## BASIC AREA - Article

# Plot size and number of repetitions in vetch 

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#### Abstract

The objectives of this work were to determine the optimum plot size and number of repetitions to evaluate the fresh matter weight of vetch (Vicia sativa L.) in sowing densities. Fortyeight uniformity trials of $6 \mathrm{~m} \times 6 \mathrm{~m}$ were conducted. Sixteen trials were evaluated in each sowing density $\left(40,60\right.$, and $\left.80 \mathrm{~kg} \cdot \mathrm{ha}^{-1}\right)$. Each trial was divided into 36 basic experimental units (BEU) of $1 \mathrm{~m} \times 1 \mathrm{~m}$, totaling $1,728 \mathrm{BEU}$. In each BEU, the fresh matter was weighed. The optimum plot size was determined by the method of maximum curvature of the coefficient of variation model. The means were compared among sowing densities by the Scott-Knott's test. The number of repetitions - for experiments with completely randomized and randomized block designs, in scenarios of i treatments


( $\mathrm{i}=3,4, \ldots, 50$ ) and d minimal differences between treatments means being detected as significant with 5\% probability by Tukey's test, expressed in percentage of the experiment average ( $d=10 \%$, $11 \%, \ldots, 20 \%)$, was determined by iterative process until convergence. The optimum plot size to evaluate the fresh matter weight of vetch is $4.52 \mathrm{~m}^{2}$ at the 3 sowing densities. Four repetitions, to evaluate up to 50 treatments, in completely randomized and randomized block designs, are enough to identify as significant, at 5\% probability by Tukey's test, differences between treatment means of $29.15 \%$ of the experiment average.
Key words: Vicia sativa L., uniformity trials, experimental design, cover crops, experimental precision.

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## INTRODUCTION

The vetch (Vicia sativa L.) is an annual winter legume used as forage for animal feed and as cover crop (Santos et al. 2012). Researches with vetch, grown isolated or in mixture with other cover crops (Lithourgidis et al. 2006; Tuna and Orak 2007; Seidel et al. 2011; Ansar et al. 2013; Cherubin et al. 2014; Basbag et al. 2015; Desalegn and Hassen 2015), have demonstrated promising aspects of this crop. In isolated cultivation, in different environmental conditions, the fresh matter weights were $20.49 \mathrm{t} \cdot \mathrm{ha}^{-1}$ (Lithourgidis et al. 2006); $19.60 \mathrm{t} \cdot \mathrm{ha}^{-1}$ (Tuna and Orak 2007); $7.40 \mathrm{t} \cdot \mathrm{ha}^{-1}$ (Seidel et al. 2011); $11.13 \mathrm{t} \cdot \mathrm{ha}^{-1}$ (Ansar et al. 2013); $25.94 \mathrm{t} \cdot \mathrm{ha}^{-1}$ (Cherubin et al. 2014); $18.32 \mathrm{t} \cdot \mathrm{ha}^{-1}$ (Basbag et al. 2015) and $14.833 \mathrm{t} \cdot \mathrm{ha}^{-1}$ (Desalegn and Hassen 2015).

In studies with vetch, the experiments have presented variability in plot size, number of repetitions and sowing density. In vetch, oat and triticale cultivation in monoculture or in mixture, Lithourgidis et al. (2006) utilized plots of $100 \mathrm{~m}^{2}(20 \times 5 \mathrm{~m})$ and evaluated the fresh matter weight in $30 \mathrm{~m}^{2}(10 \times 3 \mathrm{~m})$ of each plot. Plots with $9 \mathrm{~m}^{2}(5 \times 1.8 \mathrm{~m})$ were used by Tuna and Orak (2007) to evaluate the fresh matter weight of vetch and oat, grown isolated and in intercropping. In research of 3 cover crop species: black oat, turnip and vetch, besides a control treatment, Seidel et al. (2011) used $0.50 \mathrm{~m}^{2}$ per plot ( 2 samples of $0.25 \mathrm{~m}^{2}$ ) in the determination of the fresh matter weight. In pure and mixed cultivation with oat, wheat, vetch and barley, Ansar et al. (2013) used plots of $15 \mathrm{~m}^{2}$ and evaluated the fresh matter weight in $1 \mathrm{~m}^{2}$ of each plot.

Three samples of $0.25 \mathrm{~m}^{2}$ within each plot $\left(0.75 \mathrm{~m}^{2}\right)$ were utilized to evaluate the fresh matter weight of vetch and other 6 cover crops (triticale, white oat, black oat, forage turnip, ryegrass and linseed), grown alone or in consortium (Cherubin et al. 2014). Nine species of vetch were studied in $7.2 \mathrm{~m}^{2}$ plots $(6 \times 1.2 \mathrm{~m})$, being the fresh matter weight evaluated in useful area of $3.0 \mathrm{~m}^{2}$ $(5 \times 0.6 \mathrm{~m})$ (Basbag et al. 2015). In another experiment, the fresh matter weight of 4 vetch species was measured in plots of $6.0 \mathrm{~m}^{2}(6 \times 1 \mathrm{~m})$ (Desalegn and Hassen 2015).

These researches with vetch crop were performed in a randomized block design (RBD) with 3 (Tuna and Orak 2007; Ansar et al. 2013; Cherubin et al. 2014; Desalegn and Hassen 2015) and 4 repetitions (Lithourgidis
et al. 2006; Seidel et al. 2011; Basbag et al. 2015). Sowing densities of $170 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ (Lithourgidis et al. 2006), $120 \mathrm{~kg} \cdot \mathrm{ha}{ }^{-1}$ (Tuna and Orak 2007), $45 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ (Seidel et al. 2011), $50 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ (Ansar et al. 2013) and $30 \mathrm{~kg} \cdot \mathrm{ha}{ }^{-1}$ (Desalegn and Hassen 2015) were used. According to Santos et al. (2012), the vetch sowing density to be used varies from 40 to $60 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$.

In experiments with cover crops, such as vetch, it is important to quantify with precision the fresh matter weight in sowing densities. For this, sizing properly the plot size and the number of repetitions is essential to obtain reliable results of the treatments under evaluation. Based on data obtained in uniformity trials, it is possible to determine the optimum plot size by the method of maximum curvature of the coefficient of variation model (Paranaíba et al. 2009). In this method, it is necessary to estimate the first-order spatial autocorrelation coefficient, variance and mean as well as to calculate the optimum plot size based on these estimates. Then, based on the optimum plot size, it is possible to determine the number of repetitions in scenarios formed of combinations of treatments numbers, experimental precision and experimental designs. In this way, it is possible to establish the appropriate experimental design for the required precision, according to the experimental area, time, financial resources and labor availability.

