



Impacts of the Bolsa Família Program on the labor market of **Brazilian municipalities**

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The Bolsa Família Program (PBF) benefits 14 million Brazilian families, and is one of the most important conditional cash transfer programs in Brazil in recent years. This articles investigates the impacts of the PBF in the formal labor market of the 5,570 Brazilian municipalities, from 2004 to 2013. The research used data from all Brazilian municipalities retrieved from the Cadastro Único (database of low-income families) of the Ministry of Social Development (MDS) and the Cadastro Central de Empresas (database of corporations) of the Brazilian Institute of Geography and Statistics (IBGE). The study adopted methods to estimate the ordinary least squares (OLS), and used panel data models as well as the dynamic models of Arellano and Bond (1991) in order to test hypotheses related to the association among the benefits from PBF, formal employment, and the total of salaries and other remunerations payed. The results indicate a positive association for the two hypotheses tested, in all three models applied.

Keywords: Bolsa Família Program; job market; income distribution.

Impactos do Programa Bolsa Família no mercado de trabalho dos municípios brasileiros

Beneficiando cerca de 14 milhões de famílias brasileiras, o Programa Bolsa Família (PBF) é um dos mais importantes programas brasileiros de transferência condicionada de renda dos últimos anos. A investigação dos impactos do PBF, no mercado de trabalho formal dos 5.570 municípios brasileiros, no período de 2004 a 2013, constitui o principal objetivo deste artigo. Usando dados totais de cada um dos municípios da União, obtidos no Cadastro Único do Ministério do Desenvolvimento Social (MDS) e no Cadastro Central de Empresas do Instituto Brasileiro de Geografia e Estatística (IBGE), métodos de estimação dos mínimos quadrados ordinários (MQO), modelos de dados em painel e modelos dinâmicos de Arellano e Bond (1991), testam-se as hipóteses da associação do benefício do PBF com a quantidade de pessoal ocupado assalariado formal e com o total dos salários e de outras remunerações. Nos três modelos utilizados, os resultados indicam associação positiva para as duas hipóteses testadas. Palavras-chave: Programa Bolsa Família; mercado de trabalho; distribuição de renda.

Impactos del Programa Bolsa Familia en el mercado de trabajo de los municipios brasileños

Al beneficiar a cerca de 14 millones de familias brasileñas, el Programa Bolsa Familia (PBF) es uno de los más importantes programas brasileños de transferencia monetaria condicionada de los últimos años. La investigación de los impactos del PBF, en el mercado de trabajo formal de los 5.570 municipios brasileños, en el período de 2004 a 2013, constituye el principal objetivo de este artículo. Con datos totales de cada uno de los municipios de la República, obtenidos del Registro Único del Ministerio de Desarrollo Social (MDS) y del Registro Central de Empresas del Instituto Brasileño de Geografía y Estadística (IBGE), métodos de estimación de los mínimos cuadrados ordinarios (MCO), modelos de datos en panel y modelos dinámicos de Arellano y Bond (1991), se prueban las hipótesis de la asociación del beneficio del PBF con la cantidad de personal ocupado asalariado formal y con el total de los salarios y de otras remuneraciones. En los tres modelos utilizados, los resultados indican una asociación positiva para las dos hipótesis probadas.

Palabras clave: Programa Bolsa Familia; mercado de trabajo; distribución de ingresos.

DOI: http://dx.doi.org/10.1590/0034-761220180026x

Article received on January 30, 2018 and accepted on May 29, 2019.

[Translated version] Note: All quotes in English translated by this article's translator.

ISSN: 1982-3134 @ ①

1. INTRODUCTION

Among the questions researched in the Social Sciences, such as those aimed at studying the wealth growth of nations or those directed to the distribution of wealth and to the increase of the well-being of society, an essential third one is that which deals precisely with the interrelations between growth and income distribution. The connection between product expansion and social equity has challenged researchers for decades, without a consensus on their reciprocal causality.

Pontusson (2005) expresses the essence of the implicit debate in the mentioned interrelation. On the one hand, liberalists signal the problem of increasing public spending and its harmful effects on private investment, accusing such policies for producing market distortions, reducing productivity (Easterly & Rebelo, 1993; Okun, 1975). On the other hand, defenders of interventionist policies point to the beneficial effects of this way of economic growth, whether in the increase of the domestic market or in the multiplier effect of income in times of economic crises (Barro, 1990, 1991; Lucas, 1988; Romer, 1986).

In the Brazilian case, it is possible to verify both the significant economic growth, for example, the one from 70s, of 129% (according to the Brazilian Institute of Geography and Statistics - IBGE) and the significant increase in inequality over the period (according to IBGE). In the 80s, successive economic policies, alternating freezes, hyperinflation and low economic growth, concurring with the worsening of social inequality. In late 90s, "compensatory" public policies began to reverse the extreme inequality situation. Cash conditional transfers, such as *Bolsa Escola* (School Allowance), *Bolsa Alimentação* (Food Allowance), Ajuda Gás (Gas Allowance), *Cartão Alimentação* (Food Card), were intensified, and, in 2003, the *Bolsa Família* (Family Allowance) Program (PBF) comes, whose main objective is to unify these four programs (Soares & Sátyro, 2009).

Published studies related to the PBF highlight its beneficial effects in terms of reducing poverty, increasing school attendance and reducing inequality (Bussmann & Risco, 2012; Jacinto & Tejada, 2009). However, PBF impact on the labor market and economic efficiency remains an open question.

