

Infant malnutrition and obesity in three population-based birth cohort studies in Southern Brazil: trends and differences

Desnutrição e obesidade infantis em três coortes de nascimentos no Sul do Brasil: tendências e diferenças

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Abstract

The aim of this study was to compare the evolution of nutritional deficits and overweight in one-year-old children from three birth cohorts started in 1982, 1993 and 2004 in Pelotas, Southern Brazil. Samples from the 1982 and 1993 cohorts and all children from 2004 were weighed and measured, and their mothers interviewed. Anthropometric deficits and overweight were assessed using both NCHS and WHO growth standards. A comparison of the existence of nutritional deficits showed that, after a decline between 1982 and 1993, its prevalence stabilized between 1993 and 2004. Across the whole period, a decrease in all deficits was observed. Obesity, on the other hand, increased. A deficit in the ratio of body length to age was found to be strongly associated with family income. The group with income below one minimum wage was the only to present a significant reduction of stunting during the study period. The most significant improvements in the reduction of nutritional deficits occurred in the first half of the study period, while social differentials remained. Fighting malnutrition is still necessary among the 40% of the population considered poor, and must be accompanied by efforts to combat overweight which is being observed in all social strata.

Malnutrition; Overweight; Anthropometry; Cohort Studies

Introduction

Infant malnutrition has been declining in almost all regions of the globe, a phenomenon closely followed by a concomitant increase in obesity. A global study of trends in malnutrition, measured as height-for-age deficits, showed an overall tendency towards reduction, with West and East Africa and Central America being the exceptions. One of the strongest decreasing trends in height-for-age deficits worldwide was detected in South America, where estimated prevalence fell from almost 25% in 1980 to below 5% in 2005 ¹.

Overweight, on the other hand, is clearly on the rise, and several studies of the adult population point to an alarming scenario ^{2,3,4}. Although studies involving children are less frequent, estimates from an international survey suggest that Latin America is the region with the greatest prevalence of overweight among children under the age of five ⁵. Moreover, in South America the study estimated that 4.9% of children had z-scores for weight-for-length higher than 2 standard deviations. In Brazil, earlier studies carried out in the city of Pelotas, Southern Brazil, using the same criteria had already indicated an increased prevalence of overweight, which reached 8.1% of children in 1993 ⁶.

The present article evaluates the changes in anthropometric indicators that took place between 1982, 1993, and 2004 in the city of Pelotas, using a study design that has never before used

in Brazil. In each of these years, either all the city's children or representative samples were studied at birth and followed throughout their first year of life. The study therefore allowed us to compare prevalence rates for nutrition deficits and overweight across a 22-year period. We also present data on how nutritional status varied according to socioeconomic status, sex, and birth weight.

Methodology

Data collection in the three cohorts was carried out using similar methods, which allowed results from the three studies to be compared directly ⁷. The 1982 and 1993 studies began with the identification of hospital births throughout the entire year – totaling 6,011 and 5,304 births, respectively – and proceeded with follow-ups of samples of these children. The 2004 study began in a similar fashion, including 4,287 births, but attempted to locate all cohort subjects at ages three months and 12 months.

The first follow-up of the 1982 cohort took place in early 1983. This follow-up included children born from January to April 1982, who were then aged between 9 and 15 months (54% were in the 11-13 months age range).

For the 1993 cohort, on the other hand, we sought to evaluate at age 12 months a systematic sample of about 20% of all live births, in addition to all children born weighing under 2,500g, visiting a total of 1,364 children. In this follow-up, 1,347 children (99%) were in the 11-13 months age range. In the present analyses, data are weighted so that our results reflect the prevalence of malnutrition and overweight in the population.

The 12-month evaluation of the 2004 cohort included all subjects. Visits were planned so that children were seen within a window of at most 7 days before or after their first birthday. We were able to carry out 93% of visits within this window.