Dimensioning the optimum plot size based on the method of maximum curvature of the coefficient of variation model (Paranaíba et al. 2009) and number of repetitions in combinations of treatments and precision levels have been carried out to measure fresh matter weight of cover crops, such as: black oat (Cargnelutti Filho et al. 2014a), jack bean (Cargnelutti Filho et al. 2014b), forage pea (Cargnelutti Filho et al. 2015a), canola (Cargnelutti Filho et al. 2015b), millet in evaluation times (Burin et al. 2015), millet in sowing and cut times (Burin et al. 2016) and pigeonpea (Santos et al. 2016). For these crops, the optimal plot size oscillated between $4.14 \mathrm{~m}^{2}$ for black oat (Cargnelutti Filho et al. 2014a) and $8.39 \mathrm{~m}^{2}$ for pigeonpea (Santos et al. 2016). In these studies, in addition to the plot size, it was dimensioned the number of repetitions in combinations of treatments and precision levels in completely randomized and randomized block designs, which are references for researches with these crops.

There were not found research in the literature on the optimal plot size and number of repetitions to
assess the fresh matter weight of vetch (Vicia sativa L.). Moreover, possible difference of plot size among sowing densities is unknown. These investigations can be carried out from uniformity trials data and can provide useful information to the appropriate experimental design. Thus, the objectives of this work were to determine the optimum plot size and number of repetitions and to evaluate the fresh matter weight of vetch (Vicia sativa L.) in sowing densities.

## MATERIAL AND METHODS

The uniformity trials (experiment without treatment, in which the crop and all procedures performed during the experiment are homogeneous throughout the experimental area) were carried out with vetch (Vicia sativa L.), common cultivar, in the experimental area located at lat $29^{\circ} 42^{\prime} \mathrm{S}$, long $53^{\circ} 49^{\prime} \mathrm{W}$ and 95 m of altitude. According to Köppen, the climate is Cfa, humid subtropical, with hot summers and no dry season (Heldwein et al. 2009). The soil is classified as "Argissolo Vermelho distrófico arênico" Paleudalf (Santos et al. 2013). The sowing was broadcasted on May 13, 2015, within the indicated season to sow vetch, which extends from April to May (Santos et al. 2012), having been used the densities of 40,60 and $80 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ of seed. The following basic fertilization was utilized: $20 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ of $\mathrm{N}, 80 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ of $\mathrm{P}_{2} \mathrm{O}_{5}$ and $80 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ of $\mathrm{K}_{2} \mathrm{O}$. The cultural practices were conducted equally throughout the experimental area, as recommended for uniformity trials (Storck et al. 2011; Ramalho et al. 2012).

In the experimental area $\left(75 \times 40 \mathrm{~m} ; 3,000 \mathrm{~m}^{2}\right), 3$ areas of $25 \times 40 \mathrm{~m}\left(1,000 \mathrm{~m}^{2}\right)$ were marked. The sowing densities of 40,60 and $80 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ were held in the area 1, 2 and 3, respectively. Sixteen uniformity trials with size of $6 \mathrm{~m} \times 6 \mathrm{~m}$ $\left(36 \mathrm{~m}^{2}\right)$ were demarcated in each area ( $576 \mathrm{~m}^{2}$ of useful area $+424 \mathrm{~m}^{2}$ of border area), totaling 48 uniformity trials ( 3 areas $\times 16$ uniformity trials per area). Each uniformity trial of size $6 \mathrm{~m} \times 6 \mathrm{~m}\left(36 \mathrm{~m}^{2}\right)$ was divided into 36 basic experimental units (BEU) of $1 \mathrm{~m} \times 1 \mathrm{~m}\left(1 \mathrm{~m}^{2}\right)$, forming a matrix with 6 rows and 6 columns. At 125 days after sowing, in the flowering vetch stage, the plants were cut close to the ground in each of the $1,728 \mathrm{BEU}$ ( 3 areas $\times$ 16 uniformity trials per area $\times 36$ BEU per uniformity trial) and the fresh matter weight (in g) was immediately weighed.

For each uniformity trial, the first-order spatial autocorrelation coefficient $(\rho)$, the variance $\left(s^{2}\right)$, the mean ( m ), the coefficient of variation of the trial (CV, in \%), the optimum plot size ( Xo , in $\mathrm{m}^{2}$ ) and the coefficient of variation in the optimum plot size $\left(\mathrm{CV}_{\mathrm{Xo}}\right.$, in \%) were estimated with the fresh matter weight data of the 36 BEU. The $\rho$ estimate was obtained in the rows, according to the methodology of Paranaíba et al. (2009) and application in Cargnelutti Filho et al. (2014a).

Based on the method of maximum curvature of the coefficient of variation model developed by Paranaíba et al. (2009), the optimum plot size

$$
\begin{equation*}
\mathrm{Xo}=\left(10^{3} \sqrt{2\left(1-\rho^{2}\right) \mathrm{s}^{2} \mathrm{~m}}\right) / \mathrm{m} \tag{1}
\end{equation*}
$$

and the coefficient of variation in the optimum plot size
were determined. The means comparisons of the statistics $\rho, s^{2}, \mathrm{~m}, \mathrm{CV}, \mathrm{Xo}$ and $\mathrm{CV}_{\mathrm{Xo}_{\mathrm{o}}}$ among the sowing densities were performed at $5 \%$ probability as follows: initially, the analysis of variance was performed (one-way, i.e. sowing density with 3 levels) via bootstrap with 10,000 resampling. Then, the 3 sowing densities were compared by Scott-Knott's test via bootstrap with 10,000 resampling. In these analyzes, the trials were considered repetitions (independent samples, $\mathrm{n}=16$ trials at each sowing density). Analysis of variance and the Scott-Knott's test via bootstrap were performed with Sisvar software (Ferreira 2014). According to Ferreira (2014), these statistical procedures are adequate to prevent possible impacts of non-compliance of the assumptions of normality of errors and homogeneity of residual variances.

In order to calculate the number of repetitions, it was initiated with the least significant difference (d) of the Tukey's test, expressed as percentage of the experiment mean, which is estimated by the formula:

$$
\begin{equation*}
\mathrm{d}=\left(\mathrm{q}_{\alpha(\mathrm{i}, \mathrm{DFE})} \sqrt{\mathrm{MSE} / \mathrm{r}}\right) / \mathrm{m} \times 100 \tag{3}
\end{equation*}
$$

where: $q_{a(i, \text { DFE })}$ is the critical value of the Tukey's test at $\alpha$ level of error probability ( $\alpha=0.05$, in this study); i is the number of treatments; DFE is the number of degrees of freedom of the error, i.e. $i(r-1)$ for the completely randomized design (CRD) and (i -1$)(r-1)$ for the

RBD; MSE is the mean squared error; $r$ is the number of repetitions; $m$ is the experiment mean.

Thus, by replacing the expression of the coefficient of experimental variation

$$
\begin{equation*}
\mathrm{CV}=\sqrt{\mathrm{MSE}} / \mathrm{m} \times 100 \tag{4}
\end{equation*}
$$

in percentage in the formula for the calculation of $d$ and isolating r , it is obtained:

$$
\begin{equation*}
\mathrm{r}=\left(\mathrm{q}_{a(\mathrm{i} ; \mathrm{DFE})} \mathrm{CV} / \mathrm{d}\right)^{2} \tag{5}
\end{equation*}
$$

In the study, the CV is expressed as a percentage and corresponds to $\mathrm{CV}_{\mathrm{Xo}}$ because this is the expected CV for the experiment with the determined optimum plot size (Xo).

