

THE UNIT-LOGISTIC DISTRIBUTION: DIFFERENT METHODS OF ESTIMATION

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Received November 16, 2017 / Accepted October 12, 2018

ABSTRACT. This paper addresses the different methods of estimation of the unknown parameters of a two-parameter unit-logistic distribution from the frequentist point of view. We briefly describe different approaches, namely, maximum likelihood estimators, percentile based estimators, least squares estimators, maximum product of spacings estimators, methods of minimum distances: Cramér-von Mises, Anderson-Darling and four variants of Anderson-Darling. Monte Carlo simulations are performed to compare the performances of the proposed methods of estimation for both small and large samples. The performances of the estimators have been compared in terms of their relative bias, root mean squared error, average absolute difference between the theoretical and empirical estimate of the distribution functions and the maximum absolute difference between the theoretical and empirical distribution functions using simulated samples. Also, for each method of estimation, we consider the interval estimation using the Bootstrap confidence interval and calculate the coverage probability and the average width of the Bootstrap confidence intervals. Finally, two real data sets have been analyzed for illustrative purposes.

Keywords: Unit-Logistic distribution, Monte Carlo simulations, estimation methods, parametric Bootstrap.

1 INTRODUCTION

Tadikamalla & Johnson [30] introduced a new probability distribution with support on $(0, 1)$ and named the distribution as L_B distribution by using transformations of logistic variables. They obtained the distribution as follows:

$$X = g^{-1} \left(\frac{Y - \gamma}{\delta} \right), \quad (1)$$

where Y is a standard Logistic distribution, $g(\cdot)$ is some suitable function and $\gamma \in \mathbb{R}$, $\delta > 0$ are parameters. The choice of $g(\cdot)$ determines the support of the distribution, hence from [30], by taking:

$$g(X) = \log \left(\frac{X}{1 - X} \right), \quad (2)$$

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we can obtain the L_B distribution, hereafter we refer it as unit-Logistic distribution, with probability density function (PDF) given by:

$$f(x | \gamma, \delta) = \frac{\delta e^\gamma x^{\delta-1} (1-x)^{\delta-1}}{[x^\delta e^\gamma + (1-x)^\delta]^2}, \quad 0 < x < 1. \quad (3)$$

In spite of its versatility, this distribution did not receive much attention in the literature. Nevertheless, recently, the basic properties and a regression analysis was studied by [5]. The authors introduced an alternative parametrization, where one parameter is the median. Following this parametrization, they defined the PDF as:

$$f(x | \mu, \beta) = \frac{\beta \mu^\beta x^{\beta-1} (1-\mu)^\beta (1-x)^{\beta-1}}{[(1-\mu)^\beta x^\beta + \mu^\beta (1-x)^\beta]^2}, \quad 0 < x < 1 \quad (4)$$

where $0 < \mu < 1$ is the median of X and $\beta > 0$ is the shape parameter. The corresponding cumulative distribution function and quantile function are written respectively as:

$$F(x | \mu, \beta) = \left[1 + \left(\frac{\mu(1-x)}{x(1-\mu)} \right)^\beta \right]^{-1}, \quad 0 < x < 1 \quad (5)$$

and

$$Q(p | \mu, \beta) = \frac{\mu p^{1/\beta}}{(1-\mu)(1-p)^{1/\beta} + \mu p^{1/\beta}}, \quad 0 < p < 1. \quad (6)$$

Note that when we set $\mu = 0.5$ and $\beta = 1$ in (4), the PDF of the unit-logistic distribution simply becomes the PDF of the standard uniform distribution. As we can see in Figure 1 the unit-Logistic density is uni-modal (or uni-antimodal), increasing, decreasing, or constant, depending on the values of the parameters.

The objective of this paper is to introduce different methods of estimation for the unknown parameters that index the unit-Logistic distribution and to study the behavior of these estimators for different sample sizes and for different parameter values. In particular, we compare the maximum likelihood estimators, maximum product of spacings estimators, estimators based on percentiles, least-squares estimators, weighted least-squares estimators, Cramér-von Mises estimators and Anderson-Darling estimators and four of its variants. Since, it is difficult to compare theoretically the performances of the different estimators, we perform extensive simulations to compare the performances of the different estimation methods based on relative bias, root mean squared error, average absolute difference between the theoretical and empirical estimate of the distribution functions, and maximum absolute difference between the theoretical and empirical distribution functions. Also, for each method of estimation, we consider the interval estimation using the Bootstrap confidence interval [12] and calculate the coverage probability and the average width of the confidence interval.

The uniqueness of this study comes from the fact that thus far, no attempt has been made to compare all these estimators for the two-parameter unit-logistic distribution. Comprehensive comparisons of estimation methods for other distributions have been performed in the literature: see [13]

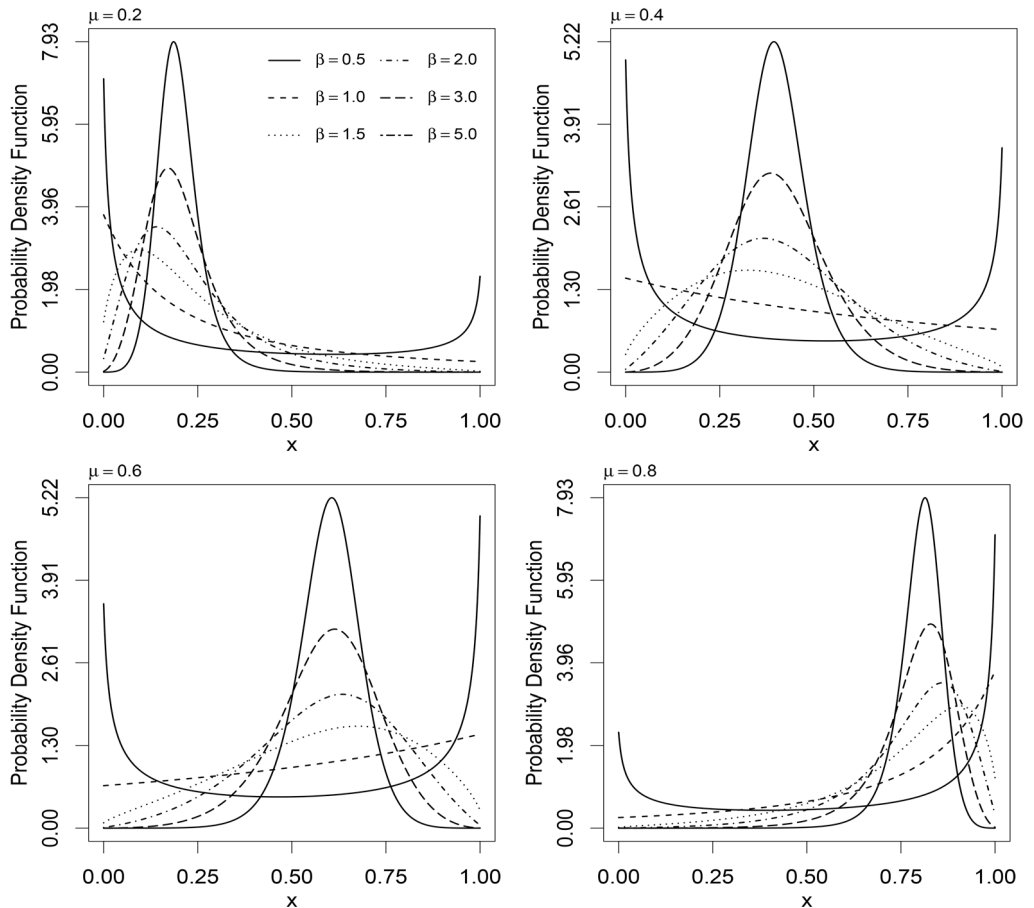


Figure 1 – Behavior of the probability density function of unit-Logistic distribution for some values of μ and β .

for generalized Exponential distribution, [17] for generalized Rayleigh distributions, [31] for Weibull distribution, [22] for weighted Lindley distribution, [10] for Marshall-Olkin extended Lindley distribution, [7] for weighted Exponential distribution, [21] for Marshall-Olkin extended Exponential distribution, [9] for the Kumaraswamy distribution, [8] for the Exponentiated Chen distribution, [27] for the Poisson-exponential distribution, [24] for the Alpha logarithmic transformed Weibull distribution and [23] for the power inverse Lindley distribution.

The final motivation of the paper is to show how different frequentist estimators of this distribution perform for different sample sizes and different parameter values and to develop a guideline for choosing the best estimation method for the unit-logistic distribution, which we think would be of interest to applied statisticians.

The paper is organized as follows. In Section 2 we discuss the eleven estimation methods considered in this paper. The performance of the proposed estimation procedures is studied through a Monte Carlo simulation and is presented in Section 3. In section 4, the methodology developed in this manuscript and the usefulness of the unit-Logistic distribution is illustrated by using two real data examples. Some concluding remarks are presented in Section 5.

2 METHODS OF ESTIMATION

In this section, we describe seven methods and four variants of AD test for estimating the parameters, μ and β , associated to the unit-Logistic distribution. For all methods, it is assumed that $\mathbf{x} = (x_1, \dots, x_n)$ is a random sample of size n from the unit-Logistic distribution with PDF given by (4) and unknown parameters μ and β . Also, let $x_{(1)} < \dots < x_{(n)}$ be the corresponding order sample statistics.

2.1 Method of Maximum Likelihood

Undoubtedly the method of maximum likelihood is the most popular method in statistical inference, mainly because of its many appealing properties. For instance, the maximum likelihood estimates are asymptotically unbiased, efficient, consistent, invariant under parameter transformation and asymptotically normally distributed (see, e.g., [18, 25, 28]).

The log-likelihood function of unit-Logistic distribution based on the random sample $\mathbf{x} = (x_1, \dots, x_n)$ can be written as:

$$\begin{aligned} \ell(\mu, \beta \mid \mathbf{x}) &= n \log \beta + n \beta \log (1 - \mu) + n \beta \log \mu \\ &+ (\beta - 1) \sum_{i=1}^n \log x_i + (\beta - 1) \sum_{i=1}^n \log (1 - x_i) \\ &- 2 \sum_{i=1}^n \log [(1 - \mu)^\beta x_i^\beta + \mu^\beta (1 - x_i)^\beta]. \end{aligned} \tag{7}$$

The maximum likelihood estimates $\hat{\mu}_{MLE}$ and $\hat{\beta}_{MLE}$, of the parameters μ and β , respectively, can be obtained by maximizing (7), or equivalently solving the following normal equations:

$$\frac{\partial \ell}{\partial \mu} = \frac{n (\beta - \mu \beta)}{\mu (1 - \mu)} - 2 \beta \sum_{i=1}^n \frac{(1 - \mu)^{\beta-1} x_i^\beta - \mu^{\beta-1} (1 - x_i)^\beta}{(1 - \mu)^\beta x_i^\beta + \mu^\beta (1 - x_i)^\beta} = 0, \tag{8}$$

$$\frac{\partial \ell}{\partial \beta} = \frac{n}{\beta} + n \log(1 - \mu) + n \log \mu + \sum_{i=1}^n \log x_i + \sum_{i=1}^n \log(1 - x_i) \tag{9}$$

$$-2 \sum_{i=1}^n \frac{x_i^\beta [\log (1 - \mu) + \log x_i] (1 - \mu) \beta + 2 \mu^\beta (1 - x_i)^\beta [\log (1 - x_i) + \log \mu]}{(1 - \mu)^\beta x_i^\beta + \mu^\beta (1 - x_i)^\beta} = 0.$$

Confidence intervals can be obtained by using the large sample distribution of the MLEs, which is normally distributed with the covariance matrix given by the inverse of the Fisher information since regularity conditions are satisfied.