The anthropometric evaluations described in the present article were carried out at around 12 months. Weight was measured by trained interviewers, whenever possible with the child undressed. In case the mother did not allow the child to be totally undressed, any clothing worn was noted, and had its weight subsequently deducted from the measured weight. In the 1982 and 1993 cohorts, children were weighed using Salter CMS mechanical scales (Salter, Tonbridge, United Kingdom) with a 25kg maximum and 100g precision. In the 2004 cohort, children were weighed together with their mothers after the mother had been weighed alone, and the weight was calculated as the difference be-

tween the two measures. Tanita electronic scales (Tanita, Tokyo, Japan) with a 150kg maximum and 100g precision were used in this study. In all studies, scales were calibrated on a weekly basis using standard weights. Length was measured for all children, regardless of age, in supine position using an AHRTAG portable infantometer with 1mm precision, that was custom built from wood for the purpose of these studies. In all three cohorts, interviewers underwent training and standardization in anthropometric techniques, and 5-10% of interviews were repeated by supervisors.

Based on the collected data, we calculated z-scores for length-for-age, weight-for-age, and weight-for-length, according to the United States National Center for Health Statistics (NCHS) reference ⁸. The same anthropometric indicators were recalculated based on the growth curves published by the World Health Organization (WHO) in 2006 ⁹, using Anthro 2005 software, which can be downloaded from the WHO website (<http://www.who.int/childgrowth/software/en/>). Children with z-scores below -2 for any of the indicators were considered to be underweight. Overweight was defined as weight-for-length z-score above +2 standard deviations. Weight gain in the first year was calculated as the difference between weight at 12 months and birth weight.

The explanatory variables used in the present analysis were: sex, birth weight, and family income. The child's sex was recorded based on data collected soon after birth, while the mother was still in hospital. Birth weight was recorded at the same time, based on the measurement made by the nurse assisting the delivery using a scale whose accuracy was checked on a weekly basis by the interviewer. Family income was self-reported by the mother at every subsequent visit, and was recorded as a multiple of the current minimum wage.

Statistical analysis included calculating the prevalence of nutritional deficits and overweight, as well as prevalence ratios for comparing different subgroups. Statistical comparisons between groups were based on χ^2 tests for heterogeneity in the case of nominal variables and linear trend tests in the case of ordinal variables. To compare mean weight gain, we used a linear regression model adjusting for cohort effect, family income, and the interaction between these two variables.

All analyses were performed using the Stata statistical package (Stata Corp., College Station, USA).

The study protocol was approved by the Medical Ethics Committee of the Federal University of Pelotas. In 1982 and 1993, verbal consent was

obtained from mothers for participation in the study. In 2004, in addition, written consent was also requested.

Results

In the 1982 cohort, we were able to trace 79% of children born between January and April of 1982. The mean age of these children was 11.2 ± 1.3 months. In 1993, the follow-up rate was 93% and the mean age was 12.2 ± 0.5 months. In 2004, the follow-up rate was 94% and the mean age 11.9 ± 0.2 months. Since the z-scores already take into consideration the child's age, eventual differences in mean age and standard deviation should not affect comparisons of nutritional status. The number of children studied at around age 1 year in each cohort is presented in Table 1, and broken down according to sex, family income, and birth weight.

Table 2 presents the prevalence of length-for-age, weight-for-age, and weight-for-length deficits, as well as overweight, in each of the cohorts and according to the NCHS⁸ and WHO⁹ reference curves. For all indicators of nutritional deficit, there was a reduction in prevalence between 1982 and 2004, but results for 1993 were inconsistent with the overall trend when the NCHS⁸ curve was used. In the case of length-for-age, the deficit increased between 1982 and 1993, with a

subsequent reduction between 1993 and 2004. Analysis of weight-for-age based on the WHO⁹ curve showed a decreasing trend, followed by stabilization. Weight-for-length deficits were less prevalent than the other two types, and show a consistent reduction across the three cohorts according to both references. Figure 1 offers a clearer perspective of trends according to the WHO⁹ reference: an important reduction in deficits in the three indicators is observed between 1982 and 1993, followed by a period of stability between 1993 and 2004.

Considering the entire period (1982 to 2004), reductions were greater according to WHO⁹ than to NCHS⁸ for all three indicators, which is reflected in the existence of a statistically significant difference between cohorts only when the WHO reference is used. Table 2 also shows an increase in overweight between 1982 and 1993, with subsequent decline in 2004. This reduction took place primarily in children from families earning between 3.1 and 6 times the minimum wage (Figure 2).