From the average $\mathrm{CV}_{\mathrm{X}_{0}}$ between the 3 sowing densities, the number of repetitions was determined (r) by iterative process until convergence, for experiments in completely randomized and randomized block designs, in scenarios formed by combinations of $\mathrm{i}(\mathrm{i}=3,4, \ldots, 50)$ and $\mathrm{d}[\mathrm{d}=10 \%$ (higher precision), $11 \%, \ldots, 20 \%$ (lower precision)]. Statistical analyzes were performed using Microsoft ${ }^{\circledR}$ Office Excel application and Sisvar software (Ferreira 2014).

## RESULTS AND DISCUSSION

The fresh matter weight of vetch (Vicia sativa L.), common cultivar, oscillated between 22.11 and $27.96 \mathrm{t} \cdot \mathrm{ha}^{-1}$, 20.99 and $26.65 \mathrm{t} \cdot \mathrm{ha}^{-1}$, and 22.22 and $29.57 \mathrm{t} \cdot \mathrm{ha}^{-1}$, respectively, for the sowing densities of 40,60 and $80 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ (Table 1). There was mean difference of fresh matter weight of vetch between the 3 sowing densities. Detailed studies into the causes of these differences were not the focus of this research. The average of fresh matter weight of vetch among 48 uniformity trials (3 sowing densities $\times 16$ uniformity trials per sowing density) was of $24.82 \mathrm{t} \cdot \mathrm{ha}^{-1}$ (Table 1). This average was greater than that obtained by Seidel et al. (2011) (7.40 t•ha ${ }^{-1}$ ), Ansar et al. (2013) (11.13 t•ha ${ }^{-1}$ ), Desalegn and Hassen (2015) (14.833 t $\cdot \mathrm{ha}^{-1}$ ), Basbag et al. (2015) ( $18.32 \mathrm{t} \cdot \mathrm{ha} \mathrm{a}^{-1}$ ), Tuna and Orak (2007) (19.60 t•ha ${ }^{-1}$ ) and Lithourgidis et al. (2006) (20.49 t•ha ${ }^{-1}$ ). Also, it was smaller than that obtained by Cherubin et al. (2014) (25.94 t•ha ${ }^{-1}$ ). Therefore, in view of the high number of uniformity trials (48) in 3 sowing
densities (large variability), allied with the appropriate plant development, it can be inferred that this database is suitable for the proposed study.

Regarding to the fresh matter weight of vetch, there was variability in the estimates of the first-order spatial autocorrelation coefficient $(\rho)$, the variance $\left(s^{2}\right)$, the mean (m) and the coefficient of variation of the trial (CV), among 16 uniformity trials performed on each sowing density ( 40,60 and $80 \mathrm{~kg} \cdot \mathrm{ha}^{-1}$ ) (Table 1). In this way, there was also variability in estimates of the optimum plot size (Xo) and the coefficient of variation in the optimum plot size $\left(\mathrm{CV}_{\mathrm{Xo}_{0}}\right)$, whose estimates are obtained based on $\rho, \mathrm{s}^{2}$ and m (Paranaíba et al. 2009). This wide variability scenario of $\rho, \mathrm{s}^{2}, \mathrm{~m}, \mathrm{CV}, \mathrm{Xo}$ and $\mathrm{CV}_{\mathrm{xo}}$ statistics among 48 uniformity trials is important for the properly dimensioning of plot size and number of repetitions. In field areas, this wide variability has also been observed in other cover crops, such as: black oat (Cargnelutti Filho et al. 2014a), jack bean (Cargnelutti Filho et al. 2014b), forage pea (Cargnelutti Filho et al. 2015a), canola (Cargnelutti Filho et al. 2015b), millet (Burin et al. 2015; Burin et al. 2016) and pigeonpea (Santos et al. 2016).

Among the 3 sowing densities, there was no significant difference between the means $\rho, \mathrm{s}^{2}, \mathrm{CV}, \mathrm{Xo}$ and $\mathrm{CV}_{\mathrm{X}}$ (Table 1). Wherefore, as Xo and $\mathrm{CV}_{\mathrm{Xo}}$ did not differentiate among sowing densities, the experimental design based on the overall average is an appropriate procedure to ensure the desired precision for these sowing densities. Thus, based on the average of 48 uniformity trials, it can be inferred that the optimum plot size to assess the fresh matter weight of vetch is $4.52 \mathrm{~m}^{2}$ and the coefficient of variation in the optimum plot size was $10.10 \%$, which is the base value to calculate the number of repetitions. Similar plot sizes to $4.52 \mathrm{~m}^{2}$ were determined to evaluate the fresh matter weight of other cover crops. The determined sizes were: $4.14 \mathrm{~m}^{2}$ for black oat (Cargnelutti Filho et al. 2014a), $5.85 \mathrm{~m}^{2}$ for jack bean (Cargnelutti Filho et al. 2014b), $5.03 \mathrm{~m}^{2}$ for forage pea (Cargnelutti Filho et al. 2015a), $6.41 \mathrm{~m}^{2}$ for canola (Cargnelutti Filho et al. 2015b), 4.46 and $4.97 \mathrm{~m}^{2}$ for millet (Burin et al. 2015, Burin et al. 2016) and $8.39 \mathrm{~m}^{2}$ for pigeonpea (Santos et al. 2016).

In researches carried out by Lithourgidis et al. (2006), Tuna and Orak (2007) and Desalegn and Hassen (2015) to assess the fresh matter weight of vetch, the authors utilized plot sizes and useful area exceeding $4.52 \mathrm{~m}^{2}$.