2.2 Method of Maximum Product of Spacings

The maximum product of spacings (MPS) method was introduced by [2, 3] as an alternative to MLE for estimating parameters of continuous univariate distributions. Ranney [26] independently derived the same method as an approximation to the Kullback–Leibler measure of information.

The uniform spacings of a random sample from unit-Logistic distribution is defined as:

$$D_i(\mu, \beta) = F(x_{(i)} | \mu, \beta) - F(x_{(i-1)} | \mu, \beta), \quad i = 1, \dots, n,$$

$$F(x_{(0)} | \mu, \beta) = 0 \quad \text{and} \quad F(x_{(n+1)} | \mu, \beta) = 0.$$

Clearly, $D_0(\mu, \beta) + D_1(\mu, \beta) + \dots + D_{n+1}(\mu, \beta) = 1$.

From [2, 3], the MPSEs, $\hat{\mu}_{MPS}$ and $\hat{\beta}_{MPS}$, are the values of μ and β , which maximize the geometric mean of the spacings:

$$G(\mu, \beta | \mathbf{x}) = \left[\prod_{i=1}^{n+1} D_i(\mu, \beta) \right]^{\frac{1}{n+1}} \tag{10}$$

or, equivalently, by maximize its logarithm:

$$H(\mu, \beta | \mathbf{x}) = \frac{1}{n+1} \sum_{i=1}^{n+1} \log D_i. \tag{11}$$

The estimators $\hat{\mu}_{MPS}$ and $\hat{\beta}_{MPS}$ of the parameters α and β can also be obtained by solving the nonlinear equations:

$$\frac{\partial}{\partial \mu} H(\mu, \beta) = \frac{1}{n+1} \sum_{i=1}^{n+1} \frac{1}{D_i(\mu, \beta)} [\Delta_1(x_{(i)} | \mu, \beta) - \Delta_1(x_{(i-1)} | \mu, \beta)] = 0,$$

$$\frac{\partial}{\partial \beta} H(\mu, \beta) = \frac{1}{n+1} \sum_{i=1}^{n+1} \frac{1}{D_i(\mu, \beta)} [\Delta_2(x_{(i)} | \mu, \beta) - \Delta_2(x_{(i-1)} | \mu, \beta)] = 0,$$

where

$$\Delta_1(x_{(i)} | \mu, \beta) = \frac{\mu^{\beta-1} \beta x_{(i)}^{-\beta} \left(\frac{\mu-1}{x_{(i)}-1}\right)^\beta}{(\mu-1) \left[x_{(i)} \mu^{-\beta} \left(\frac{\mu-1}{x_{(i)}-1}\right)^\beta + 1 \right]^2} \tag{12}$$

and

$$\Delta_2(x_{i:n} | \mu, \beta) = \frac{[\log \mu + \log(1 - x_{(i)}) - \log x_{(i)} - \log(1 - \mu)] x_{(i)}^{-\beta} \mu^\beta \left(\frac{\mu - 1}{x_{(i)} - 1}\right)^\beta}{(\mu - 1) \left[x_{(i)} \mu^{-\beta} \left(\frac{\mu - 1}{x_{(i)} - 1}\right)^\beta + 1 \right]^2}. \tag{13}$$

It is noteworthy that the MPSE is as efficient as ML estimation and consistent under more general conditions than the ML estimators [3].

2.3 Method of Percentiles

If the data come from a distribution function which has a closed form, then we can estimate the unknown parameters by fitting straight line to the theoretical points obtained from the distribution function and the sample percentile points. This method was developed by [15, 16] to estimate the parameters of the Weibull distribution.

Since the unit-Logistic distribution has an explicit cumulative distribution function, (5), it is feasible to use the same concept to derive estimators for μ and β . If p_i denotes some estimate of $F(x_{(i)} | \mu, \beta)$, then the percentiles estimators, $\widehat{\mu}_{PCE}$ and $\widehat{\beta}_{PCE}$, can be obtained by minimizing, with respect to μ and β , the nonlinear function:

$$P(\mu, \beta | \mathbf{x}) = \sum_{i=1}^n \left[x_{(i)} - \frac{\mu p_i^{1/\beta}}{(1 - \mu)(1 - p_i)^{1/\beta} + \mu p_i^{1/\beta}} \right]^2, \tag{14}$$

where $p_i = \frac{i}{n + 1}$ is an unbiased estimator of $F(x_{(i)} | \mu, \beta)$. It is to be mentioned here that there are several possible choices for p_i , interested readers may refer to [20].

2.4 Methods of Least Squares

The least square methods were originally proposed by [29] to estimate the parameters of the Beta distributions. Suppose that $F(X_{(i)})$ denotes the distribution function of the order statistics from the random sample $\mathbf{x} = (x_1, \dots, x_n)$. An important result from the probability shows that $F(X_{(i)}) \sim \text{Beta}(i, n - i + 1)$. Therefore, we have:

$$E[F(X_{(i)})] = \frac{i}{n + 1} \quad \text{and} \quad V[F(X_{(i)})] = \frac{i(n - i + 1)}{(n + 1)^2(n + 2)} \tag{15}$$

for further details see [14]. Using the expectations and variances, we obtain two variants of the least squares methods.

2.4.1 Ordinary Least Squares

In case of unit-Logistic distribution, the ordinary least square estimates $\widehat{\mu}_{OLS}$ and $\widehat{\beta}_{OLS}$ of the parameters μ and β can be obtained by minimizing the function:

$$S(\mu, \beta | \mathbf{x}) = \sum_{i=1}^n \left[F(x_{(i)} | \mu, \beta) - \frac{i}{n+1} \right]^2 \tag{16}$$

with respect to μ and β . Alternatively, these estimates can also be obtained by solving the following nonlinear equations:

$$\begin{aligned} \sum_{i=1}^n \left[F(x_{(i)} | \mu, \beta) - \frac{i}{n+1} \right] \Delta_1(x_{(i)} | \mu, \beta) &= 0, \\ \sum_{i=1}^n \left[F(x_{(i)} | \mu, \beta) - \frac{i}{n+1} \right] \Delta_2(x_{(i)} | \mu, \beta) &= 0 \end{aligned}$$

where $\Delta_1(\cdot | \mu, \beta)$ and $\Delta_2(\cdot | \mu, \beta)$ are given by Equations (12) and 13, respectively.

2.4.2 Weighted Least Squares

For the unit-Logistic distribution, the weighted least square estimates of μ and β , say $\widehat{\mu}_{WLS}$ and $\widehat{\beta}_{WLS}$, respectively are obtained by minimizing the function:

$$W(\mu, \beta | \mathbf{x}) = \sum_{i=1}^n \frac{(n+1)^2 (n+2)}{i(n-i+1)} \left[F(x_{(i)} | \mu, \beta) - \frac{i}{n+1} \right]^2 \tag{17}$$

with respect to μ and β . Equivalently, these estimates are the solution of the following nonlinear equations:

$$\begin{aligned} \sum_{i=1}^n \frac{(n+1)^2 (n+2)}{i(n-i+1)} \left[F(x_{(i)} | \mu, \beta) - \frac{i}{n+1} \right] \Delta_1(x_{(i)} | \mu, \beta) &= 0, \\ \sum_{i=1}^n \frac{(n+1)^2 (n+2)}{i(n-i+1)} \left[F(x_{(i)} | \mu, \beta) - \frac{i}{n+1} \right] \Delta_2(x_{(i)} | \mu, \beta) &= 0 \end{aligned}$$

where $\Delta_1(\cdot | \mu, \beta)$ and $\Delta_2(\cdot | \mu, \beta)$ are defined in Equations (12) and 13, respectively.

2.5 Methods of Minimum Distances

Here, we will discuss some methods based on the test statistics of Cramér-von Mises, Anderson-Darling and four variants of the Anderson-Darling test, whose acronyms are ADR, AD2R, AD2L and AD2. Mainly, these methods determine the values of parameters that minimize the distance between the theoretical and empirical cumulative distribution functions (see for further details e.g., [6, 19]). The expressions for each method are presented in Table 1.

For illustrative purposes, we have presented only the expressions used for the estimation of the parameters for the Cramér-von- Mises and Anderson-Darling methods.

Table 1 – Expressions for the methods based on the minimum distances.

Acronyms	Expressions
CvM	$W_n^2 = \frac{1}{12n} + \sum_{i=1}^n \left(x_{(i)} - \frac{2i-1}{2n}\right)^2$
AD	$A_n^2 = -n - \frac{1}{n} \sum_{i=1}^n (2i - 1) [\log x_{(i)} + \log(1 - x_{(n+1-i)})]$
ADR	$R_n^2 = \frac{n}{2} - 2 \sum_{i=1}^n x_{(i)} - \frac{1}{n} \sum_{i=1}^n (2i - 1) \log(1 - x_{(n+1-i)})$
AD2R	$r_n^2 = 2 \sum_{i=1}^n \log(1 - x_{(i)}) + \frac{1}{n} \sum_{i=1}^n \frac{2i-1}{1-x_{(n+1-i)}}$
AD2L	$l_n^2 = 2 \sum_{i=1}^n \log x_{(i)} + \frac{1}{n} \sum_{i=1}^n \frac{2i-1}{x_{(i)}}$
AD2	$a_n^2 = 2 \sum_{i=1}^n [\log x_{(i)} + \log(1 - x_{(i)})] + \frac{1}{n} \sum_{i=1}^n \left(\frac{2i-1}{x_{(i)}} + \frac{2i-1}{1-x_{(n+1-i)}}\right)$

2.5.1 Method of Cramér-von Mises

In regard to unit-Logistic distribution, the Cramér-von- Mises estimates $\widehat{\mu}_{CvM}$ and $\widehat{\beta}_{CvM}$ are obtained by minimizing with respect to μ and β the function:

$$C(\mu, \beta | \mathbf{x}) = \frac{1}{12n} + \sum_{i=1}^n \left(F(x_{(i)} | \mu, \beta) - \frac{2i - 1}{2n}\right)^2. \tag{18}$$

The estimates can also be obtained by solving the following nonlinear equations:

$$\begin{aligned} \sum_{i=1}^n \left(F(x_{(i)} | \mu, \beta) - \frac{2i - 1}{2n}\right) \Delta_1(x_{(i)} | \mu, \beta) &= 0, \\ \sum_{i=1}^n \left(F(x_{(i)} | \mu, \beta) - \frac{2i - 1}{2n}\right) \Delta_2(x_{(i)} | \mu, \beta) &= 0 \end{aligned}$$

where $\Delta_1(\cdot | \mu, \beta)$ and $\Delta_2(\cdot | \mu, \beta)$ are specified in Equations (12) and 13, respectively.

2.5.2 Method of Anderson-Darling

[1] developed a test, as an alternative to statistical tests for detecting sample distributions departure from normality. Using these test statistics, we can obtain the Anderson-Darling estimates, $\widehat{\mu}_{ADE}$ and $\widehat{\beta}_{ADE}$, by minimizing the function

$$A(\mu, \beta | \mathbf{x}) = -n - \frac{1}{n} \sum_{i=1}^n (2i - 1) \{\log F(x_{(i)} | \mu, \beta) + \log \overline{F}(x_{(n+1-i)} | \mu, \beta)\}. \tag{19}$$

with respect to μ and β . Equivalently, these estimates are the solution of the following nonlinear equations:

$$\sum_{i=1}^n (2i - 1) \left[\frac{\Delta_1(x_{(i)} | \mu, \beta)}{F(x_{(i)} | \mu, \beta)} - \frac{\Delta_1(x_{(n+1-i)} | \mu, \beta)}{\bar{F}(x_{(n+1-i)} | \mu, \beta)} \right] = 0,$$

$$\sum_{i=1}^n (2i - 1) \left[\frac{\Delta_2(x_{(i)} | \mu, \beta)}{F(x_{(i)} | \mu, \beta)} - \frac{\Delta_2(x_{(n+1-i)} | \mu, \beta)}{\bar{F}(x_{(n+1-i)} | \mu, \beta)} \right] = 0$$

where $\Delta_1(\cdot | \alpha, \beta)$ and $\Delta_2(\cdot | \alpha, \beta)$ are given by (12) and (13), respectively.

