 | 95217.62*** | | | | |

Notes: All regressions include time effects. Standard errors in parentheses. *, **, and *** mean that coefficients are statistically significant at 10%, 5%, and 1% level, respectively. The difference between the number of observations in Table 4 and Table 3 (about 86 observations and not firms) is due to the lack of information on exports for some (few) firms.

Looking at international market involvement, exports appear to have a positive effect on the efficiency of agribusiness firms. This finding corroborates the results found in previous studies, such as Granér & Isaksson (2009), for firms in Kenya, Alvarez, & López (2005) for the firms in Chile, and Baldwin & Gu (2003) for firms in Canada. Overall, these findings show that exporters are more efficient than other firms and that exporting activities could yield efficiency and productivity gains.

5 Main conclusions

The objective of this study was to analyse the effect of subsidies on the technical efficiency of firms operating in the Portuguese agribusiness sector, specifically, it focused upon four main activities: (1) manufacture of paper, (2) forestry and logging (3) manufacture of food and (4) livestock production. The models used, fixed effects regression and stochastic frontier, have unveiled significant results and suggest that subsidies are likely to enhance the average efficiency and to gear positive effect on technical efficiency of the Portuguese firms, as well as exports as found in several empirical studies which analyse other countries (Granér & Isaksson, 2009; Alvarez & López, 2005; Baldwin & Gu, 2003).

Evidence shows that subsidised firms in livestock production hold an average efficiency three times higher than firms without subsidies. In what concerns the firms in the manufacture of paper, and manufacture of food, these have materialised the highest levels of average efficiency among the sectors examined, although those that received subsidies display values between 1.40 and 1.129 times higher the efficiency level than firms without subsidies in the same sector. As for the forestry and logging sector, the results showed that

subsidies harm technical efficiency and the average value of the efficiency is between 4 and 8%, which can be explained by the less favourable natural conditions and the profile of the Portuguese rural producers, which is consistent with the study by Giannakis & Bruggeman (2015). These differences also show that subsidies alone are not enough to make productive units fully adjusted to their economic and political environments. The social complexity that is specific to each region and other local factors seems to be useful to explain the different effects of subsidies on productivity (McCloud & Kumbhakar, 2008). Hence, further research and discoveries involving different factors are needed to fully explain the ongoing agricultural problems.

The study contributes to the literature by featuring interesting results on the average efficiency of Portuguese agribusiness firms. It also contributes to political issues, providing necessary insights into the development of policies that can enhance and strengthen the efficiency and productivity of the sector. As Minviel & Latruffe (2017) argue, there is no consensus in the empirical literature on the effects of subsidies on firm's technical efficiency, therefore it is necessary to continue researching and use advanced techniques and case studies to help policymakers in their decision-making processes by providing them with insightful studies and results.

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Appendix 1 - Descriptive statistics by sector

| Variable | Manufacture of paper | | | |
|----------------------|----------------------|--------|--------|-----------|
| | Mean | Max | Min | Std. Dev. |
| Operating revenue | 6.928 | 13.809 | -0.535 | 1.908 |
| Capital | 5.412 | 13.693 | -2.920 | 2.653 |
| Material | 6.402 | 12.974 | 0.221 | 2.054 |
| Labour | 5.281 | 11.275 | 0.402 | 1.603 |
| Subsidies | 0.377 | 7.609 | -5.396 | 1.220 |
| Exports | 2.684 | 13.565 | -4.632 | 3.153 |
| Forestry and logging | | | | |
| | Mean | Max | Min | Std. Dev. |
| | 5.230 | 11.318 | -3.853 | 1.616 |
| Operating revenue | 4.597 | 12.030 | -4.304 | 2.027 |
| Capital | 4.017 | 10.811 | -5.529 | 2.266 |
| Material | 3.835 | 8.466 | -5.161 | 1.266 |
| Labour | 0.628 | 7.159 | -6.968 | 1.420 |
| Exports | 0.321 | 8.342 | -4.826 | 1.207 |
| Manufacture of food | | | | |
| | Mean | Max | Min | Std. Dev. |
| | 7.274 | 12.285 | -2.754 | 1.958 |
| Operating revenue | 6.104 | 11.632 | -5.423 | 2.039 |
| Capital | 6.918 | 12.300 | -3.799 | 2.155 |
| Material | 5.265 | 9.858 | -2.801 | 1.539 |
| Labour | 0.683 | 8.051 | -5.187 | 1.456 |
| Exports | 1.862 | 11.098 | -5.326 | 2.787 |
| Animal Production | | | | |
| | Mean | Max | Min | Std. Dev. |
| | 5.477 | 11.139 | -3.133 | 1.709 |
| Operating revenue | 5.019 | 11.061 | -3.755 | 1.967 |
| Capital | 5.042 | 11.074 | -8.186 | 2.091 |
| Material | 3.513 | 8.423 | -4.694 | 1.297 |
| Labour | 1.535 | 7.388 | -6.091 | 1.900 |
| Exports | 0.302 | 10.283 | -3.538 | 1.207 |