This article associates the total transferred by the PBF in each of the 5,570 municipalities to local labor variables, specifically the amount formalized and the amount of wage received. For this, three different estimation methods are used: ordinary least squares (OLS), panel data models and Arellano and Bond model (1991). Significant coefficients of program impact on interest variables are found, based on administrative records of the Ministry of Social Development and Fight against Hunger (MDS) between the years 2004 and 2013. The results indicate the positive association for the 2 hypotheses tested, reinforcing previous studies that suggest that income transfers arising from the PBF can cause wage increases in a new market balance of the municipality.

2. IMPACTS OF THE BOLSA FAMÍLIA (FAMILY ALLOWANCE) PROGRAM IN THE FORMAL LABOR MARKET

According to Arretche (1995, p. 3), "the provision of social services, covering the most varied forms of individual and collective life risk has become a right ensured by the State to very expressive layers of the population of developed capitalist countries". Fleury (1994, p. 11) argues that "the analysis of the State intervention in social issued, through the so-called social policies, has the existence of a positive law relationship that is established between the citizen and his State as a corollary". This

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suggests that "state intervention through social policies, regulating and/or providing conditions for the maintenance and reproduction of a portion of the population, is considered an intrinsic function of the modern State, configuring patterns of social rights proper to each nation" (Fleury, 1994, p. 11).

Kerstenetzky (2011) argues that national social policies have maintained, from the 30s until the 70s, a pattern linked to the economic development model in which the main target groups of social policies are urban workers and civil servants. With the Constitution of the Federative Republic of Brazil (CF, 1988), however, the Brazilian social assistance is now inserted in the social security scope, seeking to replace an inefficient national welfare by a public policy of social assistance and welfare promotion, integrating, alongside health and social security policies, the general system of social protection.

With the new CF (1988), the Brazilian State becomes the main responsible for improving the population quality of life: health care programs are being developed, such as the Unified Health System (SUS), and cash conditional transfers programs - *Bolsa Escola* (School Allowance), *Bolsa Alimentação* (Food Allowance), Ajuda Gás (Gas Allowance), *Cartão Alimentação* (Food Card) - in order to combat poverty and reduce social inequalities. In 2003, these cash conditional transfers programs were grouped into a single system, considered more efficient: PBF. According to Draibe (2009, p. 105), "among the key factors pointed out by analysts, design simplicity, implementation flexibility, operational agility, relatively low administrative costs are highlighted".

Studies that investigate the impacts of PBF on the labor market and aggregate product are consolidated (Barbosa & Corseuil, 2014; T. Cavalcanti & Corrêa, 2010; Foguel & Barros, 2010; Ribas & Soares, 2011; Tavares, 2010), with the proposition of new theories (D. M. Cavalcanti, Costa, Silva, & Sampaio, 2016; Denes, Komatsu, & Menezes, 2016; Nunes & Mariano, 2015; Rands Barros, & Athias, 2013). Likewise, studies on the relationship between the benefit of PBF and the labor market, which are generally based on the data analysis of the National Household Sample Survey (PNAD) or the Consumer Expenditure Survey (POF), are growing.

According to Foguel and Barros (2010), the cash conditional transfers programs in Brazil do not change the average number of hours worked by men, but increase their participation rate and lead to a drop in hours worked by women, without influencing their participation rate. The authors estimate the effects of cash conditional transfers programs on the labor supply of adult women and men using the municipalities panel continuously covered by the PNAD or the IBGE's Census in the period between 2001 and 2005. According to Oliveira and Soares (2012), except for specific demographic groups, such as women with children, the disincentive to work is very small or nonexistent. Brito and Kerstenetzky (2010) determine the effects of the PBF on the beneficiary's labor market using data from the Single Registry of 2008 and PNAD of 2006. Santos, Leichsenring, Menezes, and Mendesda-Silva (2017) analyzed the effect of the PBF on the formal employment of the poor portion of the population using data between 2008 and 2011 of the Single Resgitry combined with data from the Annual Social Information Report, concluding that the chances of PBF beneficiaries remaining in employment are higher than for non-beneficiaries.

Regarding informality, there are studies whose results suggest that the PBF increases the probability of substitution effect: changing formal to informal employment of its beneficiaries (Marinho & Mendes, 2013) and those that indicate that the program has no impact on the occupational choice of beneficiaries between formal and informal employment (Barbosa & Corseuil, 2014).

International research also investigates the relationship between compensatory policies and the labor market (Asfaw, Davis, Dewbre, Handa, & Winters, 2014; Kabber & Waddington, 2015; Skoufia & Di Maro, 2008). Some studies are looking for comparative processes in the South American reality, approaching the main characteristics of the cash conditional transfers programs and their impacts in the regional labor markets (Alzúa, Cruces, & Ripani, 2012; Economic Commission for Latin America and the Caribbean & International Labour Organization, 2015).

However, despite the content of the international surveys, as well as the national ones, the impacts of the PBF in relation to the formal labor market still remain a matter in discussion.

3. METHOD AND DATA

According to Jacinto and Tejada (2009), models regarding efficiency and equality encompass three characteristic lines of investigation: a) dualistic models and the inverted U hypothesis; b) Kuznets curve and growth models; and c) Kuznets curve and social choice. This study belongs to the third group, which incorporates theories of endogenous growth involving political elements of social choice. Thus, to achieve the main objective of this work, two hypotheses were tested:

- **Hypothesis H1:** The PBF benefit is positively associated with the amount of salaried employees employed in the municipality.
- **Hypothesis H2:** The BPF benefit is positively associated with the total wages and other remunerations in the municipality.