3 SIMULATION RESULTS

In this section we conduct a Monte Carlo simulation study to compare the performance of the frequentist estimators discussed in the previous sections. The methods are compared for sample sizes $n = \{20, 50, 100, 200\}$. We generate $M = 5.000$ pseudo-random samples from unit-Logistic distribution using the inverse transform method with parameters $\mu = \{0.2, 0.4, 0.6, 0.8\}$ and $\beta = \{0.5, 1.5, 2.0\}$.

All simulations are done in **Ox** version 7.10, (see [11]), using the **MaxBFGS** subroutine for numerical optimizations. For each estimate, we calculate the relative bias, root mean-squared error (RMSE), the average absolute difference between the theoretical and empirical estimate of the distribution functions (D_{abs}), and the maximum absolute difference between the theoretical and empirical distribution functions (D_{max}). These measures are obtained using the following formulae:

$$\text{Bias}(\hat{\Theta}) = \frac{1}{M} \sum_{i=1}^M \left(\frac{\hat{\Theta}_i - \Theta}{\Theta} \right), \tag{20}$$

$$\text{RMSE}(\hat{\Theta}) = \sqrt{\frac{1}{M} \sum_{i=1}^M (\hat{\Theta}_i - \Theta)^2}, \tag{21}$$

$$D_{abs} = \frac{1}{M \times n} \sum_{i=1}^M \sum_{j=1}^n |F(y_{ij} | \Theta) - F(y_{ij} | \hat{\Theta})|, \tag{22}$$

$$D_{max} = \frac{1}{M} \sum_{i=1}^M \max_j |F(y_{ij} | \Theta) - F(y_{ij} | \hat{\Theta})|. \tag{23}$$

where $\Theta = (\mu, \beta)$. Due to space constraint, we report the results only for $\mu = (0.2, 0.8)$ and $\beta = (0.5, 2)$. The results for other combinations are summarized by their ranks in Tables 6 and 11, however this can be obtained from the corresponding author on request.

In Tables 2–5 we report the empirical values of (20)–(23). A superscript indicates the rank of each of the estimators among all the estimators for that metric. For example, Table 2 presents

the bias of the MLE ($\hat{\beta}$) in the first row as 0.070⁹ for $n = 20$. This indicates that the bias of $\hat{\beta}$ obtained using the method of maximum likelihood ranks 9th among all other estimators.

The following observations can be drawn from Tables 2-5.

1. All the estimators reveal the property of consistency, i.e., the RMSE decreases when the sample size increases.
2. The bias of $\hat{\beta}$ decreases when n increases for all estimation methods.
3. The bias of $\hat{\mu}$ decreases when n increases for all estimation methods.
4. The bias of $\hat{\mu}$ generally decreases when β increases for any given β and n for all estimation methods.
5. The bias of $\hat{\beta}$ generally decreases when μ increases for any given μ and n for all estimation methods.
6. \hat{D}_{abs} is smaller than \hat{D}_{max} for all estimation techniques. Again, these statistics become smaller when n increases.

The overall ranks of the estimation methods are presented in Table 6. For the parameter combinations considered in our study, Anderson-Darling estimator (AD) turns out to be the best (overall score of 159) in the overall ranking closely followed by the method of weighted least square (WLS) (overall score of 178).

In the previous tables, we have obtained the point estimates of each method of estimation. However, it is also important to know the behaviour of interval estimation for each method of estimations. Therefore, we computed the parametric Bootstrap confidence interval [12] and evaluate their coverage probability and average length of the simulated confidence intervals. The results are presented in Tables 7–10.

From the results reported in Tables 7–10, it is observed that as sample sizes increases, the coverage probability increases for both the parameters as well as for the estimation methods, while the average width of the confidence intervals decreases as the sample sizes increases for both the parameters and estimation methods.

The overall positions of the interval estimates are presented in Table 11. It is observed that WLS is the best method for interval estimation based on parametric Bootstrap confidence intervals. The next best method is the AD, followed by MLE.

Thus, based on our study we may conclude that AD and WLS are the best methods for estimating the parameters of unit-Logistic distribution for both point and interval estimation. Therefore, we suggest to use AD and WLS methods of estimation for practical purposes.

Table 2 – Simulations results for $\mu = 0.2$ and $\beta = 0.5$.

n	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
20	Bias($\hat{\beta}$)	0.025 ⁴	-0.038 ⁵	0.071 ¹⁰	0.069 ⁷	0.052 ⁶	0.077 ¹¹	0.070 ⁹	-0.070 ⁸	-0.009 ³	0.000 ¹	0.002 ⁻²
	RMSE($\hat{\beta}$)	0.214 ²	0.216 ³	0.361 ¹⁰	0.366 ¹¹	0.258 ⁷	0.271 ⁸	0.226 ⁵	0.200 ¹	0.235 ⁶	0.298 ⁹	0.223 ⁴
	Bias($\hat{\mu}$)	0.108 ³	0.126 ⁹	0.088 ¹	0.167 ¹¹	0.129 ¹⁰	0.109 ⁷	0.107 ²	0.108 ⁴	0.109 ⁶	0.111 ⁸	0.108 ⁵
	RMSE($\hat{\mu}$)	0.642 ²	0.693 ¹⁰	0.676 ⁸	0.722 ¹¹	0.657 ⁷	0.645 ⁶	0.641 ¹	0.642 ³	0.644 ⁴	0.681 ⁹	0.644 ⁵
	\hat{D}_{abs}	0.098 ⁸	0.097 ²	0.098 ¹¹	0.097 ³	0.098 ¹⁰	0.098 ⁹	0.098 ⁷	0.098 ⁵	0.097 ⁴	0.096 ¹	0.097 ⁶
	\hat{D}_{max}	0.222 ⁷	0.221 ³	0.223 ¹¹	0.222 ⁸	0.222 ¹⁰	0.222 ⁵	0.222 ⁹	0.222 ⁶	0.222 ⁴	0.220 ¹	0.222 ²
Total	26 ³	32 ⁵	51 ¹⁰	51 ¹⁰	50 ⁹	46 ⁸	33 ⁶	27 ⁴	21 ¹	35 ⁷	24 ²	
50	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	Bias($\hat{\beta}$)	0.007 ⁶	-0.046 ¹¹	0.001 ²	-0.001 ¹	0.016 ⁷	0.026 ⁹	0.024 ⁸	-0.044 ¹⁰	-0.007 ⁵	-0.004 ⁴	0.002 ⁻³
	RMSE($\hat{\beta}$)	0.124 ³	0.143 ⁷	0.201 ¹⁰	0.205 ¹¹	0.142 ⁶	0.145 ⁸	0.124 ²	0.122 ¹	0.137 ⁵	0.187 ⁹	0.127 ⁴
	Bias($\hat{\mu}$)	0.046 ⁶	0.059 ¹⁰	0.076 ¹¹	0.046 ³	0.055 ⁹	0.047 ⁸	0.046 ⁴	0.046 ⁵	0.047 ⁷	0.044 ¹	0.046 ²
	RMSE($\hat{\mu}$)	0.398 ²	0.443 ⁹	0.460 ¹¹	0.455 ¹⁰	0.405 ⁷	0.400 ⁶	0.398 ¹	0.398 ³	0.400 ⁵	0.432 ⁸	0.399 ⁴
	\hat{D}_{abs}	0.062 ²	0.062 ⁶	0.062 ¹⁰	0.062 ¹¹	0.062 ³	0.062 ⁷	0.062 ⁴	0.061 ¹	0.062 ⁹	0.062 ⁸	0.062 ⁵
\hat{D}_{max}	0.152 ²	0.152 ¹	0.153 ⁹	0.153 ¹¹	0.152 ³	0.152 ⁷	0.152 ⁴	0.152 ⁵	0.152 ⁶	0.153 ¹⁰	0.152 ⁸	
Total	21 ¹	44 ⁸	53 ¹¹	47 ¹⁰	35 ⁵	46 ⁹	24 ²	24 ²	38 ⁶	40 ⁷	24 ²	
100	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	Bias($\hat{\beta}$)	0.003 ³	-0.039 ¹¹	-0.013 ⁹	-0.012 ⁶	0.008 ³	0.012 ⁷	0.012 ⁸	-0.027 ¹⁰	-0.004 ⁴	-0.003 ²	0.002 ¹
	RMSE($\hat{\beta}$)	0.087 ³	0.109 ⁸	0.148 ¹⁰	0.152 ¹¹	0.098 ⁶	0.099 ⁷	0.086 ¹	0.086 ²	0.096 ⁵	0.132 ⁹	0.088 ⁴
	Bias($\hat{\mu}$)	0.026 ⁴	0.032 ¹⁰	0.067 ¹¹	0.009 ¹	0.030 ⁹	0.026 ⁸	0.026 ⁶	0.026 ⁵	0.026 ⁷	0.025 ²	0.025 ³
	RMSE($\hat{\mu}$)	0.277 ²	0.309 ⁹	0.348 ¹¹	0.327 ¹⁰	0.281 ⁷	0.279 ⁶	0.277 ¹	0.277 ³	0.279 ⁵	0.302 ⁸	0.277 ⁴
	\hat{D}_{abs}	0.044 ³	0.044 ⁶	0.044 ⁴	0.044 ⁸	0.044 ¹⁰	0.044 ¹¹	0.044 ⁷	0.043 ¹	0.044 ⁹	0.044 ²	0.044 ⁵
\hat{D}_{max}	0.111 ¹	0.112 ³	0.112 ⁷	0.112 ⁶	0.113 ¹⁰	0.113 ¹¹	0.112 ⁵	0.112 ²	0.112 ⁹	0.112 ⁴	0.112 ⁸	
Total	16 ¹	47 ⁸	52 ¹¹	42 ⁷	47 ⁸	50 ¹⁰	28 ⁵	23 ²	39 ⁶	27 ⁴	25 ³	
200	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	Bias($\hat{\beta}$)	0.001 ¹	-0.032 ¹¹	-0.016 ⁸	-0.020 ¹⁰	0.003 ³	0.005 ⁶	0.005 ⁷	-0.017 ⁹	-0.003 ⁴	-0.003 ²	0.001 ²
	RMSE($\hat{\beta}$)	0.061 ³	0.083 ⁸	0.113 ¹⁰	0.113 ¹¹	0.068 ⁷	0.067 ⁶	0.059 ¹	0.060 ²	0.066 ⁵	0.092 ⁹	0.061 ⁴
	Bias($\hat{\mu}$)	0.012 ⁵	0.018 ¹⁰	0.054 ¹¹	-0.017 ⁹	0.014 ⁸	0.012 ²	0.012 ⁴	0.012 ⁶	0.012 ³	0.010 ¹	0.012 ⁷
	RMSE($\hat{\mu}$)	0.198 ²	0.224 ⁹	0.263 ¹¹	0.241 ¹⁰	0.199 ⁷	0.198 ⁶	0.198 ¹	0.198 ⁴	0.198 ⁵	0.214 ⁸	0.198 ³
	\hat{D}_{abs}	0.031 ²	0.031 ⁴	0.031 ¹¹	0.031 ³	0.031 ¹	0.031 ¹⁰	0.031 ⁷	0.031 ⁸	0.031 ⁵	0.031 ⁹	0.031 ⁶
\hat{D}_{max}	0.081 ²	0.081 ¹	0.082 ⁸	0.082 ³	0.082 ⁶	0.082 ¹⁰	0.082 ⁵	0.082 ¹¹	0.082 ⁴	0.082 ⁹	0.082 ⁷	
Total	15 ¹	43 ⁹	59 ¹¹	46 ¹⁰	32 ⁵	40 ⁶	25 ²	40 ⁶	26 ³	41 ⁸	29 ⁴	
Overall Total	6 ¹	30 ⁸	43 ¹¹	37 ¹⁰	27 ⁷	33 ⁹	15 ⁴	14 ³	16 ⁵	26 ⁶	11 ²	