All analyses presented from this point on are based on the WHO⁹ reference curves alone. Table 3 shows that boys tended to show greater prevalence of deficits than girls, although only the difference in length-for-age was significant, and only for 1982 and 2004. Decreasing trends in the period were similar for both sexes. Boys also showed greater prevalence of overweight when

Table 1

Number of children examined at age 12 months in the 1982, 1993, and 2004 cohorts according to sex, family income, and birth weight. Pelotas, Southern Brazil.

Variables	1982		1993		2004	
	n	%	n	%	n	%
Sex						
Male	714	49.0	671	49.2	2,024	51.8
Female	742	51.0	693	50.8	1,883	48.2
Family income (as a multiple of the minimum wage)						
≤ 1.0	252	17.4	244	18.2	792	20.3
1.1-3.0	740	51.0	591	44.1	1,803	46.3
3.1-6.0	282	19.4	298	22.2	886	22.7
6.1-10.0	91	6.3	114	8.5	216	5.5
> 10.0	85	5.9	93	6.9	199	5.1
Birth weight (g)						
< 2,500	93	6.5	132	9.7	345	8.8
2,500-2,999	378	26.2	325	23.9	973	24.9
≥ 3,000	970	67.3	904	66.4	2,588	66.3
Total	1,457	-	1,364	-	3,907	-

Table 2

Prevalence of malnutrition and overweight at age 12 months in the 1982, 1993, and 2004 cohorts, according to National Center for Health Statistics (NCHS) ⁸ and World Health Organization (WHO) ⁹ reference levels. Pelotas, Southern Brazil.

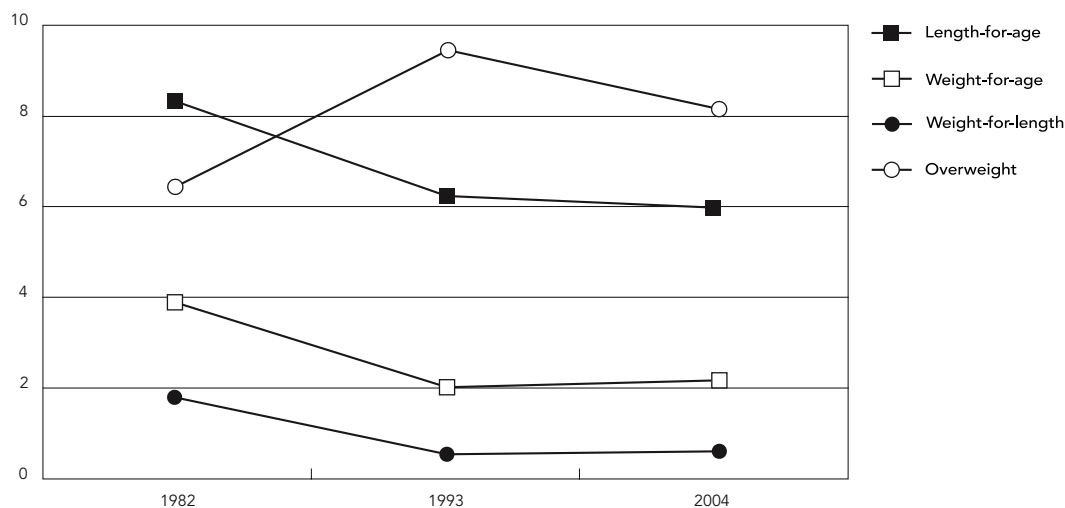
Classification		1982	1993	2004	Variation (1982-2004) *	p **
		%	%	%	%	
Length-for-age	NCHS	5.3	6.1	5.0	-5.7	0.298
	WHO	8.3	6.3	6.0	-27.7	0.004
Weight-for-age	NCHS	5.4	3.7	4.8	-11.1	0.079
	WHO	3.9	2.1	2.2	-43.6	0.002
Weight-for-length	NCHS	1.3	0.9	0.7	-46.2	0.096
	WHO	1.8	0.6	0.6	-66.7	< 0.001
Overweight	NCHS	4.0	6.7	5.5	37.5	0.006
	WHO	6.5	9.4	8.2	26.2	0.100

* $(\%2004 - \%1982)/\%1982 \times 100$;

** χ^2 test for heterogeneity of proportions.

Figure 1

Prevalence of length-for-age, weight-for-age, and weight-for-length deficits ($z < 2$ standard deviations) and overweight (z weight-for-age > 2 standard deviations) at age 12 months in children from the 1982, 1993, and 2004 cohorts, according to World Health Organization ⁹ reference levels. Pelotas, Southern Brazil.