The research covered the econometric deductive model. The first part of the statistical analysis brings the formulation of the linear model of variables of first difference, year by year, with OLS estimators. This regression model is linear for the parameters and can be represented by the function (1)¹:

$$\Delta Y_i = \beta_1 + \beta_2 \, \Delta X_i + u_i \tag{1}$$

The second and third parts of the statistical analysis bring panel data models. The second model can be represented by the function $(2)^2$, where the dependent variable is explained by a set of explanatory variables and an error term.

$$Y_{i,t} = \alpha_0 + \beta_1 X_{1i,t} + \beta_n X_{ni,t} + u_{i,t}$$
 (2)

Finally, observing that, in the regression analysis involving time series, the models can also include out-of-phase values, the third model contemplates a dynamic econometric analysis with out-of-phase

¹ Where: ΔY_i represents the differences of the dependent variable, **i** represents the ith cross-sectional unit, β_1 and β_2 represent the parameters to be estimated, ΔX_i represents the differences of the independent variable, e **u**, represents the independent error term.

² Where: \mathbf{Y}_{it} represents the ith cross-sectional unit in the tth time period, $\mathbf{\alpha}_0$ and $\boldsymbol{\beta}_1$, ..., $\boldsymbol{\beta}_n$ represent the parameters to be estimated, $\mathbf{X}_{1,it}$ represents the first exogenous variable, referring to the ith cross-sectional unit in the tth time period, $\mathbf{X}_{n,it}$ represents the nth exogenous variable, referring to the ith cross-sectional unit in the tth time period, and \mathbf{u}_{it} represents the ith error term in the tth time period.

autoregressive data, with out-of-phase values of the dependent variable, according to function $(3)^3$, where the dependent variable is explained by a set of explanatory variables and an error term.

$$Y_{it} = \alpha_0 + \beta_1 X_{1,it} + \dots + \beta_n X_{n,it} + \phi_1 Y_{it-1} + u_{i,t}$$
(3)

This self-regressive model is based on the idea that the dependent variable is not adjusted immediately due to changes in the explanatory variables and that past values tend to influence their own future values. In this study, we use the generalized method of moments (GMM) estimators, developed by Arellano and Bond (1991), which allows obtaining reliable estimators in lagged panel data models.

As variables, the following are used for each of the 5,570 (N) municipalities: a) total values received from the PBF benefits by all families living in the municipality; b) total amount of families benefited by the PBF; c) total number of employed and salaried people and with a formal contract living in the municipality; d) total value of wages and other remuneration of the employed salaried people and with a formal contract living in the municipality; e) total value of the business units of employed and salaried people with a formal contract living in the municipality; and (f) the total annual gross domestic product (GDP) of the municipality, as shown in Table 1.

BOX 1 STUDY VARIABLE

Variables per municipality	Source
Total benefit of the PBF	Single Registry (MDS)
Amount of families benefited by PBF	Single Registry (MDS)
Annual municipal GDP	Central Register of Enterprises (IBGE)
Wages and other remuneration	Central Register of Enterprises (IBGE)
Employed and salaried people on the last day of the year	Central Register of Enterprises (IBGE)
Local units	Central Register of Enterprises (IBGE)

Source: Elaborated by the authors.

The data are collected through the official governmental websites, thus, are considered a primary source from: a) IBGE, Directorate of Research, Methodology Coordination of Business Statistics, Registry and Classifications, Central Registry of Enterprises; and b) Single Registry, Social Information Matrix of the Ministry of Social Development (MDS), both referring to the years 2004 to 2013, with annual or monthly values accumulated annually.

³ Where: Y_{it} ; $X_{1,it}$; ...; $X_{n,it}$ and u_{it} have the same interpretations of function (4). Besides that, α_0 ; β_1 ; ...; β_n still represent the parameters to be estimated. The difference in this specification is that a new parameter $\phi 1$ appears representing the autoregressive parameter that needs to be estimated and relates to the first lag of the variable Y_{it} , namely, Y_{it-1} .

The choice of empirical observation interval cut-off between the 2004 and 2013 is due to 3 reasons: a) in the 2002 and 2003, PBF was either in the implementation phase or in the initial phase of operation and it is understood that, at that stage, there could be behavioral family change reversed in the short term; b) starting in 2004, the perennial effects of the program are already reflected; and c) the completion of the collection in 2013 is due to the limited availability of data on formally employed people and wages at the date of the research

The choice of the number of people employed and the wage and the average remuneration are based only on the formal work, that is, people who work with a formal contract, has the objective of analyzing the evolution not only of employment and wages, but also of the formalization of the Brazilian economy.

The choice to use the total specific data of each of the 5,570 municipalities to the detriment of the data of the PNAD or IBGE Census, which have the individual or the families as unit, is justified by the suggestion of "searching for a new database to verify conclusions from previous studies" (King, Keohane, & Verba, 1996, p. 27). Also because the purpose of the research is not about the individual or family behavior arising from the benefit of the program, but alternatively it evaluates the aggregate impact of the program. By using the total specific data of each of the municipalities of the Union rather than based on data from the PNAD or the IBGE Census, i.e., because it is not about the individual or family behavior arising from the benefit of the program, but, alternatively, to evaluate the aggregate impact of policies in the municipalities, this research allows the confrontation with the conclusions of previous studies, developed through the use of alternative databases. The data are collected directly from the official government agencies. Table 1 summarizes the total current values of the variables used.