Table 1. First-order spatial autocorrelation coefficient, variance, mean, coefficient of variation of the trial, optimum plot size, and coefficient of variation in the optimum plot size for fresh matter weight of vetch (Vicia sativa L.), common cultivar, in $\mathrm{g} \cdot \mathrm{m}^{-2}$, in 48 uniformity trials ( 16 uniformity trials in each sowing density).

| Sowing density (kg•ha ${ }^{-1}$ ) | Trial ${ }^{(1)}$ | $\rho$ | $\mathrm{s}^{2}$ | m | CV (\%) | Xo (m²) | CV $\mathrm{x}_{\mathrm{o}}$ (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 40 | 1 | 0.17 | 227,869.85 | 2,796.25 | 17.07 | 3.84 | 8.59 |
| 40 | 2 | 0.25 | 263,609.06 | 2,629.50 | 19.53 | 4.15 | 9.28 |
| 40 | 3 | 0.38 | 319,116.03 | 2,487.72 | 22.71 | 4.45 | 9.96 |
| 40 | 4 | 0.27 | 212,717.37 | 2,426.06 | 19.01 | 4.06 | 9.09 |
| 40 | 5 | 0.04 | 206,373.54 | 2,514.94 | 18.06 | 4.02 | 9.00 |
| 40 | 6 | 0.09 | 338,038.08 | 2,721.08 | 21.37 | 4.49 | 10.04 |
| 40 | 7 | -0.05 | 310,380.81 | 2,534.14 | 21.98 | 4.59 | 10.25 |
| 40 | 8 | 0.16 | 234,076.41 | 2,507.64 | 19.29 | 4.17 | 9.32 |
| 40 | 9 | 0.25 | 330,475.31 | 2,211.00 | 26.00 | 5.02 | 11.23 |
| 40 | 10 | 0.30 | 391,869.05 | 2,591.56 | 24.16 | 4.73 | 10.58 |
| 40 | 11 | 0.18 | 338,916.99 | 2,420.92 | 24.05 | 4.82 | 10.77 |
| 40 | 12 | 0.49 | 435,981.00 | 2,465.97 | 26.78 | 4.78 | 10.69 |
| 40 | 13 | 0.28 | 422,356.54 | 2,252.17 | 28.86 | 5.35 | 11.97 |
| 40 | 14 | 0.10 | 336,454.88 | 2,641.75 | 21.96 | 4.57 | 10.22 |
| 40 | 15 | 0.09 | 542,589.78 | 2,584.22 | 28.50 | 5.44 | 12.17 |
| 40 | 16 | 0.08 | 267,020.00 | 2,495.00 | 20.71 | 4.40 | 9.84 |
| Mean |  | 0.19 a | 323,615.29 a | 2,517.49 a | 22.50 a | 4.56 a | 10.19 a |
| 60 | 1 | 0.22 | 338,930.83 | 2,159.28 | 26.96 | 5.17 | 11.56 |
| 60 | 2 | 0.19 | 237,371.09 | 2,188.36 | 22.26 | 4.57 | 10.23 |
| 60 | 3 | 0.09 | 333,147.83 | 2,227.06 | 25.92 | 5.11 | 11.42 |
| 60 | 4 | -0.09 | 336,479.54 | 2,347.00 | 24.72 | 4.95 | 11.06 |
| 60 | 5 | 0.13 | 284,329.00 | 2,166.17 | 24.62 | 4.92 | 11.00 |
| 60 | 6 | 0.28 | 238,918.88 | 2,099.53 | 23.28 | 4.64 | 10.38 |
| 60 | 7 | 0.37 | 185,340.25 | 2,169.92 | 19.84 | 4.08 | 9.13 |
| 60 | 8 | -0.09 | 404,442.58 | 2,283.86 | 27.85 | 5.36 | 11.98 |
| 60 | 9 | 0.12 | 250,585.29 | 2,469.50 | 20.27 | 4.33 | 9.68 |
| 60 | 10 | -0.19 | 301,252.46 | 2,383.22 | 23.03 | 4.68 | 10.45 |
| 60 | 11 | -0.04 | 200,178.24 | 2,228.14 | 20.08 | 4.32 | 9.66 |
| 60 | 12 | 0.14 | 262,892.66 | 2,522.50 | 20.33 | 4.33 | 9.68 |
| 60 | 13 | 0.26 | 215,495.50 | 2,614.61 | 17.75 | 3.89 | 8.69 |
| 60 | 14 | 0.39 | 298,482.20 | 2,316.83 | 23.58 | 4.55 | 10.17 |
| 60 | 15 | 0.07 | 322,020.23 | 2,386.94 | 23.77 | 4.83 | 10.80 |
| 60 | 16 | 0.00 | 268,033.23 | 2,665.53 | 19.42 | 4.23 | 9.45 |
| Mean |  | 0.12 a | 279,868.74 a | 2,326.78 b | 22.73 a | 4.62 a | 10.33 a |
| 80 | 1 | 0.34 | 238,849.28 | 2,626.92 | 18.60 | 3.94 | 8.82 |
| 80 | 2 | 0.08 | 209,582.33 | 2,643.19 | 17.32 | 3.91 | 8.74 |
| 80 | 3 | 0.30 | 269,556.48 | 2,222.92 | 23.36 | 4.63 | 10.35 |
| 80 | 4 | 0.03 | 189,122.03 | 2,673.03 | 16.27 | 3.75 | 8.39 |
| 80 | 5 | 0.13 | 257,218.65 | 2,564.92 | 19.77 | 4.25 | 9.51 |
| 80 | 6 | -0.37 | 630,466.70 | 2,957.61 | 26.85 | 4.99 | 11.15 |

Table 1. Continuation...

| Sowing density (kg•ha-1) | Trial ${ }^{(1)}$ | $\rho$ | $\mathrm{s}^{2}$ | m | CV (\%) | Xo (m²) | CV $\mathrm{x}_{0}$ (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 80 | 7 | 0.51 | 521,833.94 | 2,422.33 | 29.82 | 5.10 | 11.40 |
| 80 | 8 | 0.20 | 237,172.96 | 2,484.81 | 19.60 | 4.19 | 9.37 |
| 80 | 9 | -0.19 | 235,899.26 | 2,681.33 | 18.11 | 3.98 | 8.91 |
| 80 | 10 | -0.15 | 475,498.61 | 2,854.72 | 24.16 | 4.85 | 10.85 |
| 80 | 11 | 0.26 | 227,667.28 | 2,513.75 | 18.98 | 4.06 | 9.08 |
| 80 | 12 | 0.06 | 270,870.35 | 2,689.64 | 19.35 | 4.21 | 9.41 |
| 80 | 13 | -0.41 | 312,779.69 | 2,568.17 | 21.78 | 4.29 | 9.60 |
| 80 | 14 | 0.00 | 458,435.45 | 2,590.92 | 26.13 | 5.15 | 11.52 |
| 80 | 15 | -0.10 | 262,654.18 | 2,573.86 | 19.91 | 4.28 | 9.58 |
| 80 | 16 | 0.35 | 325,484.35 | 2,593.22 | 22.00 | 4.40 | 9.83 |
| Mean |  | 0.07 a | 320,193.22 a | 2,603.83 a | 21.38 a | 4.37 a | 9.78 a |
| Overall mean |  | 0.12 | 307,892.42 | 2,482.70 | 22.20 | 4.52 | 10.10 |

${ }^{(1)}$ Each uniformity trial of size $6 \mathrm{~m} \times 6 \mathrm{~m}\left(36 \mathrm{~m}^{2}\right)$ was divided into 36 basic experimental units of $1 \mathrm{~m} \times 1 \mathrm{~m}\left(1 \mathrm{~m}^{2}\right)$, forming a matrix with 6 rows and 6 columns; ${ }^{(2)}$ For each statistics ( $\rho, \mathrm{s}^{2}, \mathrm{~m}, \mathrm{CV}, \mathrm{Xo}$, and $\mathrm{CV}_{\mathrm{x}_{0}}$ ), the means not followed by the same letter in column (comparison of means between sowing densities) differ at a $5 \%$ probability by Scott-Knott's test via bootstrap with 10,000 resampling. $\rho=$ First-order spatial autocorrelation coefficient; $\mathrm{s}^{2}=$ Variance; $\mathrm{m}=\mathrm{Mean}$; $\mathrm{CV}=$ Coefficient of variation of the trial; $\mathrm{Xo}=$ Optimum plot size; $\mathrm{CV}_{\mathrm{x}_{0}}=$ Coefficient of variation in the optimum plot size.