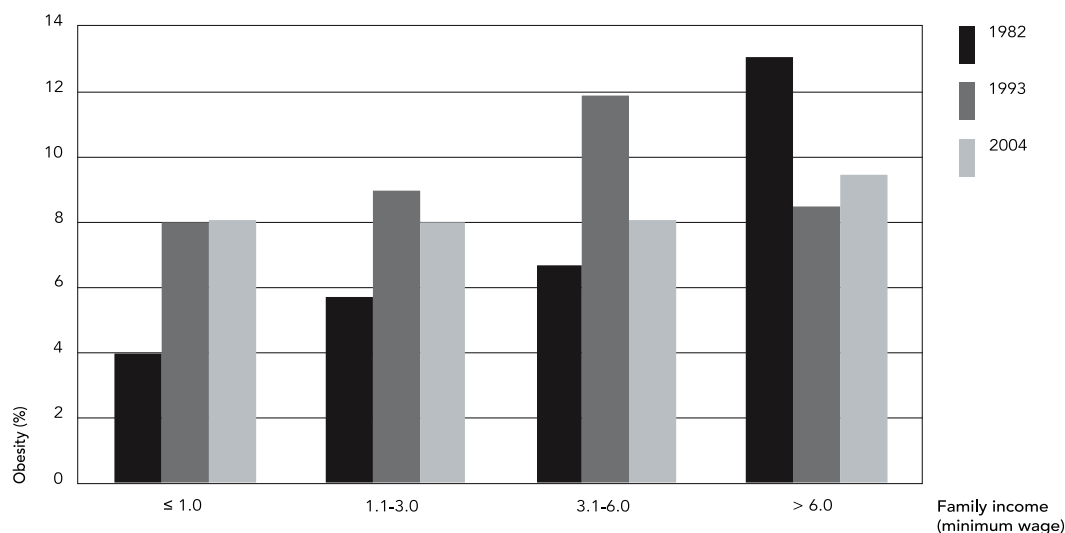


compared to girls in all three studies. These differences, however, decreased with time so that prevalence among girls was virtually equal to that of boys in 2004. Between 1982 and 2004, the prevalence of obesity among boys remained stable, whereas its prevalence among girls increased 66% ($p = 0.01$).

Table 4 presents trends in anthropometric indicators according to family income. There was an important reduction – almost 50% – in the prevalence of nutritional deficits within the two poorest strata. Prevalence in the remaining groups, which was already very low to begin with, remained stable in 2004. There were some

Figure 2

Prevalence of overweight (z-scores of weight-for-length > 2) according to family income * among children from the 1982, 1993, and 2004 cohorts, according to World Health Organization 9 reference levels. Pelotas, Southern Brazil.



* Even though family income had a significant effect for the 1982 cohort study ($p < 0.001$), its effect was not significant for the 1993 ($p = 0.284$) and the 2004 cohort study ($p = 0.499$).

Table 3

Prevalence of malnutrition and overweight at age 12 months by sex in children from the 1982, 1993, and 2004 cohorts according to World Health Organization 9 reference levels. Pelotas, Southern Brazil.

Sex	Length-for-age (%)				Weight-for-age (%)				Weight-for-length (%)				Overweight (%)			
	1982	1993	2004	p *	1982	1993	2004	p *	1982	1993	2004	p *	1982	1993	2004	p *
Male	10.4	7.4	6.8	0.003	4.1	2.9	2.5	0.04	2.1	0.5	0.7	0.006	8.3	10.7	8.6	0.9
Female	6.2	5.3	5.1	0.3	3.6	1.4	1.9	0.02	1.5	0.7	0.5	0.02	4.7	8.1	7.8	0.01
p **	0.004	0.100	0.020	-	0.700	0.070	0.200	-	0.400	0.800	0.400	-	0.005	0.100	0.400	-

* χ^2 for linear trend 1982-2004;

** χ^2 for difference between sexes.

fluctuations between 1993 and 2004 for groups with income of more than three times the minimum wage, but these were not statistically significant. In particular, the 2.5% prevalence of weight-for-height deficit in the group earning more than ten times the minimum wage in 2004 is worthy of note, given that no cases of deficit were recorded in this income group in either 1982 or 1993.