TABLE 1 TOTAL ANNUAL CURRENT VALUES OF THE VARIABLES SURVEYED

Year	PBF benefit, in R\$ millions	Amount of families benefited by BFP, on a thousand families	GDP, in R\$ billion	Wages and other remuneration, in R\$ billion	Salaried employed people, in thousand people	Local units, in thousands
2004	3,792	6,572	1,941	398	31,138	5,929
2005	5,692	8,700	2,147	453	33,092	6,248
2006	7,525	10,966	2,369	547	34,998	4,764
2007	8,965	11,043	2,691	616	37,641	4,904
2008	10,607	10,558	3,187	716	39,456	5,120
2009	12,455	12,371	3,239	800	41,348	5,393
2010	14,373	12,778	3,990	929	44,141	5,704
2011	17,358	12,851	4,146	1,065	46,399	5,723
2012	21,157	13,902	4,390	1,211	47,510	5,814
2013	24,890	14,086	5,316	1,320	47,665	5,846

Table 2 shows the descriptive statistics for the variable of total benefit of the PBF by municipality between the years of 2004 and 2013. It is observed that the average value of the benefit grew, on average, 23.2% per year (p.y.) in the period.

TABLE 2 DESCRIPTIVE STATISTICS FOR THE VARIABLE OF TOTAL BENEFIT OF THE PBF, CURRENT **VALUES, IN R\$**

Year	Average	Median	Standard deviation	Minimum	Maximum
2004	685,427.52	274,848.50	1,985,023.77	168.00	83,629,166.00
2005	1,023,681.12	443,798.50	2,784,547.98	325.00	112,404,678.00
2006	1,352,383.42	594,705.00	3,574,349.73	720.00	129,006,666.00
2007	1,611,340.69	706,125.00	4,392,482.73	1,280.00	151,772,106.00
2008	1,906,272.50	823,372.00	5,323,984.63	1,134.00	172,598,270.00
2009	2,238,444.02	949,381.50	6,453,458.76	752.00	178,942,695.00
2010	2,582,695.93	1,103,867.00	7,170,577.59	660.00	200,137,472.00
2011	3,119,206.63	1,324,975.00	9,038,200.23	708.00	243,203,214.00
2012	3,801,751.07	1,593,996.00	10,874,276.25	4,408.00	317,886,792.00
2013	4,469,403.32	1,878,116.00	12,995,376.33	620.00	430,243,182.00

Source: Elaborated by the authors.

Table 3 shows the descriptive statistics for the number of families benefited by the PBF variable by municipality between the years of 2004 and 2013. It is observed that the number of families benefited by the PBF grew, on average, 8.8% p.y. in the period and there was a small decrease in the number of families benefited in 2008.

TABLE 3 DESCRIPTIVE STATISTICS OF THE VARIABLE NUMBER OF FAMILIES BENEFITED BY THE PBF

Year	Average	Median	Standard deviation	Minimum	Maximum
2004	1,188	518	3,666	2	165,991
2005	1,565	692	4,253	1	167,038
2006	1,971	887	5,689	3	224,817
2007	1,985	879	5,742	2	204,127
2008	1,898	834	5,649	1	201,664
2009	2,223	966	6,607	2	177,882
2010	2,296	1,006	6,926	1	196,699
2011	2,309	1,007	7,009	1	195,532
2012	2,498	1,073	7,704	4	228,078
2013	2,529	1,045	8,584	3	347,191

Table 4 shows the descriptive statistics for the GDP variable by municipality between the years of 2004 and 2013. It is observed that the average value of the GDP variable grew, on average, 11.8% p.y. in the period.

TABLE 4 DESCRIPTIVE STATISTICS OF THE ANNUAL MUNICIPAL GDP, CURRENT VALUES, **IN R\$ THOUSAND**

Year	Average	Median	Standard deviation	Minimum	Maximum
2004	349,190.35	48,198.13	3,723,248.80	3,183.93	225,170,381.88
2005	385,916.48	50,904.53	4,237,693.56	3,503.53	261,455,917.21
2006	425,859.73	56,963.78	4,610,458.18	3,746.16	282,892,454.76
2007	483,619.24	64,913.05	5,240,545.80	4,512.55	323,154,665.88
2008	572,622.07	74,416.66	6,173,015.78	6,492.19	357,116,681.33
2009	582,103.15	80,588.13	6,346,376.79	5,770.28	389,317,166.69
2010	719,171.08	93,296.52	8,679,265.76	7,237.63	443,600,101.65
2011	747,384.80	103,767.21	7,806,938.17	9,741.52	477,005,597.21
2012	791,496.87	109,785.52	8,181,913.32	8,342.57	499,375,400.66
2013	954,810.11	130,191.00	9,504,351.70	9,678.00	570,706,192.00

Source: Elaborated by the authors.

Table 5 shows the descriptive statistics for the wages variable and other remunerations received by the total employed and salaried people per municipality between the years 2004 and 2013. It is observed that the average value of wages and other remunerations grew, on average, 14.3% p.y.