Table 3 – Simulations results for $\mu = 0.2$ and $\beta = 2.0$.

n	Qtd	AD	AD2	AD2L	AD2R	ADR	CM	MLE	MPS	OLS	PCE	WLS
20	Bias($\hat{\beta}$)	0.027 ³	-0.044 ⁴	0.049 ⁵	0.060 ⁷	0.056 ⁶	0.070 ¹⁰	0.067 ⁹	-0.071 ¹¹	-0.004 ²	-0.061 ⁸	0.001 ¹
	RMSE($\hat{\beta}$)	0.216 ²	0.218 ³	0.333 ¹⁰	0.340 ¹¹	0.256 ⁸	0.260 ⁹	0.221 ⁵	0.202 ¹	0.236 ⁷	0.234 ⁶	0.220 ⁴
	Bias($\hat{\alpha}$)	0.011 ⁷	0.011 ⁹	0.000 ²	0.021 ¹¹	0.013 ¹⁰	0.007 ³	0.009 ⁴	0.009 ⁵	0.010 ⁶	0.001 ¹	0.011 ⁸
	RMSE($\hat{\alpha}$)	0.160 ⁸	0.168 ¹⁰	0.168 ⁹	0.174 ¹¹	0.158 ⁴	0.158 ³	0.159 ⁷	0.157 ¹	0.158 ²	0.159 ⁶	0.158 ⁵
	\hat{D}_{abs}	0.097 ³	0.097 ⁶	0.097 ⁷	0.097 ⁵	0.098 ⁸	0.097 ²	0.097 ⁴	0.098 ¹¹	0.098 ⁹	0.098 ¹⁰	0.098 ¹⁰
\hat{D}_{max}	0.222 ⁵	0.222 ⁸	0.222 ⁶	0.222 ⁶	0.221 ⁴	0.221 ²	0.221 ³	0.222 ⁷	0.223 ¹¹	0.222 ¹⁰	0.220 ¹	0.222 ⁹
Total	28 ²	40 ⁹	39 ⁸	49 ¹¹	38 ⁷	30 ³	36 ⁴	40 ⁹	36 ⁴	23 ¹	37 ⁶	
50	Bias($\hat{\beta}$)	0.008 ⁵	-0.045 ¹¹	-0.003 ¹	0.003 ²	0.019 ⁶	0.026 ⁷	0.028 ⁸	-0.041 ¹⁰	-0.003 ³	-0.039 ⁹	0.004 ⁴
	RMSE($\hat{\beta}$)	0.127 ²	0.145 ⁷	0.210 ¹⁰	0.210 ¹¹	0.143 ⁶	0.146 ⁸	0.130 ³	0.126 ¹	0.141 ⁵	0.151 ⁹	0.130 ⁴
	Bias($\hat{\alpha}$)	0.004 ⁸	0.005 ⁹	0.008 ¹¹	0.002 ²	0.006 ¹⁰	0.001 ¹	0.003 ⁴	0.003 ⁵	0.004 ⁷	-0.002 ³	0.004 ⁶
	RMSE($\hat{\alpha}$)	0.099 ⁶	0.107 ⁹	0.115 ¹¹	0.113 ¹⁰	0.100 ⁷	0.099 ⁵	0.098 ³	0.098 ²	0.098 ⁴	0.100 ⁸	0.098 ¹
	\hat{D}_{abs}	0.062 ⁸	0.062 ⁶	0.062 ⁷	0.062 ⁵	0.062 ³	0.063 ¹⁰	0.062 ¹	0.062 ⁹	0.063 ¹¹	0.062 ²	0.062 ⁴
\hat{D}_{max}	0.153 ⁶	0.153 ⁵	0.153 ¹⁰	0.153 ⁷	0.152 ³	0.153 ⁹	0.152 ¹	0.153 ⁸	0.154 ¹¹	0.152 ²	0.153 ⁴	
Total	35 ⁴	47 ¹⁰	50 ¹¹	37 ⁷	35 ⁴	40 ⁸	20 ¹	35 ⁴	41 ⁹	33 ³	23 ²	
100	Bias($\hat{\beta}$)	0.005 ³	-0.040 ¹¹	-0.017 ⁸	-0.014 ⁷	0.009 ⁴	0.012 ⁵	0.014 ⁶	-0.026 ¹⁰	-0.002 ¹	-0.024 ⁹	0.004 ²
	RMSE($\hat{\beta}$)	0.089 ³	0.109 ⁹	0.151 ¹¹	0.151 ¹⁰	0.097 ⁶	0.097 ⁶	0.087 ²	0.086 ¹	0.097 ⁵	0.107 ⁸	0.090 ⁴
	Bias($\hat{\alpha}$)	0.003 ⁹	0.002 ⁷	0.011 ¹¹	-0.005 ¹⁰	0.001 ⁶	0.000 ¹	0.001 ⁴	0.001 ³	0.001 ²	-0.001 ⁵	0.002 ⁸
	RMSE($\hat{\alpha}$)	0.069 ²	0.076 ⁹	0.085 ¹¹	0.084 ¹⁰	0.070 ⁷	0.070 ⁴	0.070 ⁵	0.070 ³	0.070 ⁶	0.071 ⁸	0.069 ¹
	\hat{D}_{abs}	0.044 ¹¹	0.044 ²	0.044 ⁵	0.044 ¹⁰	0.044 ¹	0.044 ⁸	0.044 ⁹	0.044 ⁷	0.044 ³	0.044 ⁴	0.044 ⁶
\hat{D}_{max}	0.113 ¹¹	0.112 ³	0.112 ⁴	0.113 ¹⁰	0.112 ¹	0.113 ⁸	0.113 ⁹	0.113 ⁶	0.113 ⁵	0.113 ²	0.113 ⁷	
Total	39 ⁷	41 ⁹	50 ¹⁰	57 ¹¹	26 ²	32 ³	35 ⁶	30 ⁴	19 ¹	39 ⁷	28 ³	
200	Bias($\hat{\beta}$)	0.002 ²	-0.033 ¹¹	-0.020 ¹⁰	-0.017 ⁹	0.005 ⁴	0.007 ⁶	0.007 ⁵	-0.015 ⁸	-0.001 ¹	-0.014 ⁷	0.003 ³
	RMSE($\hat{\beta}$)	0.062 ³	0.082 ⁹	0.113 ¹¹	0.111 ¹⁰	0.069 ⁷	0.069 ⁶	0.060 ¹	0.060 ²	0.068 ⁵	0.076 ⁸	0.062 ⁴
	Bias($\hat{\alpha}$)	0.000 ⁵	0.001 ⁹	0.011 ¹¹	-0.007 ¹⁰	0.001 ⁸	0.000 ⁴	-0.000 ³	0.000 ²	0.000 ¹	-0.001 ⁶	0.001 ⁷
	RMSE($\hat{\alpha}$)	0.049 ²	0.054 ⁹	0.064 ¹¹	0.061 ¹⁰	0.050 ⁷	0.050 ³	0.050 ⁵	0.050 ⁴	0.050 ⁶	0.050 ⁸	0.049 ¹
	\hat{D}_{abs}	0.031 ³	0.031 ⁵	0.031 ⁷	0.031 ⁸	0.031 ⁶	0.031 ²	0.031 ¹⁰	0.031 ⁹	0.031 ¹¹	0.031 ¹	0.031 ⁴
\hat{D}_{max}	0.082 ⁶	0.082 ⁸	0.082 ⁵	0.082 ⁹	0.082 ³	0.081 ¹	0.082 ¹⁰	0.082 ⁷	0.083 ¹¹	0.081 ²	0.082 ⁴	
Total	21 ¹	51 ⁹	55 ¹⁰	56 ¹¹	35 ⁷	22 ²	34 ⁶	32 ⁴	35 ⁷	32 ⁴	23 ³	
Overall Total	14 ¹	37 ⁹	39 ¹⁰	40 ¹¹	20 ⁶	18 ⁵	17 ⁴	21 ⁷	21 ⁷	15 ³	14 ¹	

Table 4 – Simulations results for $\mu = 0.8$ and $\beta = 0.5$.