In spite of the reduction in deficits among lower income groups, there was still a strong

association between family income and nutritional deficit in 2004. For 2004, we calculated the mean deficits in the three highest income groups, so as to obtain a more stable estimate of actual prevalence rates. The values obtained were 0.9%, 3.5%, and 0.2% for weight-for-age, length-for-age, and weight-for-length, respectively. Comparing the poorest group with the mean of the three wealthiest groups, prevalence ratios for these indicators were 5, 2, and 10, respectively.

Table 4

Prevalence of malnutrition and overweight at age 12 months by family income in children from the 1982, 1993, and 2004 cohorts according to World Health Organization ⁹ reference levels. Pelotas, Southern Brazil.

Family income (multiple of the minimum wage)	Length-for-age (%)				Weight-for-age (%)				Weight-for-length (%)				Overweight (%)			
	1982	1993	2004	p *	1982	1993	2004	p *	1982	1993	2004	p *	1982	1993	2004	p *
≤ 1.0	16.9	10.9	8.0	< 0.001	9.6	5.600	4.4	0.003	4.0	2.1	1.5	0.020	4.0	8.0	8.1	0.050
1.1-3.0	7.9	7.4	6.8	0.300	3.9	2.000	2.2	0.030	1.9	0.3	0.6	0.005	5.7	9.0	8.0	0.100
3.1-6.0	3.9	2.2	3.6	0.900	0.7	0.800	0.5	0.60	0.7	0.2	0.1	0.090	6.7	11.9	8.1	0.900
6.1-10.0	5.5	2.9	3.3	0.400	0.0	0.300	0.9	-	0.0	0.0	0.0	-	9.9	11.0	6.5	0.200
> 10.0	2.4	4.8	3.1	0.900	0.0	0.000	2.5	-	0.0	0.0	0.5	-	16.5	5.4	12.7	0.700
p **	< 0.001	< 0.001	< 0.001	-	< 0.001	< 0.001	< 0.001	-	0.020	0.020	0.005	-	0.001	0.300	0.200	-

* χ^2 for linear trend 1982-2004;

** χ^2 for difference between income strata.

Table 5

Prevalence of malnutrition and overweight at age 12 months according to birth weight in children from the 1982, 1993, and 2004 cohorts according to World Health Organization ⁹ reference levels. Pelotas, Southern Brazil.

Birthweight (g)	Length-for-age (%)				Weight-for-age (%)				Weight-for-length (%)				Overweight (%)			
	1982	1993	2004	p *	1982	1993	2004	p *	1982	1993	2004	p *	1982	1993	2004	p *
< 2,500	36.7	20.5	21.2	0.006	22.5	11.0	11.1	0.009	13.3	3.0	2.9	< 0.001	0.0	3.5	1.8	0.600
2,500-2,999	14.4	10.3	9.5	0.020	6.7	2.4	3.6	0.040	2.1	0.4	1.2	0.500	5.0	4.3	4.5	0.800
≥ 3,000	3.5	2.7	2.6	0.200	1.1	0.6	0.5	0.070	0.6	0.3	0.1	0.010	7.6	12.1	10.5	0.040
p **	< 0.001	< 0.001	< 0.001	-	< 0.001	< 0.001	< 0.001	-	< 0.001	< 0.001	< 0.001	-	0.006	< 0.001	< 0.001	-

* χ^2 for linear trend 1982-2004;

** χ^2 for difference between birth weight categories.

Table 5 shows evidence of a strong inverse association between birth weight and prevalence of anthropometrical deficits. Between 1982 and 1993, reductions in anthropometric deficits were seen across almost all birth weight categories. This reduction was particularly important among low birth weight babies. Very small differences were detected between 1993 and 2004, however, suggesting that these indicators remained stable.

A comparison of prevalence of overweight by birth weight category in 1993 and 2004 showed no increase between these dates, in contrast with what was observed between 1982 and 1993. While there was an apparent reduction among low-birth weight babies, levels for other groups remained relatively stable (Table 5). In all three cohorts, there was a clear increasing trend for prevalence as birth weight increased.