TABLE 5 DESCRIPTIVE STATISTICS ON WAGES AND OTHER REMUNERATION, CURRENT VALUES, IN R\$ THOUSANDS

Year	Average	Median	Standard deviation	Minimum	Maximum
2004	71,641.06	3,342.26	1,101,459.89	-	67,555,828.38
2005	81,569.35	4,110.83	1,232,143.08	-	75,267,796.59
2006	98,407.89	5,216.56	1,471,465.41	12.59	89,181,133.84
2007	110,887.61	6,228.25	1,636,708.62	2.28	99,171,050.06
2008	128,854.13	7,348.09	1,893,471.47	16.29	114,669,100.24
2009	143,927.64	8,783.80	2,077,020.31	-	124,377,680.04
2010	167,492.48	10,260.29	2,376,854.55	47.70	142,150,179.34
2011	191,989.01	12,049.54	2,702,458.72	96.66	161,769,713.10
2012	218,414.95	13,972.78	3,054,433.77	8.03	182,963,919.90
2013	238,091.26	16,038.26	3,350,990.32	29.50	201,301,606.08

Table 6 shows the descriptive statistics for the total employed and salaried people variable by municipality between the years of 2004 and 2013. It is observed that the average value of employed and salaried people grew, on average, 4.9% p.y.

TABLE 6 DESCRIPTIVE STATISTICS OF THE EMPLOYED AND SALARIED PEOPLE VARIABLE

Year	Average	Median	Standard deviation	Minimum	Maximum
2004	5,608	582	61,241	+	3,543,743
2005	5,956	630	64,552	+	3,731,795
2006	6,301	683	68,326	2	3,975,545
2007	6,776	732	73,570	4	4,277,939
2008	7,103	751	77,457	1	4,528,310
2009	7,442	830	79,913	-	4,654,681
2010	7,961	875	84,378	3	4,906,877
2011	8,368	921	88,157	6	5,098,791
2012	8,568	919	90,726	1	5,239,502
2013	8,598	970	89,971	3	5,281,006

Source: Elaborated by the authors.

Table 7 shows the descriptive statistics for local units of the municipality between the years 2004 and 2013. It is observed that the average value of the local units of the municipality decreased, on average, 0.1% p.y. in the period. There is also a decrease in local units in 2006, with recovery in the immediately following year.

TABLE 7 DESCRIPTIVE STATISTICS OF THE LOCAL UNITS VARIABLE

Year	Average	Median	Standard deviation	Minimum	Maximum
2004	1,068	191	9,375	5	571,309
2005	1,125	203	9,845	4	600,936
2006	857	154	7,689	2	472,143
2007	883	160	7,992	6	491,553
2008	922	166	8,416	4	520,533
2009	971	180	8,798	7	544,983
2010	1,029	192	9,320	7	577,190
2011	1,032	187	9,423	4	584,225
2012	1,049	187	9,679	6	599,434
2013	1,054	195	9,718	7	618,713

The obtained data may not present, when descriptively analyzed, normal distribution, exposing obliquity on one side when presenting the frequency distribution. This occurrence suggests some kind of transformation and the transformation used here is that of the Napierian logarithm.

Also, for a better quality of comparison of the current monetary values, their constant values are used. Thus, the implicit GDP deflator obtained from the IBGE, the Coordination of National Accounts Research Directorate and the National Consumer Price Index (INPC) is used for the other monetary amounts, also obtained in IBGE, according to Table 8.

TABLE 8 IMPLICIT DEFLATOR OF GDP AND INPC BETWEEN 2004 AND 2013

Vern	Implicit	GDP deflator	INPC		
Year	Percentage	Accumulated divider	Percentage	Accumulated divider	
2004		1.000		1.000	
2005	0.0749	1.075	0.051	1.051	
2006	0.0669	1.147	0.028	1.080	
2007	0.0640	1.220	0.052	1.136	
2008	0.0887	1.328	0.065	1.209	
2009	0.0735	1.426	0.041	1.259	
2010	0.0857	1.548	0.065	1.341	
2011	0.0831	1.677	0.061	1.422	
2012	0.0587	1.775	0.062	1.510	
2013	0.0651	1.891	0.056	1.594	

Source: Elaborated by the authors.

4. RESULTS

This section presents, through the three models, first differences, fixed effects and Arellano-Bond, the results in the 2 hypotheses tested.

4.1 Ordinary least squares

The use of the OLS estimation method of the variables in the first difference (year by year) is a simple tool to deal with the bias problem generated by the omission of variables, once, assuming that the heterogeneity is constant in time, it will be withdrawn in the subtraction. One advantage of this model is that it removes the heterogeneity, regardless of the model, from fixed or random effects, in order

to be the most appropriate one. Such a method causes the coefficients resulting from the OLS to be consistent and non-skewed. Likewise, $\Delta u_{i,t}^{\ 4}$ can be correlated in time, indicating that more advanced models must be used (Grene, 2008).

• **Hypothesis H1:** The PBF benefit is positively associated with the amount of salaried employees employed in the municipality.

Model 1:

Model 2:

$$\ln \Delta \text{ salaried pop} =$$

$$= \beta_o + \beta_1 \ln \Delta \text{ local units} + \beta_2 \ln \Delta \text{ GDP} + \beta_3 \ln \Delta \text{ PBF beneficiaries} + u_i$$
(5)

Table 9 presents the results of the PBF impact on salaried people from the municipality at constant values of 2004.

TABLE 9 IMPACT OF THE PBF ON SALARIED EMPLOYEES OF THE MUNICIPALITY, THE NAPIERIAN LOGARITHM OF THE VARIABLES IN THE FIRST DIFFERENCE, CONSTANT VALUES FOR 2004

	Model 1	Model 2
Ln Δ local units	0.4259***	0.3929***
Ln Δ GDP of the municipality	0.2568***	0.2947***
Ln Δ PBF distributed value	0.22149***	
Ln Δ PBF number of beneficiaries		0.2033***
Constant	-3.1462***	-1.9495***
R ²	0.47	0.48

Source: Elaborated by the authors.