n	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS	
20	Bias($\hat{\beta}$)	0.028 ⁴	-0.046 ⁵	0.075 ¹⁰	0.072 ⁸	0.054 ⁶	0.079 ¹¹	0.072 ⁷	-0.072 ⁹	-0.004 ²	-0.007 ³	0.003 ¹	
	RMSE($\hat{\beta}$)	0.213 ²	0.214 ³	0.369 ¹¹	0.368 ¹⁰	0.256 ⁷	0.270 ⁸	0.228 ⁵	0.197 ¹	0.240 ⁶	0.302 ⁹	0.228 ⁴	
	Bias($\hat{\mu}$)	-0.029 ²	-0.036 ¹⁰	-0.049 ¹¹	-0.031 ⁶	-0.026 ¹	-0.030 ³	-0.031 ⁵	-0.035 ⁹	-0.032 ⁸	-0.030 ⁴	-0.032 ⁷	
	RMSE($\hat{\mu}$)	0.160 ¹	0.178 ¹⁰	0.186 ¹¹	0.174 ⁸	0.162 ²	0.162 ³	0.163 ⁴	0.166 ⁷	0.167 ⁵	0.174 ⁹	0.163 ⁵	
	\hat{D}_{abs}	0.096 ²	0.097 ⁸	0.097 ⁷	0.097 ³	0.098 ¹⁰	0.097 ⁶	0.097 ⁵	0.097 ⁵	0.096 ¹	0.098 ¹¹	0.098 ⁹	0.097 ⁴
	\hat{D}_{max}	0.220 ²	0.222 ⁸	0.221 ⁵	0.220 ¹	0.223 ¹⁰	0.222 ⁷	0.221 ⁶	0.221 ⁶	0.220 ³	0.223 ¹¹	0.222 ⁹	0.220 ⁴
	Total	13 ¹	44 ⁹	55 ¹¹	36 ⁵	36 ⁵	36 ⁵	38 ⁷	32 ⁴	30 ³	44 ⁹	43 ⁸	25 ²
50	Bias($\hat{\beta}$)	0.011 ⁶	-0.047 ¹¹	0.011 ⁵	0.004 ²	0.021 ⁷	0.030 ⁹	0.029 ⁸	-0.044 ¹⁰	-0.004 ³	0.001 ¹	0.006 ⁴	
	RMSE($\hat{\beta}$)	0.128 ²	0.146 ⁷	0.213 ¹¹	0.208 ¹⁰	0.143 ⁶	0.148 ⁸	0.129 ³	0.122 ¹	0.138 ⁵	0.188 ⁹	0.132 ⁴	
	Bias($\hat{\mu}$)	-0.012 ²	-0.015 ⁶	-0.017 ¹⁰	-0.024 ¹¹	-0.012 ¹	-0.012 ⁴	-0.012 ³	-0.015 ⁸	-0.015 ⁷	-0.016 ⁹	-0.013 ⁵	
	RMSE($\hat{\mu}$)	0.100 ¹	0.111 ⁹	0.114 ¹⁰	0.122 ¹¹	0.102 ⁵	0.101 ⁴	0.100 ³	0.103 ⁷	0.103 ⁶	0.111 ⁸	0.100 ²	
	\hat{D}_{abs}	0.062 ³	0.062 ²	0.062 ¹	0.062 ⁴	0.062 ¹⁰	0.062 ⁹	0.062 ⁵	0.062 ⁵	0.062 ⁸	0.063 ¹¹	0.062 ⁶	
	\hat{D}_{max}	0.153 ⁹	0.152 ¹	0.152 ²	0.153 ⁷	0.153 ⁶	0.153 ⁴	0.152 ³	0.152 ³	0.153 ⁵	0.154 ¹¹	0.153 ⁸	
	Total	23 ¹	36 ⁵	39 ⁷	45 ¹¹	35 ⁴	35 ⁴	38 ⁶	25 ²	39 ⁷	43 ¹⁰	41 ⁹	32 ³
100	Bias($\hat{\beta}$)	0.005 ⁴	-0.041 ¹¹	-0.012 ⁷	-0.011 ⁶	0.008 ⁵	0.014 ⁹	0.013 ⁸	-0.027 ¹⁰	-0.002 ²	0.001 ¹	0.004 ³	
	RMSE($\hat{\beta}$)	0.088 ³	0.110 ⁸	0.152 ¹¹	0.150 ¹⁰	0.099 ⁶	0.099 ⁷	0.085 ¹	0.086 ²	0.095 ⁵	0.132 ⁹	0.089 ⁴	
	Bias($\hat{\mu}$)	-0.006 ³	-0.008 ¹⁰	-0.004 ¹	-0.016 ¹¹	-0.006 ²	-0.007 ⁴	-0.007 ⁶	-0.007 ⁶	-0.008 ⁸	-0.007 ⁷	-0.008 ⁹	
	RMSE($\hat{\mu}$)	0.072 ⁵	0.078 ⁸	0.084 ¹⁰	0.088 ¹¹	0.071 ³	0.072 ⁷	0.072 ⁶	0.070 ¹	0.070 ¹	0.071 ²	0.071 ⁴	
	\hat{D}_{abs}	0.047 ⁷	0.044 ⁴	0.044 ⁹	0.044 ⁶	0.044 ¹	0.044 ¹⁰	0.044 ⁵	0.044 ⁵	0.044 ³	0.044 ¹¹	0.044 ⁸	
	\hat{D}_{max}	0.112 ⁶	0.112 ⁴	0.112 ⁸	0.113 ¹⁰	0.112 ¹	0.113 ¹¹	0.112 ²	0.112 ²	0.112 ⁵	0.112 ³	0.113 ⁹	
	Total	28 ³	45 ⁷	46 ⁸	54 ¹¹	18 ¹	18 ¹	48 ⁹	28 ³	29 ⁵	21 ²	48 ⁹	31 ⁶
200	Bias($\hat{\beta}$)	0.002 ²	-0.033 ¹¹	-0.019 ¹⁰	-0.016 ⁸	0.004 ⁵	0.007 ⁷	0.006 ⁶	-0.017 ⁹	-0.001 ¹	0.002 ³	0.003 ⁴	
	RMSE($\hat{\beta}$)	0.061 ³	0.083 ⁸	0.114 ¹¹	0.113 ¹⁰	0.068 ⁶	0.068 ⁷	0.059 ¹	0.060 ²	0.068 ⁵	0.094 ⁹	0.062 ⁴	
	Bias($\hat{\mu}$)	-0.004 ⁴	-0.005 ¹⁰	0.003 ¹	-0.014 ¹¹	-0.004 ²	-0.004 ⁷	-0.004 ⁵	-0.005 ⁹	-0.004 ³	-0.004 ⁶	-0.004 ⁶	
	RMSE($\hat{\mu}$)	0.049 ¹	0.056 ⁹	0.061 ¹⁰	0.065 ¹¹	0.050 ⁴	0.049 ³	0.049 ²	0.050 ⁷	0.050 ⁵	0.055 ⁸	0.050 ⁶	
	\hat{D}_{abs}	0.031 ²	0.031 ⁸	0.031 ⁹	0.031 ³	0.031 ⁵	0.031 ⁶	0.031 ¹⁰	0.031 ¹⁰	0.031 ⁴	0.031 ⁷	0.032 ¹¹	
	\hat{D}_{max}	0.081 ²	0.082 ⁷	0.082 ⁹	0.082 ⁴	0.082 ⁵	0.082 ⁸	0.082 ¹⁰	0.082 ¹⁰	0.082 ³	0.082 ⁶	0.082 ¹¹	
	Total	14 ¹	53 ¹¹	50 ¹⁰	47 ⁹	27 ³	27 ³	38 ⁷	34 ⁵	34 ⁵	32 ⁴	45 ⁸	22 ²
Overall Total	6 ¹	32 ⁸	36 ¹⁰	36 ¹⁰	13 ²	13 ²	29 ⁷	14 ⁴	20 ⁵	25 ⁶	34 ⁹	13 ²	

Table 5 – Simulations results for $\mu = 0.8$ and $\beta = 2.0$.

n	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
20	Bias($\hat{\beta}$)	0.026 ³	-0.040 ⁴	0.055 ⁶	0.065 ⁸	0.049 ⁵	0.074 ¹¹	0.069 ⁹	-0.071 ¹⁰	-0.008 ²	-0.064 ⁷	0.007 ¹
	RMSE($\hat{\beta}$)	0.214 ²	0.216 ³	0.332 ¹⁰	0.337 ¹¹	0.252 ⁸	0.263 ⁹	0.224 ⁴	0.197 ¹	0.236 ⁷	0.234 ⁶	0.227 ⁵
	Bias($\hat{\alpha}$)	-0.002 ⁷	-0.002 ⁶	-0.005 ¹¹	0.000 ²	-0.001 ³	-0.003 ⁹	-0.002 ⁸	-0.002 ⁴	-0.003 ¹⁰	-0.000 ¹	-0.002 ⁵
	RMSE($\hat{\alpha}$)	0.039 ⁴	0.042 ¹⁰	0.043 ¹¹	0.041 ⁹	0.040 ⁵	0.039 ²	0.039 ³	0.039 ¹	0.039 ¹	0.040 ⁶	0.040 ⁷
	\hat{D}_{abs}	0.097 ⁸	0.097 ⁴	0.097 ⁷	0.097 ²	0.099 ¹⁰	0.097 ³	0.098 ⁹	0.098 ⁹	0.097 ⁶	0.097 ⁵	0.099 ¹¹
\hat{D}_{max}	0.221 ⁷	0.220 ³	0.221 ⁸	0.219 ¹	0.219 ¹	0.223 ¹⁰	0.221 ⁶	0.222 ⁹	0.221 ⁵	0.221 ⁴	0.220 ²	0.224 ¹¹
Total	31 ⁴	30 ³	53 ¹¹	33 ⁵	41 ⁹	40 ⁷	42 ¹⁰	27 ²	34 ⁶	25 ¹	40 ⁷	
50	Bias($\hat{\beta}$)	0.012 ⁵	-0.046 ¹¹	0.003 ²	0.002 ¹	0.020 ⁶	0.031 ⁸	0.025 ⁷	-0.044 ¹⁰	-0.006 ⁴	-0.039 ⁹	0.006 ³
	RMSE($\hat{\beta}$)	0.129 ³	0.149 ⁷	0.205 ¹⁰	0.205 ¹¹	0.146 ⁶	0.150 ⁹	0.128 ²	0.125 ¹	0.140 ⁵	0.149 ⁸	0.132 ⁴
	Bias($\hat{\alpha}$)	-0.001 ⁵	-0.001 ⁴	-0.000 ²	-0.002 ¹¹	-0.001 ⁶	-0.001 ⁹	-0.001 ⁸	-0.001 ³	-0.001 ⁷	0.000 ¹	-0.001 ¹⁰
	RMSE($\hat{\alpha}$)	0.025 ³	0.027 ⁹	0.028 ¹⁰	0.028 ¹¹	0.025 ⁷	0.024 ²	0.025 ⁵	0.024 ¹	0.025 ⁶	0.025 ⁸	0.025 ⁴
	\hat{D}_{abs}	0.062 ¹	0.063 ¹⁰	0.062 ³	0.062 ⁴	0.063 ¹¹	0.062 ⁸	0.062 ⁵	0.062 ⁷	0.062 ⁹	0.062 ⁶	0.062 ²
\hat{D}_{max}	0.152 ¹	0.154 ¹¹	0.152 ⁴	0.152 ³	0.153 ¹⁰	0.152 ⁵	0.153 ⁸	0.153 ⁷	0.153 ⁹	0.152 ⁶	0.152 ²	
Total	18 ¹	52 ¹¹	31 ⁴	41 ⁸	46 ¹⁰	41 ⁸	35 ⁵	29 ³	40 ⁷	38 ⁶	25 ²	
100	Bias($\hat{\beta}$)	0.006 ³	-0.038 ¹¹	-0.011 ⁵	-0.011 ⁶	0.011 ⁴	0.015 ⁸	0.012 ⁷	-0.026 ¹⁰	-0.003 ²	-0.024 ⁹	0.003 ¹
	RMSE($\hat{\beta}$)	0.089 ³	0.111 ⁹	0.152 ¹¹	0.149 ¹⁰	0.101 ⁶	0.102 ⁷	0.086 ¹	0.086 ²	0.097 ⁵	0.108 ⁸	0.090 ⁴
	Bias($\hat{\alpha}$)	-0.000 ³	-0.000 ⁷	0.001 ¹⁰	-0.003 ¹¹	-0.000 ¹	-0.000 ⁴	-0.000 ⁶	-0.000 ²	-0.000 ⁵	0.001 ⁸	-0.001 ⁹
	RMSE($\hat{\alpha}$)	0.017 ¹	0.019 ⁹	0.021 ¹⁰	0.021 ¹¹	0.017 ⁶	0.017 ⁴	0.017 ⁵	0.017 ²	0.017 ⁷	0.018 ⁸	0.017 ³
	\hat{D}_{abs}	0.044 ¹	0.044 ¹⁰	0.044 ⁸	0.044 ²	0.044 ⁹	0.044 ⁴	0.044 ³	0.044 ⁶	0.044 ⁷	0.044 ⁵	0.045 ¹¹
\hat{D}_{max}	0.113 ³	0.113 ⁸	0.113 ⁷	0.112 ¹	0.113 ¹⁰	0.112 ²	0.113 ⁴	0.113 ⁶	0.113 ⁹	0.113 ⁵	0.114 ¹¹	
Total	14 ¹	54 ¹¹	51 ¹⁰	41 ⁸	36 ⁶	29 ⁴	26 ²	28 ³	35 ⁵	43 ⁹	39 ⁷	
200	Bias($\hat{\beta}$)	0.002 ²	-0.031 ¹¹	-0.015 ⁸	-0.015 ¹⁰	0.006 ⁴	0.007 ⁶	0.006 ⁵	-0.015 ⁹	-0.002 ¹	-0.014 ⁷	0.003 ³
	RMSE($\hat{\beta}$)	0.063 ³	0.082 ⁹	0.114 ¹¹	0.112 ¹⁰	0.070 ⁷	0.070 ⁶	0.060 ¹	0.060 ²	0.068 ⁵	0.077 ⁸	0.063 ⁴
	Bias($\hat{\alpha}$)	-0.000 ³	-0.000 ⁴	0.002 ¹⁰	-0.003 ¹¹	-0.000 ²	-0.000 ⁸	-0.000 ⁶	-0.000 ¹	-0.000 ⁵	0.000 ⁹	-0.000 ⁷
	RMSE($\hat{\alpha}$)	0.012 ¹	0.014 ⁹	0.015 ¹⁰	0.016 ¹¹	0.012 ⁷	0.012 ⁴	0.012 ³	0.012 ²	0.012 ⁵	0.012 ⁶	0.012 ⁸
	\hat{D}_{abs}	0.031 ⁹	0.031 ⁴	0.031 ¹¹	0.031 ¹⁰	0.031 ⁶	0.031 ¹	0.031 ²	0.031 ¹⁰	0.031 ⁵	0.031 ³	0.031 ⁸
\hat{D}_{max}	0.082 ¹⁰	0.081 ⁴	0.082 ¹¹	0.082 ⁶	0.082 ⁹	0.081 ²	0.081 ³	0.082 ⁸	0.081 ⁵	0.081 ¹	0.082 ⁷	
Total	28 ⁴	41 ⁹	61 ¹¹	54 ¹⁰	36 ⁷	27 ³	20 ¹	32 ⁵	26 ²	34 ⁶	37 ⁸	
Overall Total	10 ¹	34 ¹⁰	36 ¹¹	31 ⁸	32 ⁹	22 ⁵	18 ³	13 ²	20 ⁴	22 ⁵	24 ⁷	

Table 6 – Overall performance of estimation methods.