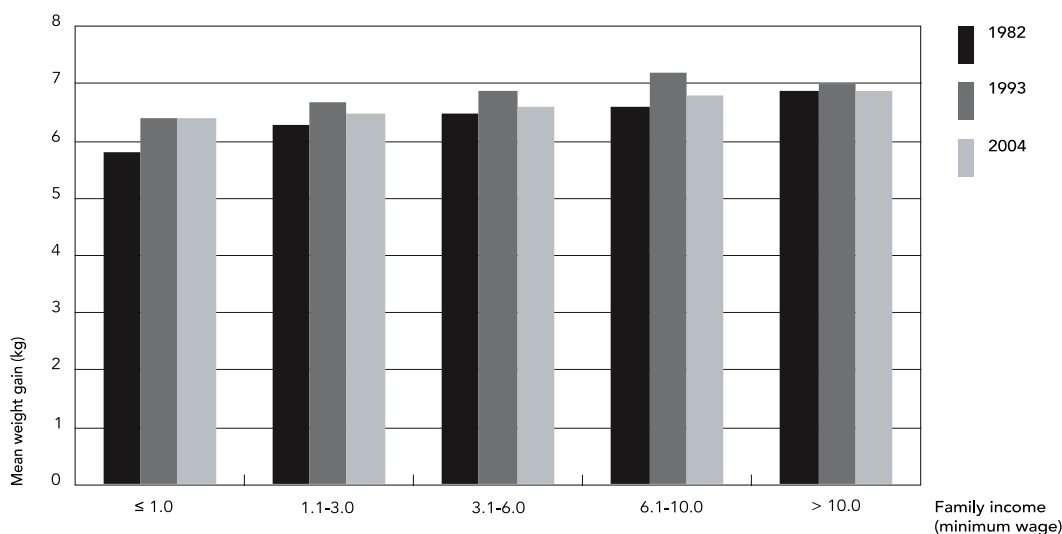
Mean weight gain in the first year of life, after adjustment for differences in age between co-

horts, was 6281.2g in 1982, 6741.3g in 1993, and 6555.9g in 2004 ($p < 0.001$). Weight gain, which had increased between 1982 and 1993 across all categories of family income, leveled out – or even fell slightly – between 1993 and 2004 (Figure 3). The three intermediate income groups, which showed the greatest weight gain between 1982 and 1993, were also those that, between 1993 and 2004, showed the greatest reductions. The difference in weight gain between the highest and lowest income groups fell from 1.1kg in 1982 to 0.5kg in 2004. The effect of the cohort design and of income on mean weight gain was tested using a linear regression model, which was significant with $p < 0.001$.

The evolution of overweight according to family income (Figure 2), shows a clear change of course. In 1982, there was a highly significant ($p < 0.001$) linear trend, which disappeared in 1993 ($p = 0.284$), and was even less evident in 2004 ($p = 0.499$). On the other hand, an anal-

Figure 3

Mean weight gain during first year of life, according to family income *, in children from the 1982, 1993, and 2004 cohorts. Pelotas, Southern Brazil.



* Family income has a significant and different effect on the three cohorts, $p < 0.001$ for interaction between family income and cohort in a linear regression model.

ysis of the evolution of overweight with time, according to income group, shows an increase among the poorest population, accompanied by stability, or even a slight reduction, in other income groups.

Discussion

The availability of data from three population-based birth cohorts provides a unique opportunity for study in Brazil. On all three occasions, these studies included over 99% of children born in the city. Losses to follow-up were much greater in 1982 than in 1993 or 2004, but an analysis of these losses showed that there was little variation in terms of socioeconomic levels⁷. In 1982, for example, we were able to locate 68% of children in the income group earning less than one minimum wage, compared to 79% of children with income higher than 10 times the minimum wage. These percentages were 92% and 96% in 1993 and 93% and 96% in 2004, respectively.

A comparison of the 1982 and 2004 cohorts shows important reductions in the prevalence of the three types of nutritional deficits studied, although prevalence rates in 1993 did not always

occupy an intermediate position between those of the 1982 and 2004 cohorts. The observed reductions provide evidence of the nutritional transition taking place in the region.