Model 1 indicates that, considering controls as number of local municipal units and the GDP of the municipality, the total distributed PBF value is positively associated to the number of employed and salaried people in the municipality. The 1% increase in the difference of the value distributed between one year and another leads to an increase of 0.22% in the difference of salaried people in the municipality between 2 years.

 $^{^4\}Delta u_{i,t}$ means the first difference also applied to the random error.

Model 2 shows that an 1% increase in the difference of the beneficiaries between 2 years leads to an increase of 0.20% in the difference of the employed people between 2 years. The 2 models confirm the hypothesis.

• **Hypothesis H2:** The BPF benefit is positively associated with the total wages and other remunerations in the municipality.

Model 1:

$$ln Δ total wages = = β0 + β1 ln Δ salaried pop + β2 ln Δ GDP + β3 ln Δ PBF distributed value + u$$
 (6)

Model 2:

$$\ln \Delta \text{ total wages} =$$

$$= \beta_0 + \beta_1 \ln \Delta \text{ salaried pop} + \beta_2 \ln \Delta \text{ GDP} + \beta_3 \ln \Delta \text{ PBF beneficiaries} + u_i$$
(7)

Table 10 presents the results of the PBF impact on the municipal total salaries at constant values for 2004.

TABLE 10 PBF IMPACT ON MUNICIPAL TOTAL SALARIES, NAPIERIAN LOGARITHM OF THE VARIABLES IN FIRST THE DIFFERENCE, CONSTANT VALUES OF 2004

	Model 1	Model 2
Ln Δ salaried population	0.6614***	0.6332***
Ln Δ GDP of the municipality	0.2009***	0.2279***
Ln Δ PBF distributed value	0.1098***	
Ln Δ PBF number of beneficiaries		0.1201***
Constant	6.6526***	7.1032***
R2	0.69	0.70

Source: Elaborated by the authors.

Model 1 indicates that the 1% increase in the difference between the 2 years in the PBF distributed value leads to an average increase of 0.11% in the total wage difference of the municipality. Model 2 indicates that the 1% increase in the difference between the 2 years in the number of PBF beneficiaries leads to an average increase of 0.12% in the total wage difference of the municipality. The 2 models confirm the hypothesis.

4.2 Panel data models

The second part of the statistical analysis presents the results of the panel data model, sets of data in which the same cross-sectional units are monitored over time. The Hausman (1978) test, for the 2 hypotheses, indicates the use of fixed effects model in detriment of the random effects model.

• Hypothesis H1: The PBF benefit is positively associated with the amount of salaried employees employed in the municipality.

Model 1:

salaried pop =
$$\beta_0 + \beta_1$$
 local units + β_2 GDP + β_3 PBF beneficiaries + u_i (8)

Model 2:

salaried pop =
$$\beta_0 + \beta_1$$
 local units + β_2 GDP + β_3 PBF distributed value + α_1 (9)

Table 11 presents the results of the PBF impact on the employed and salaried people from the municipality, at constant values of 2004.

TABLE 11 PBF IMPACT ON THE EMPLOYED AND SALARIED PEOPLE FROM THE MUNICIPALITY, **CONSTANT VALUES OF 2004**

Independent Variables	Model 1	Model 2
Local units	4.91***	4.49***
GDP of the municipality	0.000000423	0.000000371
PBF beneficiaries	2.97***	
PBF distributed value		0.00272***
Constant	-3,906.25	-2,027.73**

Source: Elaborated by the authors.

Model 1 indicates that, when considering controls such as the municipal GDP and the number of local units in the municipality, an additional beneficiary of the PBF generates an increase of 2.97 salaried employed people in the municipality, with a significance of 99%. Model 2 indicates that an increase of one monetary unit granted by the program entails an increase of 0.003 salaried person in the municipality, with a significance of 99%. Therefore, both models confirm the hypothesis that PBF is positively associated with the number of salaried employees in the formal sector.

Hausman test. H0: no correlation between the regressors and the error term.

TABLE 12 HAUSMAN TEST. HYPOTHESIS 1, MODEL 1

	EF	ER	EF-ER	SE
Local units	4.917	6.715	-1.798	0.0319
PBF beneficiaries	2.972	2.663	0.309	0.0113
GDP of the municipality	0.000000423	0.000000566	-0.000000143	
Chi ²	4548.06			
Prob>Chi ²	0			

TABLE 13 HAUSMAN TEST. HYPOTHESIS 1, MODEL 2

	EF	ER	EF-ER	SE
Local units	4.494	6.977	-2.483	0.0331
PBF distributed value	0.0027	0.0026	0.0001	0.00000464
GDP of the municipality	0.000000371	0.000000501	-0.00000013	
Chi ²	6629.41			
Prob>Chi ²	0			

Source: Elaborated by the authors.

Table 12 presents the Hausman test of Hypothesis 1 for model 1, where EF stands for fixed effects, ER stands for random effects, EF-ER is the difference between both of them and SE is the standard error, and Table 13 presents the Hausman test of Hypothesis 1 for model 2. In both models, the Hausman test indicates that the null hypothesis must be rejected ($Prob > chi^2 < 0.05$), because there is a correlation between the regressors and the error. Therefore, the fixed effects model must be used.