The plot sizes utilized by Ansar et al. (2013) ( $15 \mathrm{~m}^{2}$ ) and by Basbag et al. (2015) ( $7.2 \mathrm{~m}^{2}$ ) were also higher than $4.52 \mathrm{~m}^{2}$. However, the samples for evaluation of fresh matter weight of vetch were collected in areas smaller than $4.52 \mathrm{~m}^{2}$. On the other hand, Seidel et al. (2011) and Cherubin et al. (2014) did not report the plot size. Although, the authors utilized $0.50 \mathrm{~m}^{2}$ (Seidel et al. 2011) and $0.75 \mathrm{~m}^{2}$ per plot (Cherubin et al. 2014) to evaluate the fresh matter weight, i.e. below the optimal plot size obtained in this study $\left(4.52 \mathrm{~m}^{2}\right)$.

The number of repetitions ( r ) of plots with $4.52 \mathrm{~m}^{2}$ for fixed number of treatments gradually increases with the decreasing of the least significant difference (d) (precision increase) in CRDs and RBDs (Tables 2,3). Smaller d estimates enable identifying minor differences between treatment means as significant and it signifies higher experimental precision (Lúcio et al. 1999; Storck et al. 2011). For least significant difference fixed percentage (d), the number of repetitions (r) of plots with $4.52 \mathrm{~m}^{2}$ increased with increasing number of treatments, regardless of the design (Tables 2,3). Wherefore, these results indicate that, for fixed plot size, the higher the required precision and the higher the number of treatments to be evaluated, the larger the number of repetitions.

For experiments in CRD, the number of repetitions (r) ranged from 3.98 for 3 treatments and precision of $20 \%$ (lower precision) to 32.66 repetitions for 50
treatments and precision of $10 \%$ (higher precision) (Table 2). Moreover, for the experiments in RBDs, the number of repetitions oscillated from 4.42 ( 3 treatments and $\mathrm{d}=20 \%$ ) to 32.66 ( 50 treatments and $\mathrm{d}=10 \%$ ) (Table 3). Therefore, regardless of the experimental design, obtaining precision of $10 \%$ (higher precision) is impractical due to the elevated number of repetitions required. However, it is possible to establish the proper relations between i , d and the number of repetitions based on $\mathrm{Xo}=4.52 \mathrm{~m}^{2}$.

In practice, to carry out the experiment, it is necessary to define the integer number of repetitions. Then, fixing $r$ equal to 4 repetitions, as used in the experiments performed by Lithourgidis et al. (2006), Seidel et al. (2011) and Basbag et al. (2015), the least significant difference (d) of the Tukey's test is estimated by the expression:

$$
\begin{equation*}
\mathrm{d}=\left(\mathrm{q}_{\alpha(\mathrm{i}, \mathrm{DFE})} \mathrm{CV}\right) / \sqrt{\mathrm{r}} \tag{6}
\end{equation*}
$$

expressed as percentage of the experiment average. Thus, with 20 treatments (number of usual treatments in experiments), for the CRD,

$$
\begin{align*}
\mathrm{d}= & \left(\mathrm{q}_{5 \%(20 ; 60)} \times 10.10\right) / \sqrt{4}=  \tag{7}\\
& (5.2411998 \times 10.10) / \sqrt{4}=26.47 \%
\end{align*}
$$

and, for the RBD,

$$
\begin{align*}
\mathrm{d}= & \left(\mathrm{q}_{5 \%(20 ; 57)} \times 10.10\right) / \sqrt{4}=  \tag{8}\\
& (5.2533923 \times 10.10) / \sqrt{4}=26.53 \%
\end{align*}
$$

With 50 treatments (significant number of treatments), for the CRD,

$$
\begin{align*}
d= & \left(q_{5 \%(50 ; 150)} \times 10.10\right) / \sqrt{4}=  \tag{9}\\
& (5.7710567 \times 10.10) / \sqrt{4}=29.14 \%
\end{align*}
$$

and, for the RBD,

$$
\begin{align*}
\mathrm{d}= & \left(\mathrm{q}_{5 \%(50 ; 147)} \times 10.10\right) / \sqrt{4}=  \tag{10}\\
& (5.7736023 \times 10.10) / \sqrt{4}=29.15 \%
\end{align*}
$$

Then, it can be inferred that, to evaluate the fresh matter weight of vetch, in CRD and RBD with up to 50 treatments, 4 repetitions are enough to identify as significant, at 5\% probability by Tukey's test, differences between treatment means of $29.15 \%$ of the overall experiment average.

In order to evaluate the fresh matter weight of other cover crops, in the same scenario, i.e. for experiments in CRD and RBD with 50 treatments and 4 repetitions, the least significant difference (d) varied among crops. Compared to $d=29.15 \%$ for vetch obtained in this study, lower values (higher precision) were obtained by Cargnelutti Filho et al. (2014a) for black oat ( $\mathrm{d}=26.7 \%$ ), by Burin et al. (2016) for millet in sowing and cut times ( $\mathrm{d}=28.66 \%$ ), and by Burin et al. (2015) for millet in evaluation times ( $\mathrm{d}=28.75 \%$ ). On the other hand, higher values (lower precision) were obtained by Cargnelutti Filho et al. (2015a) for forage pea $(\mathrm{d}=32.4 \%)$, by Cargnelutti Filho et al. (2014b) for jack bean $(\mathrm{d}=37.78 \%)$, by Cargnelutti Filho et al. (2015b) for canola ( $\mathrm{d}=41.4 \%$ ) and by Santos et al. (2016) for pigeonpea ( $\mathrm{d}=54.1 \%$ ).