Reductions observed between 1982 and 2004 were always more marked when the WHO⁹ curves were used compared with the NCHS⁸ standard, which may be a result of the better quality of the former¹⁰. The WHO⁹ growth standard was constructed based on studies carried out in different countries, and, given its design, has better properties than those of the NCHS⁸ standard. The use of the two standards in the present study is intended to allow for comparison with previous studies, as well as with future studies – which will most likely use the new standard. It should be noted that certain inconsistencies were observed when trends were calculated using each of the two curves. For example, the weight-for-age deficit increased between 1993 and 2004 according to NCHS⁸, but remained stable according to the WHO⁹ standard. Generally speaking, however, time trends that used both standards were similar.

Considering the thresholds adopted (< -2 standard deviation for deficit and $> +2$ standard deviation for overweight), a maximum of 2.5% of a

healthy population would be expected to fall outside this range. According to these criteria, prevalence of the weight-for-age deficit in 2004 is within the expected values for a population comparable to the reference. Such a conclusion is misleading, however, given that the weight-for-age score is the result of a combination of length-for-age and weight-for-length¹¹. Thus, even in 2004, there is still an important deficit in length-for-age (6%), combined with excess weight-for-length – indicated by the low prevalence of deficit (0.6%) and high prevalence of excess (8.2%) shown by this indicator. This combination of deficits in linear growth and overweight leads to the prevalence of weight-for-age appearing to be normal. This is one of the reasons why the use of weight-for-age as a criterion for nutritional evaluation of children in countries undergoing nutritional transition has been subject to criticism¹².

Prevalence of length-for-age retardation was about two times higher than that expected for a healthy population, suggesting deficiencies in linear growth already at 12 months, and indicating that little progress has been made during this 22-year period, particularly over the last 11 years. The length-for-age deficit is typically more marked after the second year of life, and tends to fall again at the age of three¹³.

When compared to other Brazilian studies, however, levels in Pelotas are reasonably low. For example, the prevalence of a height-for-age deficit in under-fives ranged from 13% to 17% in Brazil's Northeast and North Regions in an evaluation carried out in 1996 using the NCHS curves¹⁴. In the country's urban Center-South, prevalence of low stature was 4.6%, slightly lower than the levels found in the 1993 and 2004 cohorts. However, the frequency of this same indicator among under-fives in the city of São Paulo, Brazil, in 1996 was similar to that of the NCHS standard (2.4%)¹⁵.

Returning to the WHO standards, in terms of weight-for-height, even though prevalence of the deficit in 1982 (1.8%) was already below the 2.5% expected for a population that is compared to the reference, there was a further reduction to 0.6% in 2004. This finding is confirmed by a number of other Brazilian surveys, which show that this type of deficit is very rare in these kinds of circumstances, even prior to the nutritional transition. In fact, the detection of a reduction in this already low prevalence, together with the maintenance of, or increase in, rates of length-for-age, support the notion that the determinants of weight gain and linear growth are – at least in part – distinct¹⁶. For instance, greater calorie availability may result in improvements in the weight-for-age ratio, but low stature may

remain due to other limiting factors, such as a lack of micronutrients in the diet.

Weight gain in each of the cohorts was adjusted for differences in age among children, in order to account for the fact that mean age was lower and variance higher in 1982. After adjustment, mean weight gain in 1993 was 431g greater than in 1982. Mean weight gain in 2004 was lower than in 1993, and Figure 3 suggests that this is due mainly to the intermediate income groups. The increased prevalence of breastfeeding seen between 1993 and 2004¹⁷ is likely to have played a role in this phenomenon, since infants who are breastfed have slower growth after the first three months of life than formula-fed infants¹⁰.