• **Hypothesis H2:** The BPF benefit is positively associated with the total wages and other remunerations in the municipality.

Model 1:

total wages =
$$\beta_0 + \beta_1$$
 salaried pop + β_2 GDP + β_3 PBF beneficiaries + u_i (10)

Model 2:

total wages =
$$\beta_0 + \beta_1$$
 salaried pop + β_2 GDP + β_3 PBF distributed value + u_i (11)

Table 14 presents the results of the PBF impact on the municipal total salaries at constant values for 2004.

TABLE 14 PBF IMPACT ON THE TOTAL MUNICIPAL WAGES, CONSTANT VALUES FOR 2004

Independent variables	Model 1	Model 2
Salaried population	26,076.13***	26,028.78***
GDP of the municipality	0.0115**	0.0114***
PBF beneficiaries	16,448.24^	
PBF distributed value		12.5
Constant	-1,040,000,000***	-1,046,000,000***

Model 1 indicates that, controlling by the municipal GDP and the number of salaried people in the formal sector, the increase of a beneficiary of the PBF leads to an increase of approximately R\$ 16,500.00 in the total wages of the municipalities. However, the equation coefficient has lower security than generally accepted in Social Sciences works. Model 2 indicates a positive relation between the value distributed by the PBF and the total wages of the municipality. However, the coefficient is not statistically different from zero. Consequently, the above models demonstrate a positive relationship between the PBF and total wages, but it is not possible to fully accept Hypothesis 2 based on the above estimates.

Hausman test. H0: no correlation between the regressors and the error term.

TABLE 15 HAUSMAN TEST. HYPOTHESIS 2, MODEL 1

	EF	ER	EF-ER	SE
Salaried population	26028.78	20884.44	-1.798	0.0319
PBF distributed value	0.0114	0.01274	0.309	0.0113
GDP of the municipality	12.5502	16.2723	-0.000000143	
Chi ²	17326.12			
Prob>Chi ²	0			

Source: Elaborated by the authors.

TABLE 16 HAUSMAN TEST. HYPOTHESIS 2, MODEL 2

	EF	ER	EF-ER	SE
Salaried population	26028.78	20884.44	-1.798	0.0319
PBF distributed value	0.0114	0.01274	0.309	0.0113
GDP of the municipality	12.5502	16.2723	-0.000000143	
Chi ²	17326.12			
Prob>Chi ²	0			

Source: Elaborated by the authors.

Table 15 presents the Hausman test of Hypothesis 2 for model 1 and Table 16 presents the Hausman test of Hypothesis 2 for model 2. In both models, the Hausman test indicates that the null hypothesis must be rejected (Prob > chi² < 0,05), because there is a correlation between the regressors and the error. Therefore, the fixed effects model must be used.

4.3 Arellano-Bond dynamic models

The same two hypotheses of the study are tested with dynamic models, which include the dependent variable lagged as regressor and the dependent variables lagged in some periods. The decision to include the lagged dependent variable can be based on theoretical arguments since, if there is reason to assume that the current values of the dependent variables are highly dependent on the past values, the inclusion of this lag reduces the bias problem due to the omission of variables. In dynamic models of this kind, fixed or random effect models must not be used, once, by construction, the lagged dependent variable is correlated with the error. This endogeneity presence makes the estimators inconsistent. In this case, the Arellano-Bond estimation method is the most appropriate one. This estimator uses the generalized method of moments (GMM) and provides efficient estimators in a dynamic model.

• Hypothesis H1: The PBF benefit is positively associated with the amount of salaried employees employed in the municipality.

Model 1:

salaried pop =
$$\beta_0$$
 + β_1 salaried pop $_{t-1}$ + β_2 local units $_t$ + β_3 local units $_{t-1}$ + β_4 local units $_{t-2}$ + β_5 GDP $_t$ + β_6 GDP $_{t-1}$ + β_7 GDP $_{t-2}$ + β_8 beneficiaries $_t$ + β_9 beneficiaries $_{t-1}$ (12) + β_{10} beneficiaries $_{t-2}$ + u_i

Model 2:

salaried pop =
$$\beta_{o}$$
 + β_{1} salaried pop $_{t-1}$ + β_{2} local units $_{t}$ + β_{3} local units $_{t-1}$ + β_{4} local units $_{t-2}$ + β_{5} GDP $_{t}$ + β_{6} GDP $_{t-1}$ + β_{7} GDP $_{t-2}$ + β_{8} distributed value $_{t}$ + β_{9} distributed value $_{t-1}$ + β_{10} distributed value $_{t-2}$ + u_{i} (13)

Table 17 presents 2 dynamic models for the Hypothesis 1 test of this study.

TABLE 17 PBF IMPACT ON THE EMPLOYED AND SALARIED PEOPLE FROM THE MUNICIPALITY, CONSTANT VALUES OF 2004, BY ARELLANO-BOND DYNAMIC MODEL

	Model 1	Model 2
Salaried population		
L1	0.21***	0.17***
Local units	7.85***	7.79***
L1	0.27	-0.49
L2	0.24	-0.049
GDP of the municipality	-0.000000141	-0.00000119
L1	-0.0000000959	0.000000217
L2	0.00000145*	0.00000166*
PBF beneficiaries	0.29	
L1	0.19	
L2	0.56***	
PBF distributed value		0.0003***
L1		0.0002
L2		0.0006*

The abbreviation L1 indicates that the variable was lagged in one period, while L2 indicates that the variable was lagged in 2 periods. The models show that there is a lagged effect on the response: in both models, the coefficients of the PBF beneficiaries and PBF total distributed values variables are significant and have a greater magnitude when they are out of date in 2 years.