By definition, in any experiment which assesses fixed effect of treatments, the conclusions are valid for the conditions under which the experiment was carried out (Storck et al. 2011). Then, the experimental plan (plot size and number of repetitions) for a certain number of treatments, required precision, and experimental design

Table 2. Number of repetitions, for experiments in completely randomized designs, in scenarios formed by combinations of $i$ treatments ( $\mathrm{i}=3$, $4, \ldots, 50$ ) and d minimal differences between treatments means being detected as significant with $5 \%$ probability by Tukey's test, expressed in percentage of the experiment average ( $\mathrm{d}=10 \%, 11 \%, \ldots, 20 \%$ ), to evaluate the fresh matter weight of vetch (Vicia sativa L.), common cultivar, based on the optimum plot size ( $\mathrm{Xo}=4.52 \mathrm{~m}^{2}$ ) and coefficient of variation in the optimum plot size $\left(\mathrm{CV}_{\mathrm{X}_{0}}=10.10 \%\right)$.

| $\mathbf{i}$ | $\mathbf{1 0 \%}$ | $\mathbf{1 1 \%}$ | $\mathbf{1 2 \%}$ | $\mathbf{1 3 \%}$ | $\mathbf{1 4 \%}$ | $\mathbf{1 5 \%}$ | $\mathbf{1 6 \%}$ | $\mathbf{1 7 \%}$ | $\mathbf{1 8 \%}$ | $\mathbf{1 9 \%}$ | $\mathbf{2 0 \%}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| $\mathbf{3}$ | 12.26 | 10.32 | 8.86 | 7.72 | 6.81 | 6.09 | 5.50 | 5.01 | 4.61 | 4.26 | 3.98 |
| $\mathbf{4}$ | 14.35 | 12.02 | 10.25 | 8.87 | 7.79 | 6.91 | 6.20 | 5.61 | 5.12 | 4.70 | 4.35 |
| $\mathbf{5}$ | 15.94 | 13.32 | 11.32 | 9.77 | 8.54 | 7.55 | 6.74 | 6.07 | 5.52 | 5.05 | 4.65 |
| $\mathbf{6}$ | 17.24 | 14.37 | 12.19 | 10.50 | 9.15 | 8.07 | 7.19 | 6.46 | 5.85 | 5.33 | 4.90 |
| $\mathbf{7}$ | 18.34 | 15.27 | 12.93 | 11.12 | 9.68 | 8.52 | 7.57 | 6.79 | 6.13 | 5.58 | 5.11 |
| $\mathbf{8}$ | 19.29 | 16.05 | 13.58 | 11.66 | 10.14 | 8.91 | 7.91 | 7.08 | 6.38 | 5.80 | 5.30 |
| $\mathbf{9}$ | 20.14 | 16.73 | 14.15 | 12.14 | 10.54 | 9.26 | 8.21 | 7.34 | 6.61 | 6.00 | 5.47 |
| $\mathbf{1 0}$ | 20.89 | 17.35 | 14.66 | 12.57 | 10.91 | 9.57 | 8.48 | 7.57 | 6.81 | 6.17 | 5.63 |
| $\mathbf{1 1}$ | 21.58 | 17.91 | 15.12 | 12.96 | 11.24 | 9.85 | 8.72 | 7.78 | 7.00 | 6.34 | 5.77 |
| $\mathbf{1 2}$ | 22.20 | 18.42 | 15.55 | 13.32 | 11.54 | 10.12 | 8.95 | 7.98 | 7.17 | 6.49 | 5.91 |
| $\mathbf{1 3}$ | 22.78 | 18.90 | 15.94 | 13.65 | 11.83 | 10.36 | 9.16 | 8.16 | 7.33 | 6.63 | 6.03 |
| $\mathbf{1 4}$ | 23.31 | 19.33 | 16.31 | 13.95 | 12.09 | 10.58 | 9.35 | 8.33 | 7.48 | 6.76 | 6.14 |
| $\mathbf{1 5}$ | 23.81 | 19.74 | 16.65 | 14.24 | 12.33 | 10.79 | 9.53 | 8.49 | 7.62 | 6.88 | 6.25 |
| $\mathbf{1 6}$ | 24.28 | 20.13 | 16.97 | 14.51 | 12.56 | 10.99 | 9.71 | 8.64 | 7.75 | 7.00 | 6.36 |
| $\mathbf{1 7}$ | 24.72 | 20.49 | 17.27 | 14.76 | 12.78 | 11.18 | 9.87 | 8.78 | 7.87 | 7.11 | 6.45 |
| $\mathbf{1 8}$ | 25.14 | 20.83 | 17.55 | 15.00 | 12.98 | 11.35 | 10.02 | 8.92 | 7.99 | 7.21 | 6.54 |
| $\mathbf{1 9}$ | 25.53 | 21.15 | 17.82 | 15.23 | 13.18 | 11.52 | 10.17 | 9.04 | 8.10 | 7.31 | 6.63 |
| $\mathbf{2 0}$ | 25.90 | 21.46 | 18.08 | 15.45 | 13.36 | 11.68 | 10.30 | 9.16 | 8.21 | 7.40 | 6.72 |
| $\mathbf{2 1}$ | 26.26 | 21.75 | 18.32 | 15.65 | 13.54 | 11.83 | 10.44 | 9.28 | 8.31 | 7.49 | 6.80 |

Table 2. Continuation..