Scenario	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
$(\mu = 0.2, \beta = 0.5)$	6 ¹	30 ⁸	43 ¹¹	37 ¹⁰	27 ⁷	33 ⁹	15 ⁴	14 ³	16 ⁵	26 ⁶	11 ²
$(\mu = 0.2, \beta = 1.5)$	19 ³	39 ¹⁰	40 ¹¹	31 ⁸	19 ³	31 ⁸	19 ³	20 ⁷	13 ²	19 ³	9 ¹
$(\mu = 0.2, \beta = 2.0)$	14 ¹	37 ⁹	39 ¹⁰	40 ¹¹	20 ⁶	18 ⁵	17 ⁴	21 ⁷	21 ⁷	15 ³	14 ¹
$(\mu = 0.4, \beta = 0.5)$	9 ¹	32 ⁹	40 ¹⁰	41 ¹¹	23 ⁶	26 ⁷	12 ³	10 ²	18 ⁵	28 ⁸	17 ⁴
$(\mu = 0.4, \beta = 1.5)$	12 ²	32 ⁹	41 ¹¹	39 ¹⁰	29 ⁸	27 ⁷	13 ³	20 ⁵	21 ⁶	13 ³	11 ¹
$(\mu = 0.4, \beta = 2.0)$	23 ⁶	28 ⁷	42 ¹¹	32 ⁹	32 ⁹	19 ⁵	17 ⁴	29 ⁸	6 ¹	11 ²	15 ³
$(\mu = 0.6, \beta = 0.5)$	18 ⁵	31 ⁹	35 ¹⁰	39 ¹¹	23 ⁷	17 ⁴	20 ⁶	16 ³	14 ²	28 ⁸	13 ¹
$(\mu = 0.6, \beta = 1.5)$	25 ⁷	36 ⁹	38 ¹⁰	38 ¹⁰	11 ²	34 ⁸	23 ⁶	12 ³	10 ¹	13 ⁴	20 ⁵
$(\mu = 0.6, \beta = 2.0)$	8 ¹	36 ¹⁰	35 ⁹	40 ¹¹	25 ⁶	26 ⁷	15 ³	16 ⁴	22 ⁵	26 ⁷	12 ²
$(\mu = 0.8, \beta = 0.5)$	6 ¹	32 ⁸	36 ¹⁰	36 ¹⁰	13 ²	29 ⁷	14 ⁴	20 ⁵	25 ⁶	34 ⁹	13 ²
$(\mu = 0.8, \beta = 1.5)$	9 ¹	36 ¹⁰	33 ⁹	39 ¹¹	15 ³	19 ⁵	14 ²	27 ⁷	18 ⁴	29 ⁸	19 ⁵
$(\mu = 0.8, \beta = 2.0)$	10 ¹	34 ¹⁰	36 ¹¹	31 ⁸	32 ⁹	22 ⁵	18 ³	13 ²	20 ⁴	22 ⁵	24 ⁷
Total	159 ¹	403 ⁹	458 ¹¹	443 ¹⁰	269 ⁷	301 ⁸	197 ³	218 ⁵	204 ⁴	264 ⁶	178 ²

Table 7 – Simulations results of interval estimation for $\mu = 0.2$ and $\beta = 0.5$.

n	Qid	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
20	CP($\hat{\beta}$)	0.944 ⁵	0.909 ⁸	0.946 ³	0.944 ⁴	0.929 ⁷	0.909 ⁹	0.903 ¹⁰	0.853 ¹¹	0.951 ²	0.958 ⁶	0.951 ¹
	CP($\hat{\alpha}$)	0.949 ²	0.962 ⁸	0.932 ¹⁰	0.923 ¹¹	0.939 ⁷	0.940 ⁶	0.941 ⁵	0.962 ⁹	0.952 ³	0.959 ⁴	0.951 ¹
	AW($\hat{\beta}$)	0.421 ³	0.398 ²	0.727 ¹⁰	0.729 ¹¹	0.510 ⁷	0.433 ⁸	0.444 ⁵	0.335 ¹	0.450 ⁶	0.589 ⁹	0.433 ⁴
	AW($\hat{\alpha}$)	0.464 ³	0.521 ¹¹	0.482 ⁶	0.517 ⁹	0.468 ⁴	0.451 ²	0.445 ¹	0.505 ⁸	0.483 ⁷	0.520 ¹⁰	0.476 ⁵
Total	13 ²	29 ⁷	29 ⁷	29 ⁷	35 ¹¹	25 ⁵	25 ⁵	21 ⁴	29 ⁷	18 ³	29 ⁷	11 ¹
50	Qid	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	CP($\hat{\beta}$)	0.949 ²	0.864 ¹¹	0.949 ⁴	0.946 ⁷	0.946 ⁶	0.938 ⁸	0.934 ⁹	0.878 ¹⁰	0.949 ³	0.947 ⁵	0.950 ¹
	CP($\hat{\alpha}$)	0.945 ⁷	0.962 ¹¹	0.950 ²	0.946 ⁵	0.943 ¹⁰	0.944 ⁸	0.946 ⁶	0.956 ⁹	0.948 ³	0.950 ¹	0.947 ⁴
	AW($\hat{\beta}$)	0.249 ³	0.259 ⁵	0.407 ¹¹	0.406 ¹⁰	0.285 ⁷	0.290 ⁸	0.248 ²	0.217 ¹	0.271 ⁶	0.368 ⁹	0.255 ⁴
AW($\hat{\alpha}$)	0.306 ³	0.354 ¹⁰	0.370 ¹¹	0.348 ⁹	0.309 ⁵	0.303 ²	0.300 ¹	0.321 ⁷	0.312 ⁶	0.336 ⁸	0.307 ⁴	
Total	15 ²	37 ¹¹	28 ⁸	31 ¹⁰	28 ⁸	26 ⁶	18 ³	27 ⁷	18 ³	23 ⁵	13 ¹	
100	Qid	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	CP($\hat{\beta}$)	0.951 ³	0.847 ¹¹	0.935 ⁸	0.930 ⁹	0.945 ⁵	0.945 ⁶	0.940 ⁷	0.896 ¹⁰	0.950 ²	0.952 ⁴	0.950 ¹
	CP($\hat{\alpha}$)	0.948 ³	0.963 ¹¹	0.960 ¹⁰	0.956 ⁹	0.949 ²	0.947 ⁵	0.947 ⁵	0.953 ⁷	0.951 ¹	0.954 ⁸	0.948 ⁴
	AW($\hat{\beta}$)	0.173 ³	0.194 ⁶	0.292 ¹⁰	0.292 ¹¹	0.195 ⁷	0.195 ⁸	0.169 ²	0.157 ¹	0.189 ⁵	0.259 ⁹	0.175 ⁴
AW($\hat{\alpha}$)	0.217 ³	0.252 ⁹	0.286 ¹¹	0.253 ¹⁰	0.220 ⁵	0.217 ²	0.215 ¹	0.224 ⁷	0.221 ⁶	0.221 ⁶	0.218 ⁴	
Total	12 ¹	37 ⁹	39 ¹⁰	39 ¹⁰	19 ⁵	21 ⁶	15 ⁴	25 ⁷	14 ³	29 ⁸	13 ²	
200	Qid	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	CP($\hat{\beta}$)	0.957 ⁷	0.829 ¹¹	0.927 ⁸	0.910 ¹⁰	0.954 ⁴	0.953 ²	0.950 ¹	0.913 ⁹	0.955 ⁶	0.955 ⁵	0.953 ²
	CP($\hat{\alpha}$)	0.945 ⁹	0.952 ⁴	0.950 ¹	0.948 ⁵	0.948 ⁵	0.945 ¹⁰	0.946 ⁸	0.948 ³	0.947 ⁷	0.949 ²	0.944 ¹¹
	AW($\hat{\beta}$)	0.121 ³	0.145 ⁸	0.215 ¹¹	0.215 ¹⁰	0.135 ⁷	0.135 ⁶	0.111 ²	0.112 ¹	0.133 ⁵	0.183 ⁹	0.122 ⁴
AW($\hat{\alpha}$)	0.154 ²	0.178 ⁹	0.215 ¹¹	0.186 ¹⁰	0.155 ⁵	0.154 ⁴	0.153 ¹	0.156 ⁷	0.155 ⁶	0.168 ⁸	0.154 ³	
Total	21 ⁴	32 ¹⁰	31 ⁹	35 ¹¹	21 ⁴	22 ⁶	12 ¹	20 ²	24 ⁷	24 ⁷	20 ²	
Overall Total	9 ²	37 ¹⁰	34 ⁹	42 ¹¹	22 ⁵	23 ⁶	12 ³	23 ⁶	16 ⁴	27 ⁸	6 ¹	

Table 8 – Simulations results of interval estimation for $\mu = 0.2$ and $\beta = 2.0$.