Robustness test. H0: There is no errors autocorrelation in the first difference.

TABLE 18 ERRORS AUTOCORRELATION TEST OF THE FIRST DIFFERENCE IN HYPOTHESIS 1

	Model 1	Model 2
Prob > z	0.024	0.0203

Source: Elaborated by the authors.

Table 18 shows the errors autocorrelation test in the first difference. H0 must be rejected. Thus, there is evidence that the Arellano-Bond model assumptions are respected.

• **Hypothesis H2:** The BPF benefit is positively associated with the total wages and other remunerations in the municipality.

Model 1:

total salaried =
$$\beta_0 + \beta_1$$
 total salaried $_{t-1} + \beta_2$ local units $_t + \beta_3$ local units $_{t-1} + \beta_4$ local units $_{t-2} + \beta_5$ GDP $_t + \beta_6$ GDP $_{t-1} + \beta_7$ GDP $_{t-2} + \beta_8$ beneficiaries $_t + \beta_9$ beneficiaries $_{t-1} + \beta_{10}$ beneficiaries $_{t-2} + u_i$ (14)

Model 2:

Table 19 presents the results of the PBF impact on the municipal total salaries at constant values for 2004.

TABLE 19 PBF IMPACT ON THE TOTAL MUNICIPAL WAGES, CONSTANT VALUES FOR 2004, BY ARELLANO-BOND DYNAMIC MODEL

	Model 1	Model 2	
Total wages	-	-	
L1	0.57***	0.55***	
Local units	116,956.4***	116,757.3***	

Continue

	Model 1	Model 2
L1	29,356.79***	29,450.9***
L2	2,807.45***	-4,693.75
GDP of the municipality		0.0012
L1	0.0016	0.0008
L2	0.004**	0.0048**
PBF beneficiaries	7,342.5***	
L1	7,137.21**	
L2	11,790.01***	
PBF distributed value		14.43***
L1		-1.19
L2		-4.78

Source: Elaborated by the authors.

Model 1 also indicates the dynamic character of the impact of the PBF beneficiaries number and the total formal wages of the municipality. The coefficient of the explanatory lagged variable in two years indicates that the increase of a PBF beneficiary leads, on average, to an increase of R\$ 11,790.01 in the total wage of the municipality. The coefficients remain positive and significant for the level variable. For the variable lagged in one year, however, the impact is reduced to R\$ 7,137.21. In model 2, the dynamic effect does not appear. Only the coefficient of the level variable is significant and indicates that an increase of R\$ 1.00 in the PBF total value distributed leads to an increase of R\$ 14.00 in the total formal wages of the municipalities. In conclusion, the robustness test also indicates that there is an evidence that the Arellano-Bond model assumptions are respected.

Robustness test. H0: There is no errors autocorrelation in the first difference.

TABLE 20 ERRORS AUTOCORRELATION TEST OF THE FIRST DIFFERENCE IN HYPOTHESIS 2

	Model 1	
Prob > z	0.0084	0.0053

Source: Elaborated by the authors.

Table 20 presents the results of the autocorrelation test of the first difference errors. H0 must be rejected. Thus, there is evidence that the Arellano-Bond model assumptions are respected.

5. FINAL CONSIDERATIONS

According to the results obtained, the three models used (first differences, fixed effects, and Arellano-Bond) indicate the positive association of the municipality PBF benefits with salaried people and with wages and other income in the municipality.

From the academic point of view, it is understood as contribution that, by using the total specific data of each one of the Union municipalities instead of basing itself on the data from PNAD or IBGE Census, this research allows the confrontation with the conclusions of previous studies, developed through the use of alternative data bases. For that matter, the result presented herein corroborates previous results, obtained with alternative bases, which indicate the positive association of PBF with the labor market.

A second academic contribution is the application of the three econometric models, Napierian logarithms of the first differences, fixed-effects panel data model and Arellano-Bond dynamic panel data model, whose convergence of results solidifies the research.

From the public policy management point of view, it is understood as real and effective contribution the finding that the PBF is positively associated to the formal labor market of the Brazilian municipalities, both regarding the increase in salaried employed people and the increase in wages and other incomes. Such results reinforce previous studies that suggested that the transfers of income from the PBF can cause salary increases in a new market balance of the municipality. The finding that the results indicate that the PBF benefit does not cause adverse effects regarding the decisions related to the labor supply is added to the previous contribution.

Thus, the expectation of maintaining this cash conditional transfer program is preserved, both as a generator of social benefits, characterized by the fight against absolute poverty, improvement in education, health and fight against family hunger, and as an endogenous factor contributing to the growth of salaried employees and to wages and other income in the municipality.

Some limitations of this research must be mentioned. A first limitation refers to the nonconsideration of the municipality worker association with his/her address, since the data related to the individual may be allocated, regarding the PBF benefit, in a municipality and, as a person formally employed, in a different municipality. A second limitation is that the research only deals with a formal work, since the source used only offers this data. Thus, the substitution effect, derived from the informal work, is not considered.

Finally, some suggestions may be offered for the continuity of this study: a) research using data from 2014 and 2016, which incorporate a new trend of the unemployment rate; b) research encompassing the individual analysis of the municipalities, in order to observe in which municipal environments there is a greater adherence to the presented results; and c) research of possible characteristics in the set of municipalities where the positive association is more latent.

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