| i | 10\% | 11\% | 12\% | 13\% | 14\% | 15\% | 16\% | 17\% | 18\% | 19\% | 20\% |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 22 | 26.60 | 22.03 | 18.56 | 15.85 | 13.71 | 11.98 | 10.56 | 9.39 | 8.41 | 7.58 | 6.87 |
| 23 | 26.92 | 22.30 | 18.78 | 16.04 | 13.87 | 12.12 | 10.68 | 9.50 | 8.50 | 7.66 | 6.95 |
| 24 | 27.24 | 22.55 | 18.99 | 16.22 | 14.02 | 12.25 | 10.80 | 9.60 | 8.59 | 7.74 | 7.02 |
| 25 | 27.53 | 22.80 | 19.20 | 16.39 | 14.17 | 12.38 | 10.91 | 9.70 | 8.68 | 7.82 | 7.08 |
| 26 | 27.82 | 23.03 | 19.39 | 16.56 | 14.31 | 12.50 | 11.02 | 9.79 | 8.76 | 7.89 | 7.15 |
| 27 | 28.10 | 23.26 | 19.58 | 16.72 | 14.45 | 12.62 | 11.12 | 9.88 | 8.84 | 7.96 | 7.21 |
| 28 | 28.37 | 23.48 | 19.77 | 16.88 | 14.59 | 12.74 | 11.22 | 9.97 | 8.92 | 8.03 | 7.28 |
| 29 | 28.62 | 23.69 | 19.95 | 17.03 | 14.71 | 12.85 | 11.32 | 10.06 | 9.00 | 8.10 | 7.33 |
| 30 | 28.87 | 23.90 | 20.12 | 17.17 | 14.84 | 12.96 | 11.41 | 10.14 | 9.07 | 8.16 | 7.39 |
| 31 | 29.12 | 24.10 | 20.28 | 17.31 | 14.96 | 13.06 | 11.51 | 10.22 | 9.14 | 8.23 | 7.45 |
| 32 | 29.35 | 24.29 | 20.44 | 17.45 | 15.08 | 13.16 | 11.59 | 10.29 | 9.21 | 8.29 | 7.50 |
| 33 | 29.58 | 24.48 | 20.60 | 17.58 | 15.19 | 13.26 | 11.68 | 10.37 | 9.27 | 8.35 | 7.56 |
| 34 | 29.80 | 24.66 | 20.75 | 17.71 | 15.30 | 13.35 | 11.76 | 10.44 | 9.34 | 8.40 | 7.61 |
| 35 | 30.01 | 24.83 | 20.90 | 17.84 | 15.41 | 13.45 | 11.84 | 10.51 | 9.40 | 8.46 | 7.66 |
| 36 | 30.22 | 25.01 | 21.04 | 17.96 | 15.51 | 13.54 | 11.92 | 10.58 | 9.46 | 8.51 | 7.71 |
| 37 | 30.42 | 25.17 | 21.18 | 18.08 | 15.61 | 13.62 | 12.00 | 10.65 | 9.52 | 8.57 | 7.75 |
| 38 | 30.62 | 25.34 | 21.32 | 18.19 | 15.71 | 13.71 | 12.07 | 10.72 | 9.58 | 8.62 | 7.80 |
| 39 | 30.81 | 25.49 | 21.45 | 18.30 | 15.81 | 13.79 | 12.15 | 10.78 | 9.64 | 8.67 | 7.84 |
| 40 | 31.00 | 25.65 | 21.58 | 18.41 | 15.90 | 13.87 | 12.22 | 10.84 | 9.69 | 8.72 | 7.89 |
| 41 | 31.18 | 25.80 | 21.70 | 18.52 | 15.99 | 13.95 | 12.29 | 10.90 | 9.75 | 8.77 | 7.93 |
| 42 | 31.36 | 25.95 | 21.83 | 18.62 | 16.08 | 14.03 | 12.35 | 10.96 | 9.80 | 8.81 | 7.97 |
| 43 | 31.53 | 26.09 | 21.95 | 18.73 | 16.17 | 14.11 | 12.42 | 11.02 | 9.85 | 8.86 | 8.01 |
| 44 | 31.71 | 26.23 | 22.07 | 18.83 | 16.25 | 14.18 | 12.48 | 11.08 | 9.90 | 8.90 | 8.05 |
| 45 | 31.87 | 26.37 | 22.18 | 18.92 | 16.34 | 14.25 | 12.55 | 11.13 | 9.95 | 8.95 | 8.09 |
| 46 | 32.04 | 26.50 | 22.29 | 19.02 | 16.42 | 14.32 | 12.61 | 11.19 | 10.00 | 8.99 | 8.13 |
| 47 | 32.20 | 26.63 | 22.40 | 19.11 | 16.50 | 14.39 | 12.67 | 11.24 | 10.05 | 9.03 | 8.17 |
| 48 | 32.35 | 26.76 | 22.51 | 19.20 | 16.58 | 14.46 | 12.73 | 11.29 | 10.09 | 9.08 | 8.21 |
| 49 | 32.50 | 26.89 | 22.62 | 19.29 | 16.66 | 14.53 | 12.79 | 11.35 | 10.14 | 9.12 | 8.24 |
| 50 | 32.66 | 27.01 | 22.72 | 19.38 | 16.73 | 14.59 | 12.85 | 11.40 | 10.18 | 9.16 | 8.28 |

Table 3. Number of repetitions, for experiments in randomized block designs, in scenarios formed by combinations of $i$ treatments ( $i=3$, $4, \ldots, 50$ ) and d minimal differences between treatments means being detected as significant with $5 \%$ probability by Tukey's test, expressed in percentage of the experiment average ( $\mathrm{d}=10 \%, 11 \%, \ldots, 20 \%$ ), to evaluate the fresh matter weight of vetch (Vicia sativa L.), common cultivar, based on the optimum plot size $\left(X o=4.52 \mathrm{~m}^{2}\right)$ and coefficient of variation in the optimum plot size $\left(\mathrm{CV}_{\mathrm{xo}}=10.10 \%\right)$.

| $\mathbf{i}$ | $\mathbf{1 0 \%}$ | $\mathbf{1 1 \%}$ | $\mathbf{1 2 \%}$ | $\mathbf{1 3 \%}$ | $\mathbf{1 4 \%}$ | $\mathbf{1 5 \%}$ | $\mathbf{1 6 \%}$ | $\mathbf{1 7 \%}$ | $\mathbf{1 8 \%}$ | $\mathbf{1 9 \%}$ | $\mathbf{2 0 \%}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :--- |
| $\mathbf{3}$ | 12.75 | 10.82 | 9.34 | 8.20 | 7.30 | 6.58 | 5.99 | 5.48 | 5.09 | 4.72 | 4.42 |
| $\mathbf{4}$ | 14.63 | 12.30 | 10.53 | 9.16 | 8.07 | 7.20 | 6.48 | 5.89 | 5.39 | 4.98 | 4.62 |
| $\mathbf{5}$ | 16.13 | 13.50 | 11.50 | 9.95 | 8.72 | 7.73 | 6.93 | 6.26 | 5.70 | 5.23 | 4.83 |
| $\mathbf{6}$ | 17.38 | 14.51 | 12.32 | 10.63 | 9.29 | 8.21 | 7.32 | 6.59 | 5.98 | 5.47 | 5.03 |
| $\mathbf{7}$ | 18.44 | 15.37 | 13.03 | 11.22 | 9.78 | 8.62 | 7.67 | 6.89 | 6.23 | 5.68 | 5.21 |
| $\mathbf{8}$ | 19.37 | 16.12 | 13.66 | 11.74 | 10.21 | 8.99 | 7.99 | 7.16 | 6.46 | 5.88 | 5.38 |
| $\mathbf{9}$ | 20.20 | 16.80 | 14.21 | 12.20 | 10.60 | 9.32 | 8.27 | 7.40 | 6.67 | 6.06 | 5.54 |
| $\mathbf{1 0}$ | 20.94 | 17.40 | 14.71 | 12.62 | 10.96 | 9.62 | 8.53 | 7.62 | 6.87 | 6.23 | 5.68 |
| $\mathbf{1 1}$ | 21.62 | 17.95 | 15.17 | 13.00 | 11.28 | 9.90 | 8.77 | 7.83 | 7.04 | 6.38 | 5.82 |

Table 3. Continuation..