n	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
20	CP($\hat{\beta}$)	0.939 ⁵	0.905 ⁸	0.958 ⁴	0.949 ¹	0.934 ⁶	0.915 ⁷	0.905 ⁹	0.857 ¹¹	0.947 ²	0.901 ¹⁰	0.955 ³
	CP($\hat{\mu}$)	0.940 ⁵	0.961 ⁶	0.935 ⁹	0.921 ¹¹	0.933 ¹⁰	0.940 ⁴	0.935 ⁸	0.960 ³	0.948 ¹	0.962 ⁷	0.947 ²
	AW($\hat{\beta}$)	1.671 ⁴	1.575 ²	2.493 ¹⁰	2.506 ¹¹	1.968 ⁸	2.041 ⁹	1.743 ⁶	1.337 ¹	1.773 ⁷	1.627 ³	1.705 ⁵
	AW($\hat{\mu}$)	0.124 ⁴	0.144 ¹¹	0.138 ⁸	0.139 ⁹	0.123 ³	0.121 ²	0.119 ¹	0.137 ⁷	0.130 ⁶	0.143 ¹⁰	0.128 ⁵
	Total	18 ³	27 ⁷	31 ¹⁰	32 ¹¹	27 ⁷	22 ⁴	24 ⁶	22 ⁴	30 ⁹	16 ²	30 ⁹
50	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	CP($\hat{\beta}$)	0.946 ²	0.865 ¹¹	0.940 ⁶	0.945 ⁵	0.946 ³	0.938 ⁷	0.923 ⁸	0.876 ¹⁰	0.946 ³	0.909 ⁹	0.949 ¹
	CP($\hat{\mu}$)	0.945 ⁷	0.965 ¹¹	0.954 ⁶	0.948 ²	0.945 ⁸	0.945 ⁹	0.948 ³	0.956 ¹⁰	0.953 ⁴	0.953 ⁴	0.951 ¹
	AW($\hat{\beta}$)	0.999 ³	1.037 ⁵	1.612 ¹⁰	1.624 ¹¹	1.144 ⁸	1.162 ⁹	0.997 ²	0.870 ¹	1.085 ⁷	1.082 ⁶	1.022 ⁴
	AW($\hat{\mu}$)	0.077 ³	0.090 ⁹	0.094 ¹¹	0.091 ¹⁰	0.078 ⁴	0.077 ²	0.076 ¹	0.081 ⁷	0.079 ⁶	0.084 ⁸	0.078 ⁵
Total	15 ³	36 ¹¹	33 ¹⁰	28 ⁸	23 ⁵	27 ⁶	14 ²	28 ⁸	21 ⁴	27 ⁶	11 ¹	
100	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	CP($\hat{\beta}$)	0.947 ⁴	0.844 ¹¹	0.930 ⁷	0.926 ⁹	0.947 ⁴	0.950 ¹	0.937 ⁶	0.897 ¹⁰	0.950 ¹	0.927 ⁸	0.949 ³
	CP($\hat{\mu}$)	0.945 ⁸	0.965 ¹¹	0.955 ⁹	0.953 ⁶	0.949 ¹	0.948 ⁴	0.944 ¹⁰	0.952 ³	0.952 ⁵	0.954 ⁷	0.948 ²
	AW($\hat{\beta}$)	0.693 ³	0.775 ⁶	1.162 ¹⁰	1.167 ¹¹	0.781 ⁷	0.782 ⁸	0.678 ²	0.627 ¹	0.756 ⁵	0.792 ⁹	0.703 ⁴
	AW($\hat{\mu}$)	0.054 ³	0.063 ⁹	0.070 ¹¹	0.067 ¹⁰	0.055 ⁵	0.054 ²	0.054 ¹	0.056 ⁷	0.055 ⁶	0.058 ⁸	0.054 ⁴
Total	18 ⁵	37 ¹⁰	37 ¹⁰	36 ⁹	17 ³	15 ²	19 ⁶	21 ⁷	17 ³	32 ⁸	13 ¹	
200	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
	CP($\hat{\beta}$)	0.954 ⁵	0.829 ¹¹	0.927 ⁸	0.915 ¹⁰	0.948 ⁴	0.948 ³	0.945 ⁶	0.918 ⁹	0.950 ¹	0.936 ⁷	0.949 ²
	CP($\hat{\mu}$)	0.944 ⁷	0.961 ¹⁰	0.947 ⁴	0.950 ¹	0.944 ⁸	0.942 ⁹	0.938 ¹¹	0.950 ²	0.946 ⁶	0.951 ³	0.947 ⁵
	AW($\hat{\beta}$)	0.485 ³	0.580 ⁹	0.859 ¹⁰	0.861 ¹¹	0.547 ¹	0.541 ⁶	0.471 ²	0.450 ¹	0.531 ⁵	0.575 ⁸	0.490 ⁴
	AW($\hat{\mu}$)	0.038 ³	0.044 ⁹	0.051 ¹¹	0.049 ¹⁰	0.039 ⁵	0.038 ⁴	0.038 ¹	0.039 ⁷	0.039 ⁶	0.040 ⁸	0.038 ²
Total	18 ²	39 ¹¹	33 ¹⁰	32 ⁹	24 ⁷	22 ⁶	20 ⁵	19 ⁴	18 ²	26 ⁸	13 ¹	
Overall Total	13 ³	39 ¹⁰	40 ¹¹	37 ⁹	22 ⁶	18 ⁴	19 ⁵	23 ⁷	11 ²	31 ⁸	4 ¹	

Table 9 – Simulations results of interval estimation for $\mu = 0.8$ and $\beta = 0.5$.

n	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
20	CP($\hat{\beta}$)	0.944 ⁵	0.907 ⁸	0.942 ⁶	0.946 ⁴	0.935 ⁷	0.906 ⁹	0.901 ¹⁰	0.859 ¹¹	0.953 ³	0.952 ¹	0.948 ²
	CP($\hat{\alpha}$)	0.946 ²	0.958 ⁵	0.925 ¹⁰	0.921 ¹¹	0.927 ⁹	0.939 ⁶	0.936 ⁸	0.962 ⁷	0.944 ⁴	0.944 ³	0.949 ¹
	AW($\hat{\beta}$)	0.423 ³	0.395 ²	0.735 ¹¹	0.731 ¹⁰	0.511 ⁷	0.544 ⁸	0.444 ⁵	0.334 ¹	0.452 ⁶	0.594 ⁹	0.434 ⁴
	AW($\hat{\alpha}$)	0.465 ⁴	0.527 ¹¹	0.520 ⁹	0.490 ⁷	0.455 ³	0.452 ²	0.447 ¹	0.511 ⁸	0.485 ⁶	0.525 ¹⁰	0.479 ⁵
Total	14 ²	26 ⁷	36 ¹¹	32 ¹⁰	26 ⁷	25 ⁶	24 ⁵	27 ⁹	19 ³	23 ⁴	12 ¹	
50	CP($\hat{\beta}$)	0.946 ³	0.864 ¹¹	0.944 ⁶	0.941 ⁷	0.946 ⁴	0.930 ⁸	0.923 ⁹	0.877 ¹⁰	0.953 ²	0.951 ¹	0.946 ⁴
	CP($\hat{\alpha}$)	0.945 ⁷	0.958 ¹⁰	0.953 ⁵	0.949 ²	0.941 ¹¹	0.945 ⁹	0.945 ⁶	0.955 ⁸	0.947 ⁴	0.947 ³	0.950 ¹
	AW($\hat{\beta}$)	0.250 ³	0.258 ⁵	0.411 ¹¹	0.408 ¹⁰	0.286 ⁷	0.291 ⁸	0.249 ²	0.216 ¹	0.271 ⁶	0.370 ⁹	0.256 ⁴
	AW($\hat{\alpha}$)	0.305 ⁴	0.354 ¹⁰	0.349 ⁹	0.371 ¹¹	0.303 ³	0.302 ²	0.300 ¹	0.323 ⁷	0.313 ⁶	0.338 ⁸	0.307 ⁵
Total	17 ²	36 ¹¹	31 ¹⁰	30 ⁹	25 ⁶	27 ⁸	18 ³	26 ⁷	18 ³	21 ⁵	14 ¹	
100	CP($\hat{\beta}$)	0.949 ²	0.841 ¹¹	0.931 ⁹	0.936 ⁸	0.951 ¹	0.945 ⁵	0.942 ⁷	0.895 ¹⁰	0.953 ⁴	0.957 ⁶	0.948 ³
	CP($\hat{\alpha}$)	0.943 ⁹	0.961 ¹¹	0.953 ³	0.954 ⁴	0.944 ⁶	0.941 ¹⁰	0.944 ⁸	0.956 ⁶	0.949 ¹	0.945 ⁵	0.949 ²
	AW($\hat{\beta}$)	0.173 ³	0.193 ⁶	0.292 ¹⁰	0.292 ¹¹	0.195 ⁷	0.196 ⁸	0.169 ²	0.157 ¹	0.189 ⁵	0.260 ⁹	0.176 ⁴
	AW($\hat{\alpha}$)	0.217 ³	0.253 ⁹	0.254 ¹⁰	0.285 ¹¹	0.217 ⁵	0.217 ²	0.215 ¹	0.225 ⁷	0.221 ⁶	0.238 ⁸	0.217 ⁴
Total	17 ³	37 ¹¹	32 ⁹	34 ¹⁰	19 ⁵	25 ⁷	18 ⁴	24 ⁶	16 ²	28 ⁸	13 ¹	
200	CP($\hat{\beta}$)	0.953 ⁶	0.828 ¹¹	0.925 ⁸	0.914 ¹⁰	0.947 ⁵	0.947 ⁴	0.949 ²	0.914 ⁹	0.954 ⁷	0.948 ³	0.950 ¹
	CP($\hat{\alpha}$)	0.948 ²	0.956 ¹⁰	0.946 ⁹	0.958 ¹¹	0.948 ⁵	0.949 ¹	0.946 ⁸	0.952 ⁴	0.953 ⁷	0.948 ³	0.947 ⁶
	AW($\hat{\beta}$)	0.121 ³	0.145 ⁸	0.215 ¹⁰	0.215 ¹¹	0.135 ⁷	0.135 ⁶	0.117 ²	0.112 ¹	0.133 ⁵	0.183 ⁹	0.122 ⁴
	AW($\hat{\alpha}$)	0.154 ³	0.178 ⁹	0.186 ¹⁰	0.215 ¹¹	0.155 ⁵	0.154 ⁴	0.153 ¹	0.157 ⁷	0.155 ⁶	0.167 ⁸	0.154 ²
Total	14 ³	38 ¹⁰	37 ⁹	43 ¹¹	22 ⁶	15 ⁴	13 ¹	21 ⁵	25 ⁸	23 ⁷	13 ¹	
Overall Total	10 ²	39 ⁹	39 ⁹	40 ¹¹	24 ⁵	25 ⁷	13 ³	27 ⁸	16 ⁴	24 ⁵	4 ¹	

Table 10 – Simulations results of interval estimation for $\mu = 0.8$ and $\beta = 2.0$.