In the three cohorts, there was a clear inverse relationship between family income and the prevalence of a nutritional deficit. Similar trends were found in several other Brazilian studies^{18,19,20}. Prevalence of a deficit in the three highest income groups (above three minimum wages) in the three cohorts are already within the expected values for a healthy population, according to the WHO⁹ standard. This is not surprising, given that the new WHO⁹ curves are based on studies carried out in six cities, one of which is Pelotas. The criterion for including children from Pelotas in the curves was family income equal to or greater than six times the minimum wage²¹. Given the low prevalence of a deficit among children from wealthy families, additional progress in this group, even in 1982, was unlikely. For this reason, differences between the poorest and wealthiest families tended to decrease in these 22 years, given the certain degree of improvement among lower income groups. This is in agreement with the “reverse equity hypothesis”²², which predicts that new interventions in health initially benefit the wealthy, thus leading to an increase in inequality, such as that observed in 1982. Such interventions – improvements in sanitation and in the treatment of infections that contribute to malnutrition, for example²³ – would have reached the poor during this 22-year period, while at the same time further improvements in the prevalence of malnutrition among the wealthy were not possible. This would lead – as in fact it did – to a reduction in inequality. Nevertheless, the length-for-age deficit in 2004 among the children of families earning less than or the equivalent of the minimum wage remained at 8%, a level that is unacceptable given the region's level of development and the overall prevalence in Pelotas.

The relationship between overweight and family income, on the other hand, was quite interesting. The prevalence of overweight increased markedly among the poor in the 22 years of the study, but fell among the wealthy. As per

the reverse equity hypothesis, one possibility is that wealthier families are already showing concern for infant obesity, and therefore adopting preventive measures, whereas poorer families are yet to modify their behaviors. A factor that may influence both the reduction of overweight among wealthier children and the stagnation of rates among the population as a whole between 1993 and 2004 is the marked increase in both exclusive and total breastfeeding¹⁷. Breastfed children are known to show lower prevalence rates of obesity²⁴, a fact that is also reflected in the new WHO growth curves¹⁰.

Differences in the prevalence of deficit according to birth weight confirmed the importance of the latter's role as a determinant of malnutrition. However, children born weighing 3,000g or more showed a markedly greater risk of becoming obese at age 12 months than did other babies. In the current epidemiological transition through which Brazil is passing, both nutritional

deficits in low-birth weight children and obesity among children with high birth weight must be faced simultaneously.

This comparison of three birth cohorts revealed paradoxical trends in terms of nutritional status, with an important reduction in indicators of weight deficit, a slight improvement in length deficits, and an increase in the prevalence of overweight. The reasons that explain why progress regarding the prevalence of stunting was only slight need to be further investigated, while the increasing trend towards greater rates of overweight amongst infants must be the subject of interventions by healthcare professionals. The persistence of social determinants of the deficit – although somewhat attenuated by the already very low levels reached by the wealthy – and the increasing problem of obesity among the poor pose important challenges, that will require effective action by healthcare professionals.

Resumo

O estudo teve como objetivo comparar a evolução dos déficits nutricionais e sobrepeso em crianças de um ano, com base em três coortes de nascimentos iniciadas em 1982, 1993 e 2004 em Pelotas, Rio Grande do Sul, Brasil. Amostras das coortes de 1982 e 1993 e todas as crianças da coorte de 2004 foram pesadas e medidas, e suas mães foram entrevistadas. Déficits antropométricos e sobrepeso foram analisados utilizando os padrões de crescimento definidos pelo NCHS e OMS. A comparação dos déficits nutricionais mostrou que, depois de declinar entre 1982 e 1993, a prevalência se estabilizou entre 1993 e 2004. Durante o período inteiro, foi observada uma diminuição em todos os déficits. Inversamente, houve um aumento na obesidade. Baixa estatura para a idade esteve fortemente associada com renda familiar. O grupo de renda familiar abaixo de um salário mínimo foi o único que mostrou uma redução significativa em baixa estatura para idade durante o período estudado. Houve melhorias importantes na redução dos déficits nutricionais na primeira metade do período de estudo, embora ainda persistam desigualdades sociais. Combater a desnutrição ainda é necessário nos 40% mais pobres da população, em paralelo com a luta contra o sobrepeso, que vem sendo observado em todas as classes sociais.

Desnutrição; Sobrepeso; Antropometria; Estudos de Coorte

Contributors

A. J. D. Barros and A. Matijasevich were responsible for the data analysis and for preparing the first draft of the article. C. G. Victora, I. S. Santos, C. L. Araújo and F. C. Barros helped in the interpretation of results and revision of the manuscript and contributed towards the final draft. All authors read and approved the final version.

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