| i | 10\% | 11\% | 12\% | 13\% | 14\% | 15\% | 16\% | 17\% | 18\% | 19\% | 20\% |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 12 | 22.24 | 18.46 | 15.59 | 13.35 | 11.58 | 10.15 | 8.99 | 8.02 | 7.21 | 6.53 | 5.94 |
| 13 | 22.81 | 18.93 | 15.98 | 13.68 | 11.86 | 10.39 | 9.19 | 8.20 | 7.36 | 6.66 | 6.06 |
| 14 | 23.34 | 19.36 | 16.34 | 13.98 | 12.12 | 10.61 | 9.38 | 8.36 | 7.51 | 6.79 | 6.17 |
| 15 | 23.84 | 19.77 | 16.67 | 14.27 | 12.36 | 10.82 | 9.56 | 8.52 | 7.64 | 6.91 | 6.28 |
| 16 | 24.30 | 20.15 | 16.99 | 14.53 | 12.58 | 11.01 | 9.73 | 8.66 | 7.77 | 7.02 | 6.38 |
| 17 | 24.74 | 20.51 | 17.29 | 14.78 | 12.80 | 11.20 | 9.89 | 8.80 | 7.90 | 7.13 | 6.47 |
| 18 | 25.15 | 20.85 | 17.57 | 15.02 | 13.00 | 11.37 | 10.04 | 8.94 | 8.01 | 7.23 | 6.56 |
| 19 | 25.55 | 21.17 | 17.84 | 15.25 | 13.19 | 11.54 | 10.18 | 9.06 | 8.12 | 7.33 | 6.65 |
| 20 | 25.92 | 21.47 | 18.09 | 15.46 | 13.38 | 11.70 | 10.32 | 9.18 | 8.23 | 7.42 | 6.73 |
| 21 | 26.27 | 21.76 | 18.34 | 15.67 | 13.55 | 11.85 | 10.45 | 9.29 | 8.33 | 7.51 | 6.81 |
| 22 | 26.61 | 22.04 | 18.57 | 15.86 | 13.72 | 11.99 | 10.58 | 9.40 | 8.42 | 7.59 | 6.89 |
| 23 | 26.94 | 22.31 | 18.79 | 16.05 | 13.88 | 12.13 | 10.70 | 9.51 | 8.52 | 7.67 | 6.96 |
| 24 | 27.25 | 22.56 | 19.00 | 16.23 | 14.03 | 12.26 | 10.81 | 9.61 | 8.60 | 7.75 | 7.03 |
| 25 | 27.54 | 22.81 | 19.21 | 16.40 | 14.18 | 12.39 | 10.92 | 9.71 | 8.69 | 7.83 | 7.10 |
| 26 | 27.83 | 23.04 | 19.40 | 16.57 | 14.32 | 12.51 | 11.03 | 9.80 | 8.77 | 7.90 | 7.16 |
| 27 | 28.11 | 23.27 | 19.59 | 16.73 | 14.46 | 12.63 | 11.13 | 9.89 | 8.85 | 7.97 | 7.22 |
| 28 | 28.37 | 23.49 | 19.78 | 16.89 | 14.59 | 12.74 | 11.23 | 9.98 | 8.93 | 8.04 | 7.28 |
| 29 | 28.63 | 23.70 | 19.95 | 17.04 | 14.72 | 12.86 | 11.33 | 10.06 | 9.00 | 8.11 | 7.34 |
| 30 | 28.88 | 23.91 | 20.12 | 17.18 | 14.85 | 12.96 | 11.42 | 10.14 | 9.08 | 8.17 | 7.40 |
| 31 | 29.12 | 24.11 | 20.29 | 17.32 | 14.97 | 13.07 | 11.51 | 10.22 | 9.15 | 8.23 | 7.46 |
| 32 | 29.36 | 24.30 | 20.45 | 17.46 | 15.08 | 13.17 | 11.60 | 10.30 | 9.21 | 8.29 | 7.51 |
| 33 | 29.58 | 24.48 | 20.61 | 17.59 | 15.20 | 13.26 | 11.68 | 10.38 | 9.28 | 8.35 | 7.56 |
| 34 | 29.80 | 24.66 | 20.76 | 17.72 | 15.30 | 13.36 | 11.77 | 10.45 | 9.34 | 8.41 | 7.61 |
| 35 | 30.02 | 24.84 | 20.90 | 17.84 | 15.41 | 13.45 | 11.85 | 10.52 | 9.41 | 8.47 | 7.66 |
| 36 | 30.22 | 25.01 | 21.05 | 17.96 | 15.52 | 13.54 | 11.93 | 10.59 | 9.47 | 8.52 | 7.71 |
| 37 | 30.43 | 25.18 | 21.19 | 18.08 | 15.62 | 13.63 | 12.00 | 10.66 | 9.53 | 8.57 | 7.76 |
| 38 | 30.62 | 25.34 | 21.32 | 18.20 | 15.71 | 13.71 | 12.08 | 10.72 | 9.58 | 8.62 | 7.80 |
| 39 | 30.82 | 25.50 | 21.45 | 18.31 | 15.81 | 13.80 | 12.15 | 10.78 | 9.64 | 8.67 | 7.85 |
| 40 | 31.00 | 25.65 | 21.58 | 18.42 | 15.90 | 13.88 | 12.22 | 10.85 | 9.70 | 8.72 | 7.89 |
| 41 | 31.19 | 25.80 | 21.71 | 18.52 | 16.00 | 13.96 | 12.29 | 10.91 | 9.75 | 8.77 | 7.93 |
| 42 | 31.36 | 25.95 | 21.83 | 18.63 | 16.09 | 14.04 | 12.36 | 10.97 | 9.80 | 8.82 | 7.98 |
| 43 | 31.54 | 26.09 | 21.95 | 18.73 | 16.17 | 14.11 | 12.42 | 11.03 | 9.85 | 8.86 | 8.02 |
| 44 | 31.71 | 26.23 | 22.07 | 18.83 | 16.26 | 14.18 | 12.49 | 11.08 | 9.90 | 8.91 | 8.06 |
| 45 | 31.88 | 26.37 | 22.18 | 18.93 | 16.34 | 14.26 | 12.55 | 11.14 | 9.95 | 8.95 | 8.10 |
| 46 | 32.04 | 26.51 | 22.30 | 19.02 | 16.42 | 14.33 | 12.61 | 11.19 | 10.00 | 8.99 | 8.14 |
| 47 | 32.20 | 26.64 | 22.41 | 19.12 | 16.50 | 14.40 | 12.67 | 11.25 | 10.05 | 9.04 | 8.17 |
| 48 | 32.35 | 26.77 | 22.51 | 19.21 | 16.58 | 14.47 | 12.73 | 11.30 | 10.10 | 9.08 | 8.21 |
| 49 | 32.51 | 26.89 | 22.62 | 19.30 | 16.66 | 14.53 | 12.79 | 11.35 | 10.14 | 9.12 | 8.25 |
| 50 | 32.66 | 27.02 | 22.72 | 19.38 | 16.73 | 14.60 | 12.85 | 11.40 | 10.19 | 9.16 | 8.28 |

is valid for the location of the uniformity trial. However, considering the lack of research in this sense for the vetch crop, the information provided in this study is extremely important as a reference point for future experiments with vetch. Generalized conclusions for vetch crop may be performed from more uniformity trials with variation of other factors, such as: location, cultivar, species, sowing and harvest season.

## CONCLUSION

The optimum plot size to evaluate the fresh matter weight of vetch is $4.52 \mathrm{~m}^{2}$ at the 3 sowing densities.

Four repetitions, to evaluate up to 50 treatments, in completely randomized and randomized block designs, are enough to identify as significant, at 5\% probability by Tukey's test, differences between treatment means of $29.15 \%$ of the experiment average.

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