n	Qtd	AD	AD2	AD2L	AD2R	ADR	CvM	MLE	MPS	OLS	PCE	WLS
20	CP($\hat{\beta}$)	0.944 ⁵	0.907 ⁸	0.953 ³	0.953 ²	0.935 ⁶	0.908 ⁷	0.903 ⁹	0.867 ¹¹	0.951 ¹	0.901 ¹⁰	0.946 ⁴
	CP($\hat{\mu}$)	0.944 ⁴	0.963 ⁷	0.931 ¹⁰	0.919 ¹¹	0.932 ⁹	0.942 ⁵	0.940 ⁶	0.964 ⁸	0.950 ¹	0.955 ³	0.946 ²
	AW($\hat{\beta}$)	1.670 ⁴	1.584 ²	2.505 ¹⁰	2.513 ¹¹	1.958 ⁸	2.047 ⁹	1.746 ⁶	1.339 ¹	1.767 ⁷	1.622 ³	1.711 ⁵
	AW($\hat{\mu}$)	0.124 ⁴	0.144 ¹¹	0.139 ⁹	0.136 ⁷	0.123 ³	0.121 ²	0.119 ¹	0.136 ⁸	0.130 ⁶	0.143 ¹⁰	0.127 ⁵
Total	17 ³	28 ⁸	32 ¹¹	31 ¹⁰	26 ⁶	23 ⁵	22 ⁴	28 ⁸	15 ¹	26 ⁶	16 ²	
50	CP($\hat{\beta}$)	0.943 ⁴	0.858 ¹¹	0.946 ²	0.942 ⁵	0.939 ⁶	0.928 ⁷	0.927 ⁸	0.873 ¹⁰	0.947 ¹	0.912 ⁹	0.944 ³
	CP($\hat{\mu}$)	0.945 ⁵	0.960 ¹¹	0.951 ¹	0.949 ²	0.943 ⁹	0.947 ¹⁰	0.941 ¹⁰	0.957 ⁸	0.946 ⁴	0.955 ⁶	0.948 ³
	AW($\hat{\beta}$)	1.002 ³	1.036 ⁵	1.625 ¹¹	1.623 ¹⁰	1.145 ⁸	1.167 ⁹	0.995 ²	0.866 ¹	1.082 ⁷	1.082 ⁶	1.023 ⁴
	AW($\hat{\mu}$)	0.077 ³	0.090 ⁹	0.090 ¹⁰	0.093 ¹¹	0.077 ⁴	0.076 ²	0.076 ¹	0.081 ⁷	0.079 ⁶	0.084 ⁸	0.078 ⁵
Total	15 ¹	36 ¹¹	24 ⁵	28 ⁹	27 ⁸	25 ⁶	21 ⁴	26 ⁷	18 ³	29 ¹⁰	15 ¹	
100	CP($\hat{\beta}$)	0.946 ⁴	0.851 ¹¹	0.932 ⁸	0.933 ⁷	0.946 ³	0.933 ⁶	0.942 ⁵	0.895 ¹⁰	0.951 ¹	0.925 ⁹	0.948 ²
	CP($\hat{\mu}$)	0.951 ³	0.955 ⁹	0.956 ¹⁰	0.953 ⁵	0.948 ⁴	0.949 ²	0.946 ⁷	0.959 ¹¹	0.945 ⁸	0.953 ⁶	0.950 ¹
	AW($\hat{\beta}$)	0.694 ³	0.777 ⁶	1.170 ¹¹	1.169 ¹⁰	0.782 ⁷	0.784 ⁸	0.677 ²	0.626 ¹	0.755 ⁵	0.793 ⁹	0.702 ⁴
	AW($\hat{\mu}$)	0.054 ³	0.063 ⁹	0.066 ¹⁰	0.069 ¹¹	0.055 ⁴	0.054 ²	0.054 ¹	0.056 ⁷	0.055 ⁶	0.058 ⁸	0.055 ⁵
Total	13 ²	35 ¹⁰	39 ¹¹	33 ⁹	18 ⁴	18 ⁴	15 ³	29 ⁷	20 ⁶	32 ⁸	12 ¹	
200	CP($\hat{\beta}$)	0.946 ⁴	0.841 ¹¹	0.930 ⁸	0.916 ⁹	0.948 ²	0.944 ⁵	0.944 ⁶	0.914 ¹⁰	0.949 ¹	0.930 ⁷	0.947 ³
	CP($\hat{\mu}$)	0.947 ⁵	0.959 ¹¹	0.951 ¹	0.952 ⁴	0.948 ⁵	0.954 ⁸	0.948 ³	0.952 ⁶	0.951 ²	0.956 ¹⁰	0.946 ⁹
	AW($\hat{\beta}$)	0.485 ³	0.581 ⁹	0.863 ¹⁰	0.863 ¹¹	0.543 ⁷	0.540 ⁶	0.469 ²	0.450 ¹	0.531 ⁵	0.575 ⁸	0.489 ⁴
	AW($\hat{\mu}$)	0.038 ²	0.044 ⁹	0.049 ¹⁰	0.051 ¹¹	0.039 ⁵	0.038 ⁴	0.038 ¹	0.039 ⁷	0.039 ⁶	0.040 ⁸	0.038 ³
Total	16 ³	40 ¹¹	29 ⁸	35 ¹⁰	19 ⁴	23 ⁶	12 ¹	24 ⁷	14 ²	33 ⁹	19 ⁴	
Overall Total	9 ²	40 ¹¹	35 ⁹	38 ¹⁰	22 ⁶	21 ⁵	12 ³	29 ⁷	12 ³	33 ⁸	8 ¹	

Table 11 – Overall performance of estimation methods with respect the interval estimation.

Scenario	AD	AD2	AD2L	AD2R	ADR	CVM	MLE	MPS	OLS	PCE	WLS
($\mu = 0.2, \beta = 0.5$)	9 ²	37 ¹⁰	34 ⁹	42 ¹¹	22 ⁵	23 ⁶	12 ³	23 ⁶	16 ⁴	27 ⁸	6 ¹
($\mu = 0.2, \beta = 1.5$)	6 ¹	39 ¹⁰	42 ¹¹	34 ⁹	26 ⁷	29 ⁸	12 ³	24 ⁵	15 ⁴	25 ⁶	7 ²
($\mu = 0.2, \beta = 2.0$)	13 ³	39 ¹⁰	40 ¹¹	37 ⁹	22 ⁶	18 ⁴	19 ⁵	23 ⁷	11 ²	31 ⁸	4 ¹
($\mu = 0.4, \beta = 0.5$)	9 ²	39 ¹⁰	38 ⁹	39 ¹⁰	23 ⁵	27 ⁷	10 ³	27 ⁷	16 ⁴	24 ⁶	7 ¹
($\mu = 0.4, \beta = 1.5$)	8 ¹	42 ¹¹	36 ⁹	41 ¹⁰	24 ⁶	25 ⁷	14 ³	26 ⁸	19 ⁵	14 ³	9 ²
($\mu = 0.4, \beta = 2.0$)	11 ³	40 ¹⁰	40 ¹⁰	37 ⁹	29 ⁸	20 ⁵	7 ¹	26 ⁷	12 ⁴	21 ⁶	8 ²
($\mu = 0.6, \beta = 0.5$)	9 ²	38 ¹⁰	33 ⁹	42 ¹¹	26 ⁶	30 ⁷	14 ⁴	30 ⁷	11 ³	20 ⁵	5 ¹
($\mu = 0.6, \beta = 1.5$)	19 ⁶	41 ¹¹	37 ⁹	40 ¹⁰	14 ³	19 ⁶	14 ³	28 ⁸	15 ⁵	12 ²	9 ¹
($\mu = 0.6, \beta = 2.0$)	5 ¹	38 ¹⁰	36 ⁹	41 ¹¹	27 ⁷	20 ⁵	17 ³	30 ⁸	19 ⁴	21 ⁶	6 ²
($\mu = 0.8, \beta = 0.5$)	10 ²	39 ⁹	39 ⁹	40 ¹¹	24 ⁵	25 ⁷	13 ³	27 ⁸	16 ⁴	24 ⁵	4 ¹
($\mu = 0.8, \beta = 1.5$)	5 ¹	35 ⁹	35 ⁹	41 ¹¹	25 ⁶	24 ⁵	11 ²	29 ⁸	12 ³	27 ⁷	14 ⁴
($\mu = 0.8, \beta = 2.0$)	9 ²	40 ¹¹	35 ⁹	38 ¹⁰	22 ⁶	21 ⁵	12 ³	29 ⁷	12 ³	33 ⁸	8 ¹
Total	113 ²	467 ¹⁰	445 ⁹	472 ¹¹	284 ⁷	281 ⁶	155 ³	322 ⁸	174 ⁴	279 ⁵	87 ¹

4 ILLUSTRATIVE EXAMPLES

In this section, the performance of the eleven estimation methods is compared through two real data applications.

The first data (data set I) is available in software R and corresponds to 48 observations of twelve core samples from petroleum reservoirs that were sampled by four cross-sections. The second data set (data set-II) can be found in [4] and represents the total milk production in the first birth of 107 cows from SINDI race.

The parameter estimates and their corresponding Bootstrap confidence intervals for all estimation methods considered are summarized in Tables 12 and 13. We also present the results of formal goodness-of-fit tests, the Kolmogorov-Smirnov (KS) test, in order to show that the unit-Logistic distribution can be used to model these two data sets.

From Table 12 we can see that all estimates provide a good fit to the data set. It is also observed that the AD2L and MPS estimators give the shortest confidence intervals for μ and β , respectively.

Table 12 – Parameter estimates, 95% confidence intervals based on parametric Bootstrap and K-S test: data set I.

Method	μ	LCL	UCL	β	LCL	UCL	KS (<i>p</i> -value)
MLE	0.2033	0.1847	0.2240	3.8276	3.0530	5.0733	0.0979 (0.7469)
MPS	0.2019	0.1792	0.2236	3.5707	2.6905	4.3026	0.0907 (0.8242)
PCE	0.2058	0.1808	0.2298	3.2550	2.3271	4.0154	0.1114 (0.5907)
OLS	0.2014	0.1792	0.2226	3.6391	2.7666	4.7659	0.0879 (0.8520)
WLS	0.2034	0.1814	0.2249	3.7032	2.8428	4.8165	0.0992 (0.7326)
CvM	0.2013	0.1793	0.2228	3.7596	3.0188	5.0458	0.0865 (0.8649)
AD	0.2034	0.1834	0.2257	3.7135	2.9679	4.7314	0.0990 (0.7351)
ADR	0.2018	0.1798	0.2262	3.4077	2.6352	4.6341	0.0991 (0.7335)
AD2R	0.2002	0.1736	0.2273	3.1894	2.0618	4.8186	0.1192 (0.5030)
AD2L	0.1958	0.1788	0.2163	4.9271	3.2415	7.3709	0.1379 (0.3205)
AD2	0.2103	0.1864	0.2363	3.6283	2.4818	4.5025	0.1374 (0.3249)

L(U)CL lower (upper) confidence limit.

The results in Table 13 indicate that the CvM estimates do not provide a good fit to this data set as per KS statistic is concerned. It is also observed that MLE has the lowest value of KS. It is also noteworthy, that MLE and ADR have the shortest confidence intervals for μ and β .

5 CONCLUDING REMARKS

In this paper, we have performed an extensive simulation study to compare eleven aforementioned methods of estimation. We have compared estimators with respect to bias, root mean-squared error, the average absolute difference between the theoretical and empirical estimate of

Table 13 – Parameter estimates, 95% confidence intervals based on parametric Bootstrap and KS test: data set II.

Method	μ	LCL	UCL	β	LCL	UCL	KS (p -value)
MLE	0.4729	0.4317	0.5148	1.9103	1.6565	2.2716	0.0571 (0.8767)
MPS	0.4723	0.4019	0.5355	1.8338	1.3818	2.2098	0.0618 (0.8081)
PCE	0.4686	0.4037	0.5298	1.9449	1.4745	2.3979	0.0580 (0.8642)
OLS	0.4789	0.4166	0.5341	2.0873	1.5868	2.7337	0.0695 (0.6788)
WLS	0.4762	0.4137	0.5331	2.0752	1.5931	2.6993	0.0682 (0.7026)
CvM	0.2013	0.1793	0.2228	3.7596	3.0188	5.0458	0.7350 (0.0000)
AD	0.4733	0.4304	0.5163	1.9681	1.6871	2.3073	0.0610 (0.8201)
ADR	0.4798	0.4441	0.5181	2.2317	1.8949	2.7488	0.0765 (0.5584)
AD2R	0.4876	0.4442	0.5254	2.4237	1.7670	3.1201	0.0848 (0.4254)
AD2L	0.4981	0.4313	0.5785	1.3952	0.9988	1.7848	0.1480 (0.0184)
AD2	0.4419	0.3867	0.4974	1.6351	1.2378	1.8874	0.1207 (0.0886)

L(U)CL lower (upper) confidence limit.

the distribution functions, and the maximum absolute difference between the theoretical and empirical distribution functions. We have also calculated the coverage probability and the average width of the Bootstrap confidence intervals. We have also compared estimators by two real data applications. The simulation results show that AD estimators is the best performing estimator in terms of biases and RMSE. The next best performing estimators is the WLS estimators, followed by MLE. The real data applications show that the AD2L and MPS estimators give the shortest confidence intervals for μ and β , respectively for the data set I and MLE and ADR have the shortest confidence intervals for the data set II. Hence, we can argue that the AD estimators, weighted least squares estimators, AD2L, MPS, ADR and ML estimators are among the best performing estimators for unit-logistic distribution.

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